

### Offline Components: Collaborative Filtering in Cold-start Situations

### Problem







### **Model Choices**

- Feature-based (or content-based) approach
  - Use features to predict response
    - (regression, Bayes Net, mixture models, ...)
  - Limitation: need predictive features
    - Bias often high, does not capture signals at granular levels
- Collaborative filtering (CF aka Memory based)
  - Make recommendation based on past user-item interaction
    - User-user, item-item, matrix factorization, ...
    - See [Adomavicius & Tuzhilin, TKDE, 2005], [Konstan, SIGMOD'08 Tutorial], etc.
  - Better performance for old users and old items
  - Does not naturally handle new users and new items (cold-start)



### **Collaborative Filtering (Memory based methods)**

User-User Similarity

$$p_{a,j} = \overline{v}_a + \kappa \sum_{i=1}^n w(a,i)(v_{i,j} - \overline{v}_i)$$

Item-Item similarities, incorporating both

**Estimating Similarities** 

- Pearson's correlation
- Optimization based (Koren et al)



### How to Deal with the Cold-Start Problem

- Heuristic-based approaches
  - Linear combination of regression and CF models
  - Filterbot
    - Add user features as psuedo users and do collaborative filtering
  - Hybrid approaches
    - Use content based to fill up entries, then use CF
- Matrix Factorization
  - Good performance on Netflix (Koren, 2009)
- Model-based approaches
  - Bilinear random-effects model (probabilistic matrix factorization)
    - Good on Netflix data [Ruslan et al ICML, 2009]
  - Add feature-based regression to matrix factorization
    - (Agarwal and Chen, 2009)
  - Add topic discovery (from textual items) to matrix factorization
    - (Agarwal and Chen, 2009; Chun and Blei, 2011)

### Per-item regression models

- When tracking users by cookies, distribution of visit patters could get extremely skewed
  - Majority of cookies have 1-2 visits
- Per item models (regression) based on user covariates attractive in such cases

 $y_{ijt} \sim \text{Bernoulli}(p_{ijt})$  $s_{ijt} = \log \frac{p_{ijt}}{1 - p_{ijt}}$  $s_{ijt} = \mathbf{x}'_{it}\mathbf{A}\mathbf{x}_j + \mathbf{x}'_{it}\mathbf{v}_{jt}$ 

### Several per-item regressions: Multi-task learning

• Agarwal, Chen and Elango, KDD, 2010

$$s_{ijt} = x'_{it}Ax_j + x'_{it}B\theta_j \cdot B$$
 estimated  
•retrospective data  
Affinity to  
old items





# Per-user, per-item models via bilinear random-effects model

### Motivation

- Data measuring k-way interactions pervasive
  - Consider k = 2 for all our discussions
    - E.g. User-Movie, User-content, User-Publisher-Ads,....
  - Power law on both user and item degrees
- Classical Techniques
  - Approximate matrix through a singular value decomposition (SVD)
    - After adjusting for marginal effects (user pop, movie pop,..)
  - Does not work
    - Matrix highly incomplete, severe over-fitting
  - Key issue
    - Regularization of eigenvectors (factors) to avoid overfitting



### Early work on complete matrices

- Tukey's 1-df model (1956)
  - Rank 1 approximation of small nearly complete matrix
- Criss-cross regression (Gabriel, 1978)
- Incomplete matrices: Psychometrics (1-factor model only; small data sets; 1960s)
- Modern day recommender problems
  - Highly incomplete, large, noisy.



### **Latent Factor Models**





### **Factorization – Brief Overview**

• Latent user factors: • Latent movie factors:  $(\alpha_i, \mathbf{u_i}=(u_{i1},...,u_{in}))$   $(\beta_j, \mathbf{v_j}=(v_{j1},...,v_{jn}))$ Interaction

$$E(y_{ij}) = \mu + \alpha_i + \beta_j + u'_i B v_j$$

- (Nn + Mm) \_\_\_\_\_\_ will overfit for moderate values of n,m
- Key technical issue: *Regularization*



### Latent Factor Models: Different Aspects

- Matrix Factorization
  - Factors in Euclidean space
  - Factors on the simplex
- Incorporating features and ratings simultaneously
- Online updates



#### Maximum Margin Matrix Factorization (MMMF)

- Complete matrix by minimizing loss (hinge,squared-error) on observed entries subject to constraints on trace norm
  - Srebro, Rennie, Jakkola (NIPS 2004)
    - Convex, Semi-definite programming (expensive, not scalable)
- Fast MMMF (Rennie & Srebro, ICML, 2005)
  - Constrain the Frobenious norm of left and right eigenvector matrices, not convex but becomes scalable.
- Other variation: Ensemble MMMF (DeCoste, ICML2005)
  - Ensembles of partially trained MMMF (some improvements)



### Matrix Factorization for Netflix prize data

• Minimize the objective function

$$\sum_{ij \in obs} (r_{ij} - u_i^T v_j)^2 + \lambda (\sum_i |u_i|^2 + \sum_j |v_j|^2)$$

- Simon Funk: Stochastic Gradient Descent
- Koren et al (KDD 2007): Alternate Least Squares
   They move to SGD later in the competition



#### Probabilistic Matrix Factorization (Ruslan & Minh, 2008, NIPS)



•Optimization is through Gradient Descent (Iterated conditional modes)

•Other variations like constraining the mean through sigmoid, using "who-ratedwhom"

•Combining with Boltzmann Machines also improved performance

$$r_{ij} \sim N(\mathbf{u}_i^T \mathbf{v}_j, \sigma^2)$$
$$\mathbf{u}_i \sim MVN(\mathbf{0}, a_u I)$$
$$\mathbf{v}_j \sim MVN(\mathbf{0}, a_v I)$$

# Bayesian Probabilistic Matrix Factorization (Ruslan and Minh, ICML 2008)

- Fully Bayesian treatment using an MCMC approach
- Significant improvement
- Interpretation as a fully Bayesian hierarchical model shows why that is the case
  - Failing to incorporate uncertainty leads to bias in estimates
  - Multi-modal posterior, MCMC helps in converging to a better one

	Latent di	atent dimension r		2	5	10	15
		ICI	M9'	736.	9729	.9799	.9802
		MCE	M .97	728 .	9722	.9714	.9715
MCEM also more resistant to over-fitting							
•r	1	3	5	10	20	30	40
/ar-comp: a	0.234	0.123	0.075	0.047	0.028	0.020	0.017



• 1

## Non-parametric Bayesian matrix completion (Zhou et al, SAM, 2010)

• Specify rank probabilistically (automatic rank selection)

$$y_{ij} \sim N(\sum_{k=1}^{r} z_k u_{ik} v_{jk}, \sigma^2)$$

$$z_k \sim Ber(\pi_k)$$

$$\pi_k \sim Beta(a/r, b(r-1)/r)$$

$$z_k \sim Ber(1, a/(a+b(r-1)))$$

$$E(\# \text{Factors}) = ra/(a+b(r-1))$$

### How to incorporate features: Deal with both warm start and cold-start

- Models to predict ratings for new pairs
  - Warm-start: (user, movie) present in the training data with large sample size
  - Cold-start: At least one of (user, movie) new or has small sample size
    - Rough definition, warm-start/cold-start is a continuum.

### Challenges

- Highly incomplete (user, movie) matrix
- Heavy tailed degree distributions for users/movies
  - Large fraction of ratings from small fraction of users/movies
- Handling both warm-start and cold-start effectively in the presence of predictive features



### **Possible approaches**

- Large scale regression based on covariates
  - Does not provide good estimates for heavy users/movies
  - Large number of predictors to estimate interactions
- Collaborative filtering
  - Neighborhood based
  - Factorization
    - · Good for warm-start; cold-start dealt with separately
- Single model that handles cold-start and warm-start
  - Heavy users/movies  $\rightarrow$  User/movie specific model
  - Light users/movies  $\rightarrow$  fallback on regression model
  - Smooth fallback mechanism for good performance





### Add Feature-based Regression into Matrix Factorization

**RLFM: Regression-based Latent Factor Model** 

### **Regression-based Factorization Model (RLFM)**

- Main idea: Flexible prior, predict factors through regressions
- Seamlessly handles cold-start and warm-start
- Modified state equation to incorporate covariates



### **RLFM: Model**

Rating:

user i gives item j

$$\begin{array}{l} y_{ij} \sim N(\mu_{ij}, \sigma^2) \\ y_{ij} \sim Bernoulli(\mu_{ij}) \\ y_{ij} \sim Poisson(\mu_{ij}N_{ij}) \end{array} \end{array}$$

Gaussian Model Logistic Model (for binary rating) Poisson Model (for counts)

$$t(\boldsymbol{\mu}_{ij}) = x_{ij}^{t}b + \boldsymbol{\alpha}_{i} + \boldsymbol{\beta}_{j} + \boldsymbol{u}_{i}^{t}\boldsymbol{v}_{j}$$

- Bias of user *i*:  $\alpha_i = g_0^t x_i + \varepsilon_i^{\alpha}, \quad \varepsilon_i^{\alpha} \sim N(0, \sigma_{\alpha}^2)$
- Popularity of item *j*:  $\beta_j = d_0^t x_j + \varepsilon_j^{\beta}$ ,  $\varepsilon_j^{\beta} \sim N(0, \sigma_{\beta}^2)$
- Factors of user *i*:  $u_i = Gx_i + \varepsilon_i^u$ ,  $\varepsilon_i^u \sim N(0, \sigma_u^2 I)$
- Factors of item *j*:  $v_i = Dx_i + \varepsilon_i^v$ ,  $\varepsilon_i^v \sim N(0, \sigma_v^2 I)$

#### **Could use other classes of regression models**



### Graphical representation of the model



### **Advantages of RLFM**

- Better regularization of factors
  - Covariates "shrink" towards a better centroid
- Cold-start: Fallback regression model (FeatureOnly)

$$y_{ij} \sim N(m_{ij}, \sigma^2)$$
$$m_{ij} = x'_{ij} \boldsymbol{b} + g'_0 w_i + d'_0 z_j + w'_i G' D z_j$$

### **RLFM: Illustration of Shrinkage**

Plot the first factor value for each user (fitted using Yahoo! FP data)





### Induced correlations among observations

Hierarchical random-effects model

$$y_{ij} \sim N(m_{ij}, \sigma^2)$$
$$\downarrow^{x'_{ij}} \boldsymbol{b} + \alpha_i + \beta_j + u'_i v_j$$

Marginal distribution obtained by integrating out random effects

$$\begin{aligned} \alpha_i &= g'_0 w_i + \epsilon_i^{\alpha}, \quad \epsilon_i^{\alpha} \sim N(0, a_{\alpha}) \\ \beta_j &= d'_0 z_j + \epsilon_j^{\beta}, \quad \epsilon_j^{\beta} \sim N(0, a_{\beta}) \\ u_i &= G w_i + \epsilon_i^{u}, \quad \epsilon_i^{u} \sim MVN(\mathbf{0}, A_u) \\ v_j &= D z_j + \epsilon_j^{v}, \quad \epsilon_j^{v} \sim MVN(\mathbf{0}, A_v) \end{aligned}$$



### **Closer look at induced marginal correlations**

$$E(y_{ij}) = x'_{ij}\boldsymbol{b} + g'_0w_i + d'_0z_j + w'_iG'Dz_j$$
$$Var(y_{ij}) = \sigma^2 + a_\alpha + a_\beta + tr(A_uA_v) + z'_jD'A_uDz_j + w'_iG'A_vGw_i$$
$$cov(y_{ij}, y_{ij^*}) = a_\alpha + z'_jD'A_uDz_{j^*}$$
$$cov(y_{ij}, y_{i^*j}) = a_\beta + w'_iG'A_vGw_{i^*}$$

### Model fitting: EM for our class of models

### $\boldsymbol{Y}: \text{ Data}$

 $\Delta$ : Latent variables

## $\Theta$ : hyper-parameters Model: $p(\mathbf{Y}|\mathbf{\Delta}, \Theta)p(\mathbf{\Delta}|\Theta)$

Output needed: Mode:  $max_{\Theta}p(\Theta|Y)$ 

$$p(\boldsymbol{\Delta}|\boldsymbol{Y}) \approx p(\boldsymbol{\Delta}|\boldsymbol{Y}, \hat{\boldsymbol{\Theta}})$$



### The parameters for RLFM

• Latent parameters

$$\Delta = (\{\alpha_i\}, \{\beta_j\}, \{u_i\}, \{v_j\})$$

• Hyper-parameters

$$\Theta = (\mathbf{b}, \mathbf{G}, \mathbf{D}, \mathbf{A}_{u} = \mathbf{a}_{u}\mathbf{I}, \mathbf{A}_{v} = \mathbf{a}_{v}\mathbf{I})$$



$$\begin{split} log(p(\boldsymbol{\Theta}|\boldsymbol{Y})) &= log(p(\boldsymbol{\Theta}, \boldsymbol{\Delta}|\boldsymbol{Y})) - log(p(\boldsymbol{\Delta}|\boldsymbol{\Theta}, \boldsymbol{Y}))\\ log(p(\boldsymbol{\Theta}|\boldsymbol{Y})) &= E_{old}(log(p(\boldsymbol{\Theta}, \boldsymbol{\Delta}|\boldsymbol{Y}))) - E_{old}(log(p(\boldsymbol{\Delta}|\boldsymbol{\Theta}, \boldsymbol{Y})))\\ E_{old}: \text{ Expectation w.r.t. } p(\boldsymbol{\Delta}|\boldsymbol{\Theta}_{old}, \boldsymbol{Y}) \end{split}$$

Second term: Minimized  $\therefore$  at  $\Theta_{old}$ Find new value of  $\Theta$  that increases first term



The EM algorithm

Initialize  $\Theta$ Iterate

E-step :  $E_{old}(log(p(\Theta, \Delta | Y)))$ M-step :  $argmax_{\Theta}E_{old}(log(p(\Theta, \Delta | Y)))$ 



### **Computing the E-step**

- Often hard to compute in closed form
- Stochastic EM (Markov Chain EM; MCEM)
  - Compute expectation by drawing samples from

### $p(\boldsymbol{\Delta}|\boldsymbol{\Theta}_{old}, \boldsymbol{Y})$

- Effective for multi-modal posteriors but more expensive
- Iterated Conditional Modes algorithm (ICM)
  - Faster but biased hyper-parameter estimates

Approximate  $E_{old}(log(p(\Theta, \Delta | Y)))$ by  $log(p(\Theta_{old}, \hat{\Delta} | Y))$ 

 $\hat{\boldsymbol{\Delta}} = argmax_{\boldsymbol{\Delta}} log(p(\boldsymbol{\Theta}_{old}, \boldsymbol{\Delta} | \boldsymbol{Y}))$ 



### Monte Carlo E-step

• Through a vanilla Gibbs sampler (conditionals closed form)

Let 
$$o_{ij} = y_{ij} - \alpha_i - \beta_j - x'_{ij} \boldsymbol{b}$$
  
 $Var[u_i | \text{Rest}] = (A_u^{-1} + \sum_{j \in \mathcal{J}_i} \frac{v_j v'_j}{\sigma_{ij}^2})^{-1}$   
 $E[u_i | \text{Rest}] = Var[u_i | \text{Rest}](A_u^{-1} Gw_i + \sum_{j \in \mathcal{J}_i} \frac{o_{ij} v_j}{\sigma_{ij}^2})$ 

- Other conditionals also Gaussian and closed form
- Conditionals of users (movies) sampled simultaneously
- Small number of samples in early iterations, large numbers in later iterations



### M-step (Why MCEM is better than ICM)

• Update G, optimize

$$(E^*(u_{il}) - Gw_i)'(E^*(u_{il}) - Gw_i)$$

$$\hat{a_{u}} = \frac{\sum_{i=1}^{M} (E^{*}(u_{i}) - \hat{G}w_{i})' (E^{*}(u_{i}) - \hat{G}w_{i}) + \sum_{k=1}^{r} Var^{*}(u_{ikl})}{Mr}$$

Ignored by ICM, underestimates factor variability Factors over-shrunk, posterior not explored well

### **Experiment 1: Better regularization**

- MovieLens-100K, avg RMSE using pre-specified splits
- ZeroMean, RLFM and FeatureOnly (no cold-start issues)
- Covariates:
  - Users : age, gender, zipcode (1<sup>st</sup> digit only)
  - Movies: genres

	RLFM	ZeroMean	FeatureOnly
MovieLens-100K	0.8956	0.9064	1.0968



### **Experiment 2: Better handling of Cold-start**

- MovieLens-1M; EachMovie
- Training-test split based on timestamp
- Same covariates as in Experiment 1.

	MovieLens-1M			EachMovie		
Model	30%	60%	75%	30%	60%	75%
RLFM	0.9742	0.9528	0.9363	1.281	1.214	1.193
ZeroMean	0.9862	0.9614	0.9422	1.260	1.217	1.197
<i>FeatureOnly</i>	1.0923	1.0914	1.0906	1.277	1.272	1.266
FilterBot	0.9821	0.9648	0.9517	1.300	1.225	1.199
MostPopular	0.9831	0.9744	0.9726	1.300	1.227	1.205
Constant Model	1.118	1.123	1.119	1.306	1.302	1.298



### **Experiment 4: Predicting click-rate on articles**

- Goal: Predict click-rate on articles for a user on F1 position
- Article lifetimes short, dynamic updates important
- User covariates:
  - Age, Gender, Geo, Browse behavior
- Article covariates
  - Content Category, keywords
- 2M ratings, 30K users, 4.5 K articles



### **Results on Y! FP data**





### Some other related approaches

- Stern, Herbrich and Graepel, WWW, 2009
  - Similar to RLFM, different parametrization and expectation propagation used to fit the models
- Porteus, Asuncion and Welling, AAAI, 2011
  - Non-parametric approach using a Dirichlet process
- Agarwal, Zhang and Mazumdar, Annals of Applied Statistics, 2011
  - Regression + random effects per user regularized through a Graphical Lasso





### Add Topic Discovery into Matrix Factorization

fLDA: Matrix Factorization through Latent Dirichlet Allocation

### **fLDA:** Introduction

Model the rating y<sub>ij</sub> that user i gives to item j as the user's affinity to the topics that the item has

User *i* 's affinity to topic *k* 

$$y_{ij} = \dots + \sum_{k} s_{ik} \overline{z}_{jk}$$

Pr(item *j* has topic k) estimated by averaging the LDA topic of each word in item *j* 

Old items:  $z_{jk}$ 's are Item latent factors learnt from data with the LDA prior New items:  $z_{jk}$ 's are predicted based on the bag of words in the items

- Unlike regular unsupervised LDA topic modeling, here the LDA topics are learnt in a supervised manner based on past rating data
- fLDA can be thought of as a "multi-task learning" version of the supervised LDA model [Blei'07] for cold-start recommendation



### LDA Topic Modeling (1)

- LDA is effective for unsupervised topic discovery [Blei'03]
  - It models the generating process of a corpus of items (articles)
  - For each topic k, draw a word distribution  $\Phi_k = [\Phi_{k1}, ..., \Phi_{kW}] \sim \text{Dir}(\eta)$
  - For each item *j*, draw a topic distribution  $\theta_j = [\theta_{j1}, \dots, \theta_{jK}] \sim \text{Dir}(\lambda)$
  - For each word, say the *n*th word, in item j,
    - Draw a topic  $z_{jn}$  for that word from  $\theta_j = [\theta_{j1}, ..., \theta_{jK}]$
    - Draw a word  $w_{jn}$  from  $\Phi_k = [\Phi_{k1}, ..., \Phi_{kW}]$  with topic  $k = z_{jn}$





### LDA Topic Modeling (2)

- Model training:

  - EM + Gibbs sampling is a popular method
- Inference for new items
  - Compute the item topic distribution based on the prior parameters and  $\Phi$  estimated in the training phase
- Supervised LDA [Blei'07]
  - Predict a target value for each item based on supervised LDA topics
     Regression weight for topic k
     One regression per user

$$y_j = \sum_k s_k \overline{z}_{jk} \quad \text{vs.}$$

s.  $y_{ij} = \dots + \sum_{k} S_{ik} z_{jk}$ Same set of topics across different regressions

Target value of item *j* 

Pr(item *j* has topic *k*) estimated by averaging the topic of each word in item *j* 



### **fLDA: Model**

Rating:

user i gives item j  $y_{ij} \sim N(\mu_{ij}, \sigma^2)$   $y_{ij} \sim Bernoulli(\mu_{ij})$  $y_{ij} \sim Poisson(\mu_{ij}N_{ij})$  Gaussian Model Logistic Model (for binary rating) Poisson Model (for counts)

$$t(\boldsymbol{\mu}_{ij}) = x_{ij}^{t} b + \boldsymbol{\alpha}_{i} + \boldsymbol{\beta}_{j} + \sum_{k} s_{ik} \overline{z}_{jk}$$

- Bias of user *i*:
- Popularity of item j:
- Topic affinity of user *i*:  $s_i = Hx_i + \mathcal{E}_i^s$ ,  $\mathcal{E}_i^s \sim N(0, \sigma_s^2 I)$
- Pr(item *j* has topic *k*):
- $\overline{z}_{jk} = \sum_{n} 1(z_{jn} = k) / (\# \text{ words in item } j)$ The LDA topic of the *n*th word in item *j*

 $\boldsymbol{\alpha}_{i} = \boldsymbol{g}_{0}^{t} \boldsymbol{x}_{i} + \boldsymbol{\varepsilon}_{i}^{\alpha}, \quad \boldsymbol{\varepsilon}_{i}^{\alpha} \sim N(0, \boldsymbol{\sigma}_{\alpha}^{2})$ 

 $\beta_i = d_0^t x_i + \varepsilon_i^{\beta}, \quad \varepsilon_i^{\beta} \sim N(0, \sigma_{\beta}^2)$ 

Observed words:

$$W_{jn} \sim LDA(\lambda, \eta, z_{jn})$$
  
The *n*th word in item *j*

Y

### **Model Fitting**

- Given:
  - Features  $X = \{x_i, x_j, x_{ij}\}$
  - Observed ratings  $y = \{y_{ij}\}$  and words  $w = \{w_{jn}\}$
- Estimate:
  - Parameters:  $\Theta = [b, g_0, d_0, H, \sigma^2, a_{\alpha}, a_{\beta}, A_s, \lambda, \eta]$ 
    - Regression weights and prior parameters
  - Latent factors:  $\Delta = \{\alpha_i, \beta_j, s_i\}$  and  $z = \{z_{jn}\}$ 
    - User factors, item factors and per-word topic assignment
- Empirical Bayes approach:
  - Maximum likelihood estimate of the parameters:

$$\hat{\Theta} = \arg\max_{\Theta} \Pr[y, w \mid \Theta] = \arg\max_{\Theta} \int \Pr[y, w, \Delta, z \mid \Theta] d\Delta dz$$

- The posterior distribution of the factors:

$$\Pr[\Delta, z \mid y, \hat{\Theta}]$$



### The EM Algorithm

- Iterate through the E and M steps until convergence
  - Let  $\hat{\Theta}^{(n)}$  be the current estimate
  - E-step: Compute

$$f(\Theta) = E_{(\Delta, z \mid y, w, \hat{\Theta}^n)}[\log \Pr(y, w, \Delta, z \mid \Theta)]$$

- · The expectation is not in closed form
- We draw Gibbs samples and compute the Monte Carlo mean
- M-step: Find

$$\hat{\Theta}^{(n+1)} = \arg\max_{\Theta} f(\Theta)$$

It consists of solving a number of regression and optimization problems



### **Supervised Topic Assignment**



### fLDA: Experimental Results (Movie)

- Task: Predict the rating that a user would give a movie
- Training/test split:
  - Sort observations by time
  - First 75%  $\rightarrow$  Training data
  - Last 25%  $\rightarrow$  Test data
- Item warm-start scenario
  - Only 2% new items in test data

Model	Test RMSE
RLFM	0.9363
fLDA	0.9381
Factor-Only	0.9422
FilterBot	0.9517
unsup-LDA	0.9520
MostPopular	0.9726
Feature-Only	1.0906
Constant	1.1190

fLDA is as strong as the best method

It does not reduce the performance in warm-start scenarios



### fLDA: Experimental Results (Yahoo! Buzz)

- Task: Predict whether a user would buzz-up an article
- Severe item cold-start



### **Experimental Results: Buzzing Topics**

#### 3/4 topics are interpretable; 1/2 are similar to unsupervised topics

Top Terms (after stemming)	Торіс
bush, tortur, interrog, terror, administr, CIA, offici,	CIA interrogation
suspect, releas, investig, georg, memo, al	
mexico, flu, pirat, swine, drug, ship, somali, border,	Swine flu
mexican, hostag, offici, somalia, captain	
NFL, player, team, suleman, game, nadya, star, high,	NFL games
octuplet, nadya_suleman, michael, week	
court, gai, marriag, suprem, right, judg, rule, sex,	Gay marriage
pope, supreme_court, appeal, ban, legal, allow	
palin, republican, parti, obama, limbaugh, sarah, rush,	Sarah Palin
gop, presid, sarah_palin, sai, gov, alaska	
idol, american, night, star, look, michel, win, dress,	American idol
susan, danc, judg, boyl, michelle_obama	
economi, recess, job, percent, econom, bank, expect,	Recession
rate, jobless, year, unemploy, month	
north, korea, china, north_korea, launch, nuclear,	North Korea issues
rocket, missil, south, said, russia	



### **fLDA Summary**

- fLDA is a useful model for cold-start item recommendation
- It also provides interpretable recommendations for users
  - User's preference to interpretable LDA topics
- Future directions:
  - Investigate Gibbs sampling chains and the convergence properties of the EM algorithm
  - Apply fLDA to other multi-task prediction problems
    - fLDA can be used as a tool to generate supervised features (topics) from text data



### Summary

- Regularizing factors through covariates effective
- Regression based factor model that regularizes better and deals with both cold-start and warm-start in a single framework in a seamless way looks attractive
- Fitting method scalable; Gibbs sampling for users and movies can be done in parallel. Regressions in M-step can be done with any off-the-shelf scalable linear regression routine
- Distributed computing on Hadoop: Multiple models and average across partitions



### Hierarchical smoothing

Advertising application, Agarwal et al. KDD 10

• Product of states for each node pair



Spike and Slab prior

 $(S_z, E_z, \lambda_z)$ 

 $\pi(\phi;a,P) = P1(\phi=1) + (1-P) \operatorname{Gamma}(\phi;1,1/a)$ 

- Known to encourage parsimonious solutions
  - Several cell states have no corrections
- Not used before for multi-hierarchy models, only in regression
- We choose P = .5 (and choose "a" by cross-validation)
  - a psuedo number of successes

