

# Statistical Decision Theory

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Consider a parameter  $\theta \in \Theta$ . We observe data  $x$  sampled from the distribution parametrized by  $\theta$ . Let  $\hat{\theta} \equiv \hat{\theta}(x)$  be an estimator of  $\theta$  based on data  $x$ . We are going to compare different estimators.

Let a *loss function*  $L(\theta, \hat{\theta}) : \Theta \times \Theta \mapsto \mathbb{R}_+$  be defined. For example,

$$L(\theta, \hat{\theta}) = (\theta - \hat{\theta})^2 \quad (1)$$

$$L(\theta, \hat{\theta}) = \begin{cases} 0 & \theta = \hat{\theta} \\ 1 & \theta \neq \hat{\theta} \end{cases} \quad (2)$$

$$L(\theta, \hat{\theta}) = \int p(x; \theta) \log \left( \frac{p(x; \theta)}{p(x; \hat{\theta})} \right) dx \quad (3)$$

The *risk*  $R(\theta, \hat{\theta})$  is the average loss, averaged over training sets sampled from the true  $\theta$ :

$$R(\theta, \hat{\theta}) = \mathbb{E}_\theta[L(\theta, \hat{\theta}(x))] = \int p(x; \theta) L(\theta, \hat{\theta}(x)) dx \quad (4)$$

Recall that  $\mathbb{E}_\theta$  means the expectation over  $x$  drawn from the distribution with fixed parameter  $\theta$ , not the expectation over different  $\theta$ .

**Example 1** Let  $X \sim N(\theta, 1)$ . Let  $\hat{\theta}_1 = X$  and  $\hat{\theta}_2 = 3.14$ . Assume squared error loss. Then  $R(\theta, \hat{\theta}_1) = 1$  (hint: variance),  $R(\theta, \hat{\theta}_2) = \mathbb{E}_\theta(\theta - 3.14)^2 = (\theta - 3.14)^2$ . (hint: no  $X$  here) Over the whole range of possible  $\theta \in \mathbb{R}$ , neither estimator consistently dominates.

**Example 2** Let  $X_1, \dots, X_n \sim \text{Bernoulli}(\theta)$ . Consider squared error loss. Let  $\hat{\theta}_1 = \frac{\sum X_i}{n}$ , the sample mean. Let  $\hat{\theta}_2 = \frac{\alpha + \sum X_i}{\alpha + \beta + n}$  which is the “smoothed” estimate, i.e., the posterior mean under a Beta( $\alpha, \beta$ ) prior. Let  $\hat{\theta}_3 = X_1$ , the first sample. Then,  $R(\theta, \hat{\theta}_1) = \mathbb{V}(\frac{\sum X_i}{n}) = \frac{\theta(1-\theta)}{n}$  and  $R(\theta, \hat{\theta}_3) = \mathbb{V}(X_1) = \theta(1-\theta)$ . So  $\hat{\theta}_3$  is out as a learning algorithm. But what about  $\hat{\theta}_2$ ?

$$R(\theta, \hat{\theta}_2) = \mathbb{E}_\theta(\theta - \hat{\theta}_2)^2 \quad (5)$$

$$= \mathbb{V}_\theta(\hat{\theta}_2) + (\text{bias}(\hat{\theta}_2))^2 \quad (6)$$

$$= \frac{n\theta(1-\theta)}{(n+\alpha+\beta)^2} + \left( \frac{n\theta+\alpha}{n+\alpha+\beta} - \theta \right)^2 \quad (7)$$

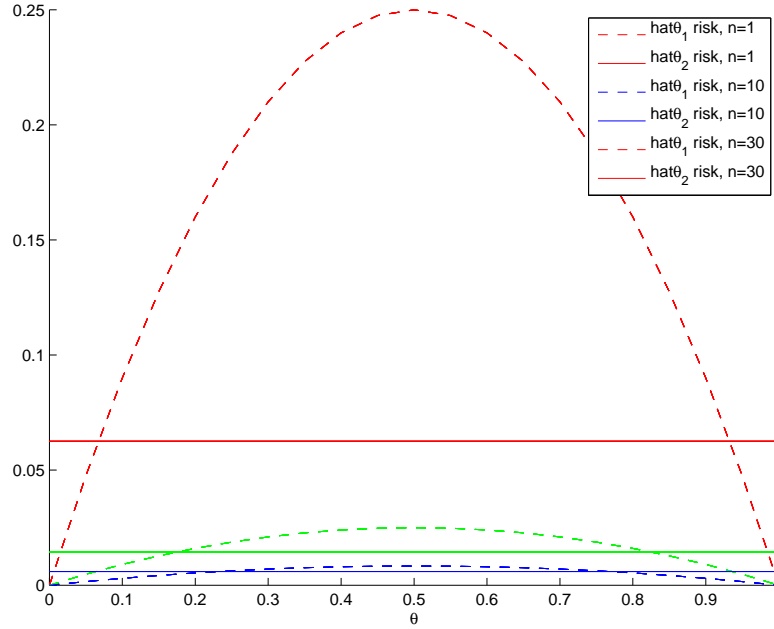
It is not difficult to show that one can make  $\theta$  disappear from the risk (i.e., task insensitivity) by setting

$$\alpha = \beta = \sqrt{n}/2$$

with

$$R(\theta, \hat{\theta}_2) = \frac{1}{4(\sqrt{n}+1)^2}$$

It turns out this particular choice of  $\alpha, \beta$  leads to a so-called minimax estimator  $\hat{\theta}_2$ , as we will show later. However, there is no dominance between  $\hat{\theta}_1$  and  $\hat{\theta}_2$  as the figure below shows:



The *maximum risk* is

$$R^{max}(\hat{\theta}) = \sup_{\theta} R(\theta, \hat{\theta}) \quad (8)$$

The *Bayes risk* under prior  $f(\theta)$  is

$$R_f^{Bayes}(\hat{\theta}) = \int R(\theta, \hat{\theta}) f(\theta) d\theta. \quad (9)$$

Accordingly, two different criteria to define “the best estimator” (or the best learning algorithm) is the *Bayes rule* and the *minimax rule*, respectively. An estimator  $\hat{\theta}^{Bayes}$  is a Bayes rule with respect to the prior  $f$  if

$$\hat{\theta}^{Bayes} = \arg \inf_{\hat{\theta}} \int R(\theta, \hat{\theta}) f(\theta) d\theta, \quad (10)$$

where the infimum is over all estimators  $\hat{\theta}$ . An estimator  $\hat{\theta}^{minimax}$  that minimizes the maximum risk is a *minimax rule*:

$$\hat{\theta}^{minimax} = \arg \inf_{\hat{\theta}} \sup_{\theta} R(\theta, \hat{\theta}), \quad (11)$$

where again the infimum is over all estimators  $\hat{\theta}$ .

We list the following theorems without proof. For details see AoS p.197.

**Theorem 1** Let  $f(\theta)$  be a prior,  $x$  a sample, and  $f(\theta | x)$  the corresponding posterior. If  $L(\theta, \hat{\theta}) = (\theta - \hat{\theta})^2$  then the Bayes rule is the posterior mean:

$$\hat{\theta}^{Bayes}(x) = \int \theta f(\theta | x) d\theta = \mathbb{E}(\theta | X = x). \quad (12)$$

If  $L(\theta, \hat{\theta}) = |\theta - \hat{\theta}|$  then the Bayes rule is the posterior median. If  $L(\theta, \hat{\theta})$  is zero-one loss then the Bayes rule is the posterior mode.

**Theorem 2** Suppose that  $\hat{\theta}$  is the Bayes rule with respect to some prior  $f$ . Suppose further that  $\hat{\theta}$  has a constant risk:  $R(\theta, \hat{\theta}) = c$  for all  $\theta \in \Theta$ . Then  $\hat{\theta}$  is minimax.

**Example 3** In example 2 we made the choice  $\alpha = \beta = \sqrt{n}/2$  so that the risk  $R(\theta, \hat{\theta}_2) = \frac{1}{4(\sqrt{n}+1)^2}$  is a constant. Also,  $\hat{\theta}_2$  is the posterior mean and hence by Theorem 1 is a Bayes rule under the prior  $\text{Beta}(\sqrt{n}/2, \sqrt{n}/2)$ . Putting them together, by Theorem 2  $\hat{\theta}_2$  is minimax.