

THE MOTIVATING PROBLEM

- **Non-parametric Bayesian** construction relating multivariate covariates to “manifold-valued” responses.
- Extension of locally-defined models on Riemannian manifolds to “global” models.

WHAT'S NEW?

- DPMM for manifold-valued data e.g., sphere [Straub et al., AISTAT 2015], SPD [Cherian et al. CVPR 2011]
- Inspired by DP-GLM in Euclidean space [Hannah et al., JMLR 2011] and Multivariate general linear model (MGLM) on manifolds [Kim et al. CVPR 2014]
- **This paper : DP-MGLM for Riemannian manifolds (SPD(n)) w.r.t. the GL-invariant metric**
- **Main technical highlight:** New Hamiltonian Monte Carlo (HMC) algorithm for parameters on manifolds and tangent spaces $(B, V) \in \mathcal{M} \times T_B\mathcal{M}^d$.
- New distribution proposed for $(B, V) \in \mathcal{M} \times T_B\mathcal{M}^d$.
- A direct by-product is “clustering based on nonlinear correlation” between covariates and response.

EUCLIDEAN AND RIEMANNIAN MGLMS: BASIC OPERATION

Operation	Subtraction	Addition	Distance	Mean	Covariance
Euclidean	$\vec{x}_i \vec{x}_j = x_j - x_i$	$x_i + \vec{x}_j \vec{x}_k$	$\ \vec{x}_i \vec{x}_j\ $	$\sum_{i=1}^n \vec{x}_i = 0$	$\mathbb{E}[(x_i - \bar{x})(x_i - \bar{x})^T]$
Riemannian	$\vec{x}_i \vec{x}_j = \text{Log}(x_i, x_j)$	$\text{Exp}(x_i, \vec{x}_j \vec{x}_k)$	$\ \text{Log}(x_i, x_j)\ _x$	$\sum_{i=1}^n \text{Log}(\bar{x}, x_i) = 0$	$\mathbb{E}[\text{Log}(\bar{x}, x_i)\text{Log}(\bar{x}, x_i)^T]$

Euclidean GLM

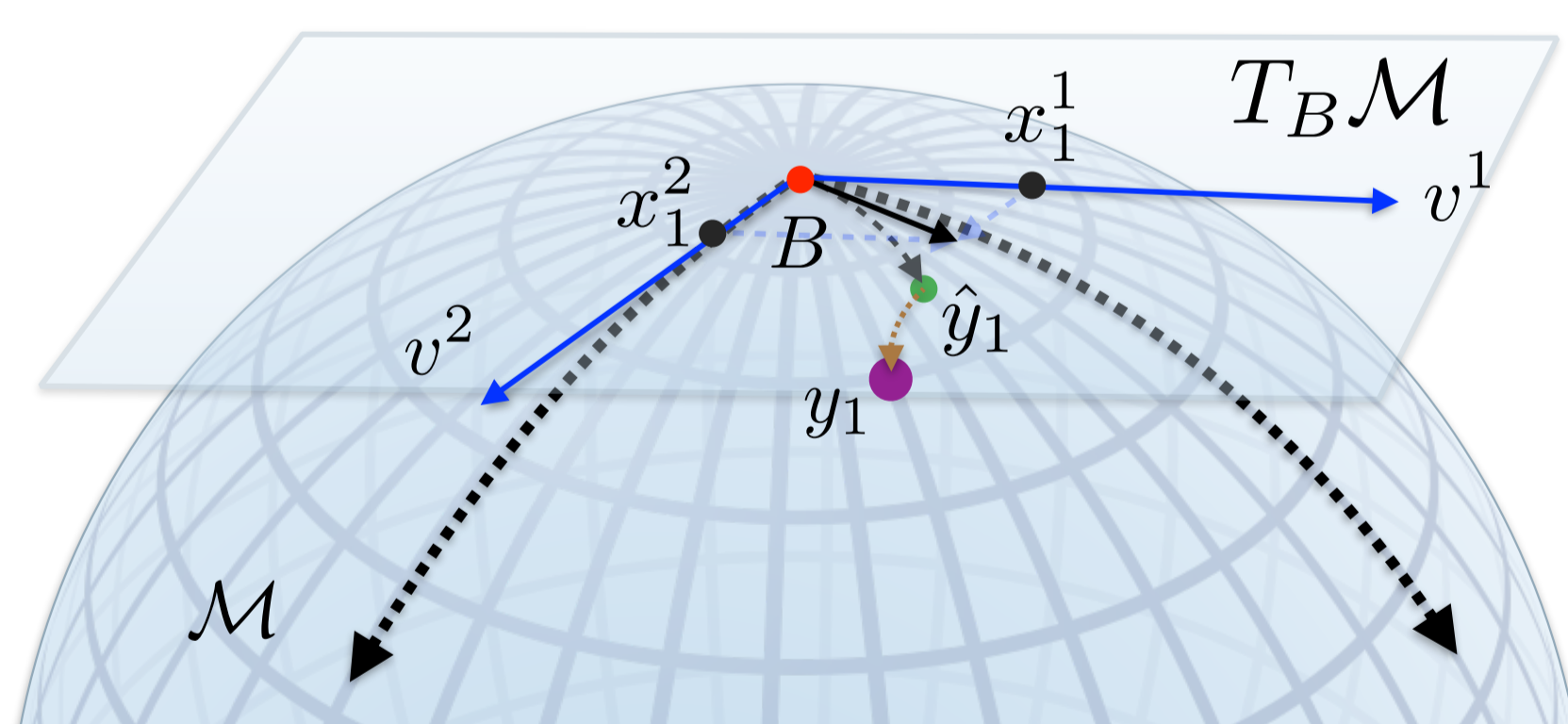
$$y = \alpha + \beta^1 x^1 + \beta^2 x^2 + \dots + \beta^d x^d + \epsilon,$$

where $x \in \mathbf{R}^d$, $y \in \mathbf{R}^d$ and $\beta \in \mathbf{R}^d$.

Riemannian MGLM

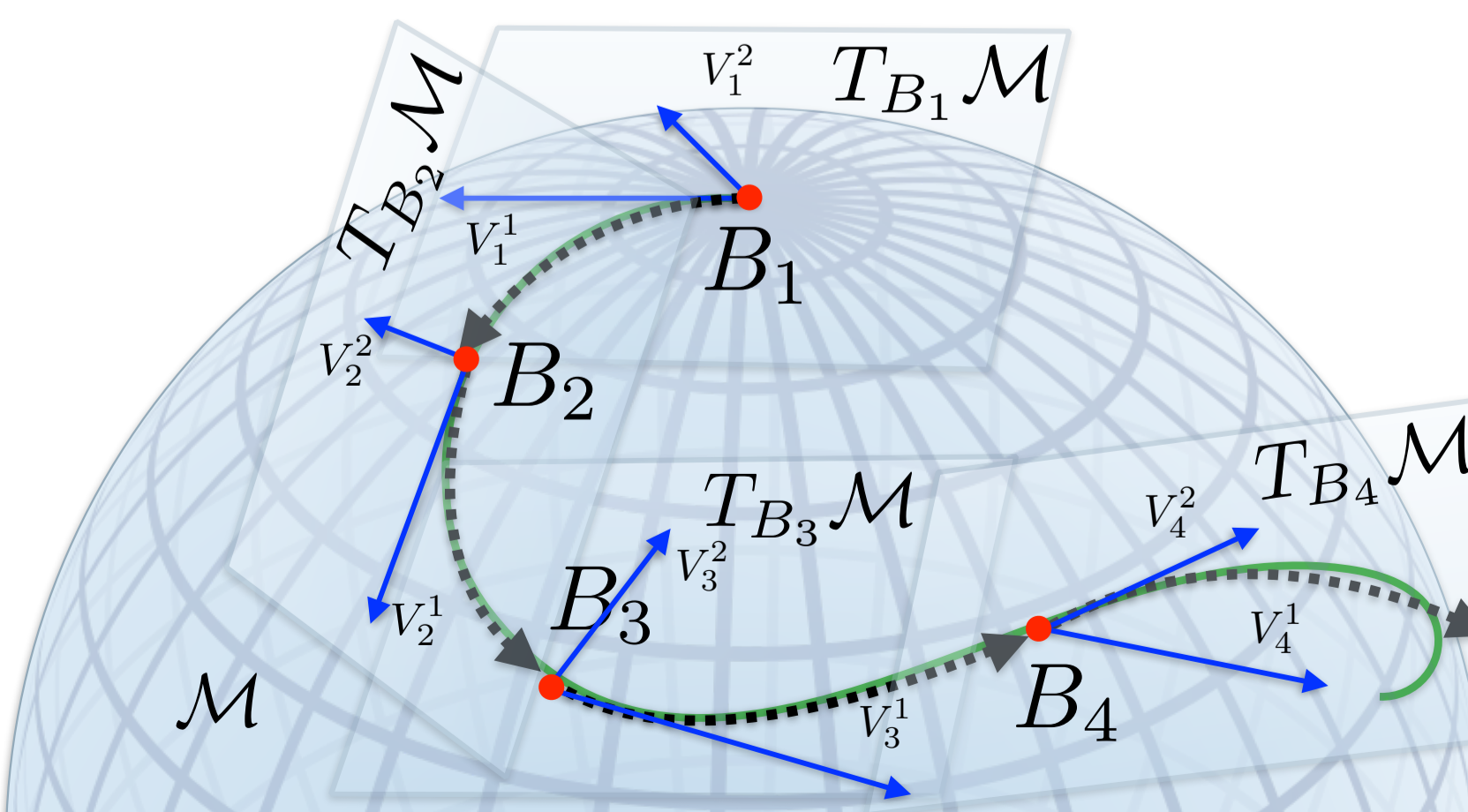
$$y = \text{Exp}\left(\text{Exp}\left(B, V^1 x^1 + V^2 x^2 + \dots + V^d x^d\right), \epsilon\right),$$

where $x \in \mathbf{R}^d$, $y \in \mathcal{M}$, $B \in \mathcal{M}$, and $V^j \in T_B\mathcal{M}$.



DP-MGLM ON MANIFOLDS

$$y = \text{Exp}\left(\text{Exp}\left(B_i, V_i^1 x_i^1 + V_i^2 x_i^2 + \dots + V_i^d x_i^d\right), \epsilon\right) \quad (1)$$



DP-MGLM ON MANIFOLDS

$$G \sim DP(G_0, \nu)$$

$$\theta_i = (\theta_{x_i}, \theta_{y_i}) | G \sim G$$

$$x_i | \theta_{x_i} \sim f_x(\theta_{x_i})$$

$$y_i | x_i, \theta_{y_i} \sim \text{MGLM}(x_i, \theta_{y_i})$$

$$G \sim DP(G_0, \nu)$$

$$\theta_i := (\underbrace{\mu_{x_i}, \sigma_{x_i}^2}_{\theta_{x_i}}, \underbrace{B_i, V_i}_{\theta_{y_i}}) | G \sim G,$$

$$x_i | \theta_{x_i} \sim \mathcal{N}(\theta_{x_i}, \sigma_{x_i}^2),$$

$$y_i | x_i, \theta_{y_i}, \sigma_{y_i}^2 \sim \mathcal{N}_{\text{SPD}}(\hat{y}_i, \sigma_{y_i}^2),$$

where $\hat{y}_i = \text{Exp}(B_i, V_i x)$.

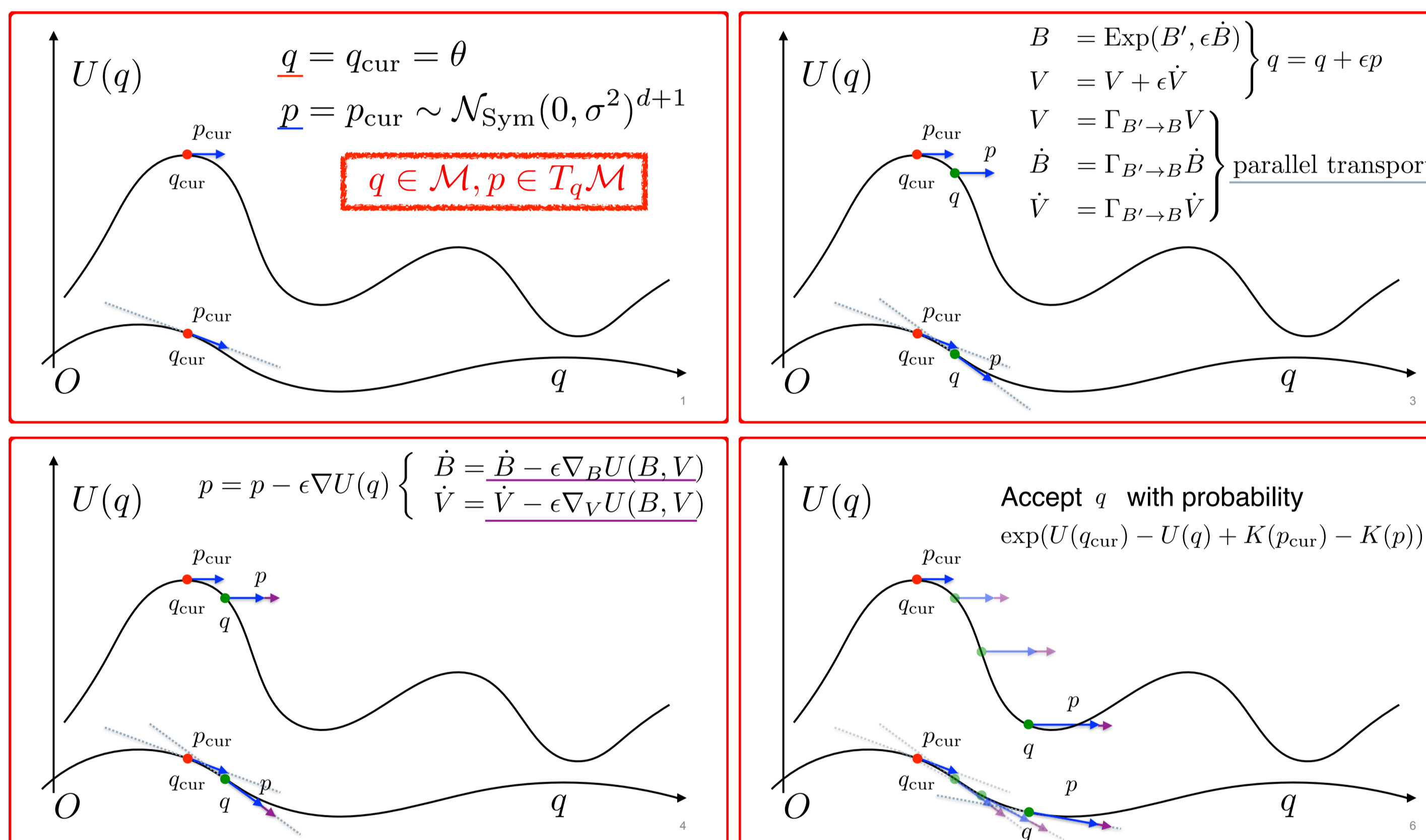
HAMILTONIAN MONTE CARLO SAMPLING (HMC)

$$H(q, p) = U(q) + K(p)$$

$$U(q) := -\log f(q) \text{ and } K(p) := \frac{1}{2} p^T M^{-1} p$$

where q is parameters to sample, p is momentum vector, $M \succ 0$, $M^T = M$ is “mass matrix” [Duane et al., “Hybrid monte carlo.”, Physics letters B, 1987].

HMC ALGORITHM FOR DP-MGLM ON RIEMANNIAN MANIFOLDS



- 1: Input: $(B_{cur}, V_{cur}) \in \mathcal{M} \times T_B\mathcal{M}^n$, Leapfrog parameters $\epsilon \in \mathbf{R}_{++}$, $L \in \mathbf{Z}_{++}$
- 2: Output: $(B_{next}, V_{next}) \in \mathcal{M} \times T_B\mathcal{M}^n$
- 3: Sample $(\hat{B}_{cur}, \hat{V}_{cur}) \in T_B\mathcal{M} \times T_B\mathcal{M}^n$ from independent normal distribution w.r.t. Riemannian metric.
- 4: Initialize $(B, V, \hat{B}, \hat{V}) \leftarrow (B_{cur}, V_{cur}, \hat{B}_{cur}, \hat{V}_{cur})$
- 5: $\hat{B} \leftarrow \hat{B} - \frac{\epsilon}{2} \nabla_B U(B, V)$ and $\hat{V} \leftarrow \hat{V} - \frac{\epsilon}{2} \nabla_V U(B, V)$
- 6: **for** $i \in \{1, \dots, L\}$ **do**
- 7: $B' \leftarrow B$, $B \leftarrow \text{Exp}(B, \epsilon \hat{B})$, $V \leftarrow V + \epsilon \hat{V}$
- 8: $(V, \hat{B}, \hat{V}) \leftarrow (\Gamma_{B' \rightarrow B} V, \Gamma_{B' \rightarrow B} \hat{B}, \Gamma_{B' \rightarrow B} \hat{V})$
- 9: /* Parallel transport */
- 10: **if** i is not L **then**
- 11: $B \leftarrow B - \epsilon \nabla_B U(B, V)$ and $\hat{V} \leftarrow \hat{V} - \epsilon \nabla_V U(B, V)$
- 12: **end if**
- 13: **end for**
- 14: $\hat{B} \leftarrow \hat{B} - \frac{\epsilon}{2} \nabla_B U(B, V)$ and $\hat{V} \leftarrow \hat{V} - \frac{\epsilon}{2} \nabla_V U(B, V)$
- 15: Accept (\hat{B}, \hat{V}) with probability
- 16: $\min[1, \exp(H(\hat{B}_{cur}, V_{cur}, B_{cur}, V_{cur}) - H(\hat{B}, \hat{V}, B, V))]$

LEMMA

Let $(B, V) \in \text{SPD}(n) \times \text{Sym}(n)$ be a sample drawn using (2), then V is Normally distributed w.r.t. a GL-invariant metric at the tangent space $T_B\mathcal{M}$ at B . For each B , the probability density function of V is proportional to $\exp(-\frac{1}{2} \|V\|_B^2)$ at $T_B\mathcal{M}$, when $\mu_V = 0$.

$$B | \mu_B, \sigma_B^2 \sim \mathcal{N}_{\text{SPD}}(B | \mu_B, \sigma_B^2), V | \mu_V, B \sim \mathcal{N}_{\text{Sym}}(V | \mu_V, B) \quad (2)$$

SYNTHETIC EXPERIMENTS: RIEMANNIAN DP-MGLM VS. MGLM

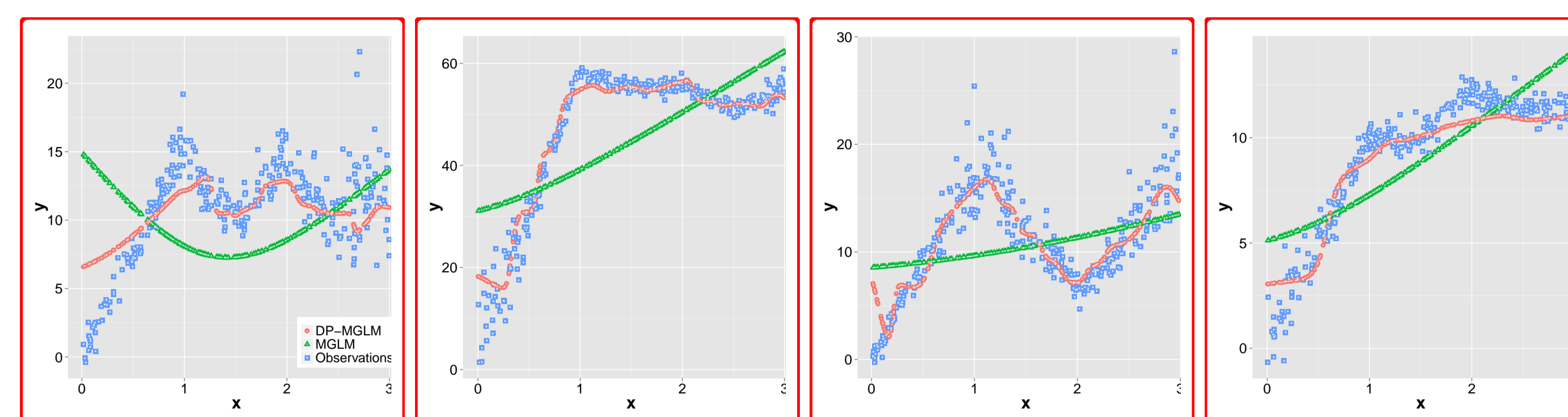


Figure : The figure shows the models fitted in response $y \in \mathcal{M}$ versus covariate $x \in \mathbf{R}$. The prediction of DP-MGLM is shown using a single sample from the posterior, $\theta^{(i)}$. To visualize the response variable $Y \in \text{SPD}(3)$, we project the variables onto the axis obtained by PGA (y -axis). The x -axis is the covariate $x \in \mathbf{R}$. Red and blue correspond to our predictions and the measurements respectively.

SYNTHETIC EXPERIMENTS: CLUSTERING BY NONLINEAR CORRELATION

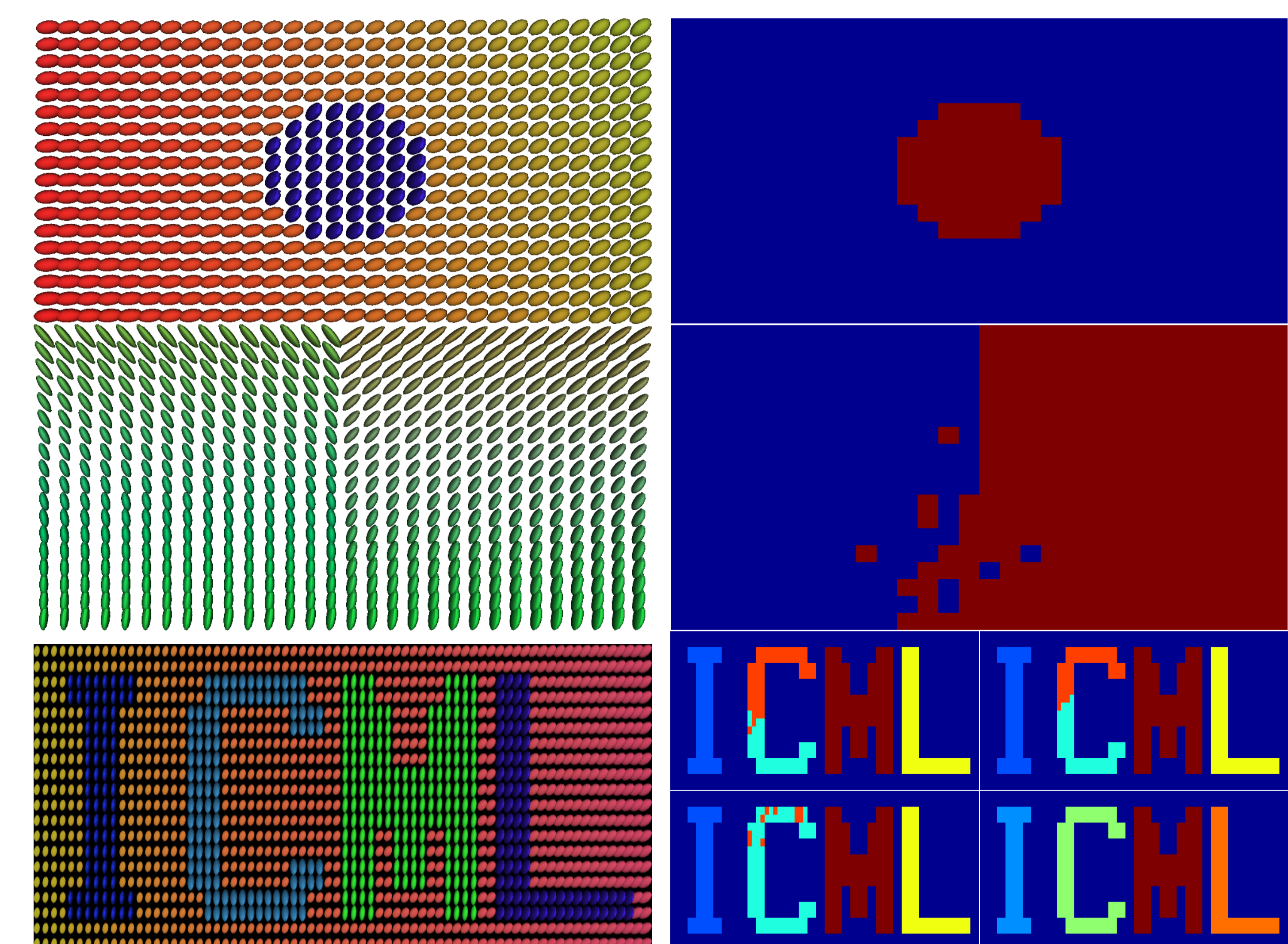


Figure : (Column 1) shows SPD(3) response variables in each voxel; the corresponding covariates are the grid positions (horizontal, vertical coordinates). (Column 2) shows a clustering result. A clustering is shown based on the posterior samples $\theta^{(i)}$.

SYNTHETIC EXPERIMENTS: CLUSTERING EFFECT

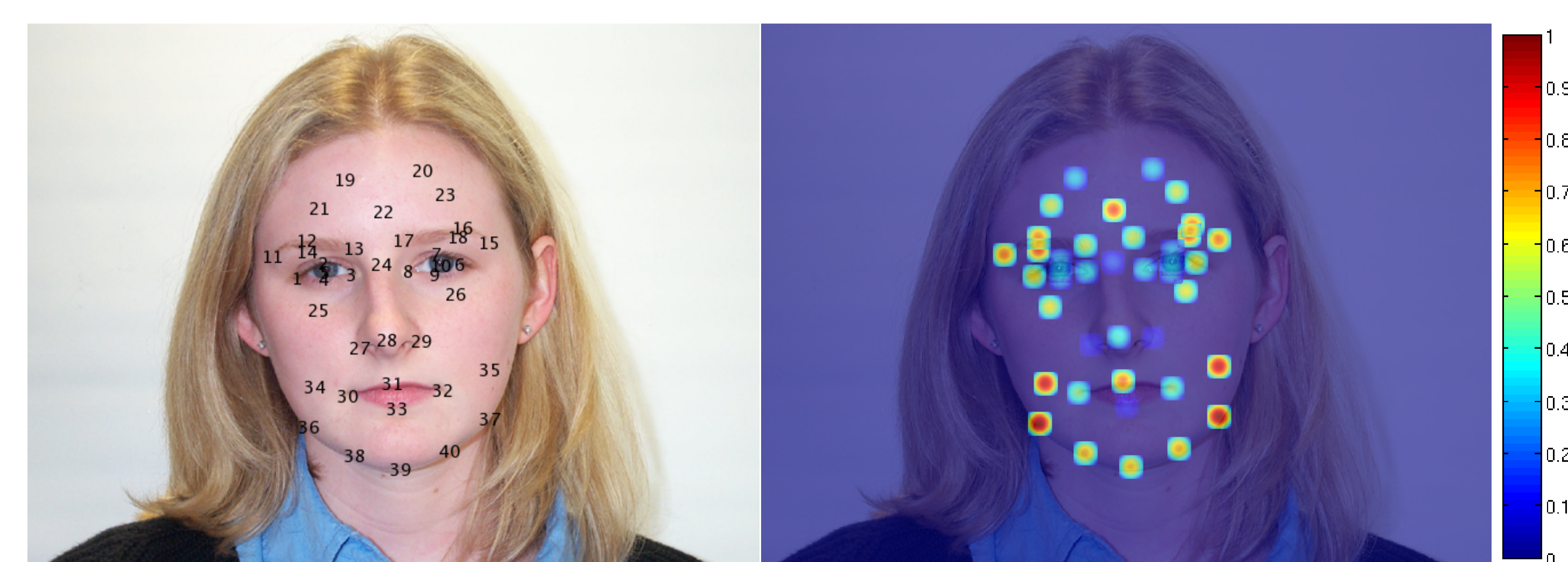


Figure : Left image shows 40 landmarks (indexed by numbers) on an example image. Right image of the bottom row shows correlation magnitude of the landmark's variation with age as a heat map. (Best viewed in color.)