

---

# Supplementary Material for “Human Memory Search as Initial-Visit Emitting Random Walk”

---

**Kwang-Sung Jun\***, **Xiaojin Zhu<sup>†</sup>**, **Timothy Rogers<sup>‡</sup>**

\*Wisconsin Institute for Discovery, <sup>†</sup>Department of Computer Sciences, <sup>‡</sup>Department of Psychology  
University of Wisconsin-Madison

kjun@discovery.wisc.edu, jerryzhu@cs.wisc.edu, ttrogers@wisc.edu

**Zhuoran Yang**

Department of Mathematical Sciences  
Tsinghua University  
yzr11@mails.tsinghua.edu.cn

**Ming Yuan**

Department of Statistics  
University of Wisconsin-Madison  
myuan@stat.wisc.edu

## A Derivative of (6) w.r.t. $\beta$

Give a censored list  $\mathbf{a}$ , define a mapping  $\sigma$  that maps a state to its position in  $\mathbf{a}$ ; that is,  $\sigma(a_i) = i$ . Let  $\mathbf{N}^{(k)} = (\mathbf{I} - \mathbf{Q}^{(k)})^{-1}$ . Hereafter, we drop the superscript  $(k)$  from  $\mathbf{Q}$ ,  $\mathbf{R}$  and  $\mathbf{N}$  when it's clear from the context.

Using  $\partial(A^{-1})_{k\ell}/\partial A_{ij} = -(A^{-1})_{ki}(A^{-1})_{j\ell}$ , the following identity becomes useful:

$$\begin{aligned} \frac{\partial N_{k\ell}}{\partial Q_{ij}} &= \frac{\partial((\mathbf{I} - \mathbf{Q})^{-1})_{k\ell}}{\partial Q_{ij}} \\ &= \sum_{c,d} \frac{\partial((\mathbf{I} - \mathbf{Q})^{-1})_{k\ell}}{\partial(\mathbf{I} - \mathbf{Q})_{cd}} \frac{\partial(\mathbf{I} - \mathbf{Q})_{cd}}{\partial Q_{ij}} \\ &= \sum_{c,d} ((\mathbf{I} - \mathbf{Q})^{-1})_{kc} ((\mathbf{I} - \mathbf{Q})^{-1})_{d\ell} \mathbb{1}_{\{c=i, d=j\}} \\ &= ((\mathbf{I} - \mathbf{Q})^{-1})_{ki} ((\mathbf{I} - \mathbf{Q})^{-1})_{j\ell} \\ &= N_{ki} N_{j\ell}. \end{aligned}$$

The derivative of  $\mathbf{P}$  w.r.t.  $\beta$  is given as follows:

$$\begin{aligned} \frac{\partial P_{rc}}{\partial \beta_{ij}} &= \mathbb{1}\{r = i\} \left( \frac{\mathbb{1}\{j = c\} e^{\beta_{ic}} (\sum_{\ell=1}^n e^{\beta_{i\ell}}) - e^{\beta_{ic}} e^{\beta_{ij}}}{(\sum_{\ell=1}^n e^{\beta_{i\ell}})^2} \right) \\ &= \mathbb{1}\{r = i\} (-P_{ic} P_{ij} + \mathbb{1}\{j = c\} P_{ic}). \end{aligned}$$

The derivative of  $\log \mathbb{P}(a_{k+1} | a_{1:k})$  with respect to  $\beta$  is

$$\begin{aligned} \frac{\partial \log \mathbb{P}(a_{k+1} | a_{1:k})}{\partial \beta_{ij}} &= \mathbb{P}(a_{k+1} | a_{1:k})^{-1} \sum_{\ell=1}^k \frac{\partial(N_{k\ell} R_{\ell 1})}{\partial \beta_{ij}} \\ &= \mathbb{P}(a_{k+1} | a_{1:k})^{-1} \left( \sum_{\ell=1}^k \frac{\partial N_{k\ell}}{\partial \beta_{ij}} R_{\ell 1} + N_{k\ell} \frac{\partial R_{\ell 1}}{\partial \beta_{ij}} \right) \end{aligned}$$

We need to compute  $\frac{\partial N_{k\ell}}{\partial \beta_{ij}}$ :

$$\begin{aligned}\frac{\partial N_{k\ell}}{\partial \beta_{ij}} &= \sum_{c,d=1}^k \frac{\partial((\mathbf{I} - \mathbf{Q})^{-1})_{k\ell}}{\partial(\mathbf{I} - \mathbf{Q})_{cd}} \frac{\partial(\mathbf{I} - \mathbf{Q})_{cd}}{\partial \beta_{ij}} \\ &= \sum_{c,d=1}^k (-1)N_{kc}N_{d\ell} \cdot (-1)\mathbb{1}_{\{a_c=i\}}(-P_{ia_d}P_{ij} + \mathbb{1}_{\{a_d=j\}}P_{ia_d}) \\ &= \mathbb{1}_{\{\sigma(i) \leq k\}}N_{k\sigma(i)} \sum_{d=1}^k N_{d\ell}(-P_{ia_d}P_{ij} + \mathbb{1}_{\{a_d=j\}}P_{ia_d}),\end{aligned}$$

where  $\sigma(i) \leq k$  means item  $i$  appeared among the first  $k$  items in the censored list  $\mathbf{a}$ .

Then,

$$\begin{aligned}\sum_{\ell=1}^k \frac{\partial N_{k\ell}}{\partial \beta_{ij}} R_{\ell 1} &= \mathbb{1}_{\{\sigma(i) \leq k\}}N_{k\sigma(i)} \sum_{\ell,d=1}^k N_{d\ell}(-P_{ia_d}P_{ij} + \mathbb{1}_{\{a_d=j\}}P_{ia_d})R_{\ell 1} \\ &= \mathbb{1}_{\{\sigma(i) \leq k\}}N_{k\sigma(i)} \left( -P_{ij} \sum_{d=1}^k P_{ia_d} \sum_{\ell=1}^k N_{d\ell} R_{\ell 1} + \sum_{d=1}^k \mathbb{1}_{\{a_d=j\}} P_{ia_d} \sum_{\ell=1}^k N_{d\ell} R_{\ell 1} \right) \\ &= \mathbb{1}_{\{\sigma(i) \leq k\}}N_{k\sigma(i)} \left( -P_{ij}(\mathbf{QNR})_{\sigma(i)1} + \mathbb{1}_{\{\sigma(j) \leq k\}}P_{ij}(\mathbf{NR})_{\sigma(j)1} \right)\end{aligned}$$

and

$$\begin{aligned}\sum_{\ell=1}^k N_{k\ell} \frac{\partial R_{\ell 1}}{\partial \beta_{ij}} &= \sum_{\ell=1}^k N_{k\ell} \mathbb{1}_{\{\ell=\sigma(i)\}} \left( -P_{ia_{k+1}}P_{ij} + \mathbb{1}_{\{a_{k+1}=j\}}P_{ia_{k+1}} \right) \\ &= \mathbb{1}_{\{\sigma(i) \leq k\}}N_{k\sigma(i)} \left( -P_{ia_{k+1}}P_{ij} + \mathbb{1}_{\{a_{k+1}=j\}}P_{ia_{k+1}} \right).\end{aligned}$$

Putting everything together,

$$\begin{aligned}&\frac{\partial \log \mathbb{P}(a_{k+1} \mid a_{1:k})}{\partial \beta_{ij}} \\ &= \frac{\mathbb{1}_{\{\sigma(i) \leq k\}}N_{k\sigma(i)}}{\mathbb{P}(a_{k+1} \mid a_{1:k})} \left( -P_{ij}(\mathbf{QNR})_{\sigma(i)1} + \mathbb{1}_{\{\sigma(j) \leq k\}}P_{ij}(\mathbf{NR})_{\sigma(j)1} \right. \\ &\quad \left. - P_{ia_{k+1}}P_{ij} + \mathbb{1}_{\{a_{k+1}=j\}}P_{ia_{k+1}} \right) \\ &= \frac{\mathbb{1}_{\{\sigma(i) \leq k\}}N_{k\sigma(i)}P_{ij}}{\mathbb{P}(a_{k+1} \mid a_{1:k})} \left( -(\mathbf{QNR})_{\sigma(i)1} + \mathbb{1}_{\{\sigma(j) \leq k\}}(\mathbf{NR})_{\sigma(j)1} P_{ia_{k+1}} \left( \frac{\mathbb{1}_{\{a_{k+1}=j\}}}{P_{ij}} - 1 \right) \right)\end{aligned}$$

for all  $i \neq j$ .

## B The Proof of Theorem 2

We first claim that (i) there must be a recurrent state  $i$  in a censored list where  $i \in W_k$  for some  $k$ . Then, it suffices to show that given (i) is true, (ii) recurrent states outside  $W_k$  cannot appear, (iii) every states in  $W_k$  must appear, and (iv) a transient state cannot appear after a recurrent state.

(i): suppose there is no recurrent state in a censored list  $\mathbf{a} = (a_{1:M})$ . Then, every state  $a_i, i \in [M]$ , is a transient state. Since the underlying random walk runs indefinitely in finite state space, there must be a state  $a_j, j \in [M]$ , that is visited infinitely many times. This contradicts the fact that  $a_j$  is a transient state.

Suppose a recurrent state  $i \in W_k$  was visited. Then,

(ii): the random walk cannot escape  $W_k$  since  $W_k$  is closed.

(iii): the random walk will reach to every state in  $W_k$  in finite time since  $W_k$  is finite and irreducible.

(iv): the same reason as (iii).

## C The Proof of Theorem 4

It suffices to show that  $\mathbb{P}(a_{k+1} \mid a_{1:k}; \mathbf{P}) = \mathbb{P}(a_{k+1} \mid a_{1:k}; \mathbf{P}')$ , where  $\mathbf{a} = (a_1, \dots, a_M)$  and  $k \leq M - 1$ . Define submatrices  $(\mathbf{Q}, \mathbf{R})$  and  $(\mathbf{Q}', \mathbf{R}')$  from  $\mathbf{P}$  and  $\mathbf{P}'$ , respectively, as in (2). Note that  $\mathbf{Q}' = \text{diag}(q_{1:k}) + (\mathbf{I} - \text{diag}(q_{1:k}))\mathbf{Q}$  and  $\mathbf{R}' = (\mathbf{I} - \text{diag}(q_{1:k}))\mathbf{R}$ .

$$\begin{aligned} \mathbb{P}(a_{k+1} \mid a_{1:k}; \mathbf{P}') &= (\mathbf{I} - \text{diag}(q_{1:k}) - (\mathbf{I} - \text{diag}(q_{1:k}))\mathbf{Q})^{-1} (\mathbf{I} - \text{diag}(q_{1:k}))\mathbf{R} \\ &= (\mathbf{I} - \mathbf{Q})^{-1} (\mathbf{I} - \text{diag}(q_{1:k}))^{-1} (\mathbf{I} - \text{diag}(q_{1:k}))\mathbf{R} \\ &= \mathbb{P}(a_{k+1} \mid a_{1:k}; \mathbf{P}) \end{aligned}$$

## D The Proof of Theorem 5

Suppose  $(\pi, \mathbf{P}) \neq (\pi', \mathbf{P}')$ . We show that there exists a censored list  $\mathbf{a}$  such that  $\mathbb{P}(\mathbf{a}; \pi, \mathbf{P}) \neq \mathbb{P}(\mathbf{a}; \pi', \mathbf{P}')$ .

**Case 1:**  $\pi \neq \pi'$ .

It follows that  $\pi_i \neq \pi'_i$  for some  $i$ . Note that the marginal probability of observing  $i$  as the first item in a censored list is  $\sum_{\mathbf{a} \in \mathcal{D}: a_1=i} \mathbb{P}(\mathbf{a}; \pi, \mathbf{P}) = \pi_i$ . Then,

$$\sum_{\mathbf{a} \in \mathcal{D}: a_1=i} \mathbb{P}(\mathbf{a}; \pi, \mathbf{P}) = \pi_i \neq \pi'_i = \sum_{\mathbf{a} \in \mathcal{D}: a_1=i} \mathbb{P}(\mathbf{a}; \pi', \mathbf{P}')$$

which implies that there exists a censored list  $\mathbf{a}$  for which  $\mathbb{P}(\mathbf{a}; \pi, \mathbf{P}) \neq \mathbb{P}(\mathbf{a}; \pi', \mathbf{P}')$ .

**Case 2:**  $\pi = \pi'$  but  $\mathbf{P} \neq \mathbf{P}'$ .

It follows that  $P_{ij} \neq P'_{ij}$  for some  $i$  and  $j$ . Then, we compute the marginal probability of observing  $(i, j)$  as the first two items in a censored list, which results in

$$\sum_{\mathbf{a} \in \mathcal{D}: \substack{a_1=i, \\ a_2=j}} \mathbb{P}(\mathbf{a}; \pi, \mathbf{P}) = \pi_i P_{ij} \neq \pi'_i P'_{ij} = \sum_{\mathbf{a} \in \mathcal{D}: \substack{a_1=i, \\ a_2=j}} \mathbb{P}(\mathbf{a}; \pi', \mathbf{P}')$$

Then, there exists a censored list  $\mathbf{a}$  for which  $\mathbb{P}(\mathbf{a}; \pi, \mathbf{P}) \neq \mathbb{P}(\mathbf{a}; \pi', \mathbf{P}')$ .

## E Results Required for Theorem 6

Throughout, assume  $\boldsymbol{\theta} = (\boldsymbol{\pi}^\top, \mathbf{P}_1, \dots, \mathbf{P}_n)^\top$ . Let  $\text{supp}(\boldsymbol{\theta})$  be the set of nonzero dimensions of  $\boldsymbol{\theta}$ :  $\text{supp}(\boldsymbol{\theta}) = \{i \mid \theta_i > 0\}$ . Lemma 1 shows conditions on which  $\mathcal{Q}^*(\boldsymbol{\theta})$  and  $\widehat{\mathcal{Q}}_m(\boldsymbol{\theta})$  are above  $-\infty$ .

**Lemma 1.** *Assume A1. Then,*

$$\text{supp}(\boldsymbol{\theta}) \supseteq \text{supp}(\boldsymbol{\theta}^*) \iff \mathcal{Q}^*(\boldsymbol{\theta}) > -\infty \quad (7)$$

$$\text{supp}(\boldsymbol{\theta}) \supseteq \text{supp}(\boldsymbol{\theta}^*) \implies \widehat{\mathcal{Q}}_m(\boldsymbol{\theta}) > -\infty, \forall m. \quad (8)$$

*Proof.* Define two vectors of probabilities w.r.t.  $\boldsymbol{\theta}$  and  $\boldsymbol{\theta}^*$ :  $\mathbf{q} = [q_{\mathbf{a}} = \mathbb{P}(\mathbf{a}; \boldsymbol{\theta})]_{\mathbf{a} \in \mathcal{D}}$  and  $\mathbf{q}^* = [q_{\mathbf{a}}^* = \mathbb{P}(\mathbf{a}; \boldsymbol{\theta}^*)]_{\mathbf{a} \in \mathcal{D}}$ . Note that

$$\text{supp}(\mathbf{q}) \supseteq \text{supp}(\mathbf{q}^*) \iff \mathcal{Q}^*(\boldsymbol{\theta}) > -\infty$$

by the definition of  $\mathcal{Q}^*(\boldsymbol{\theta})$ . Thus, for (7), it suffices to show that

$$\text{supp}(\boldsymbol{\theta}) \supseteq \text{supp}(\boldsymbol{\theta}^*) \iff \text{supp}(\mathbf{q}) \supseteq \text{supp}(\mathbf{q}^*).$$

( $\implies$ ) The LHS implies that the directed graph induced by  $\boldsymbol{\theta}$  includes the graph induced by  $\boldsymbol{\theta}^*$ ; a path that is possible w.r.t.  $\boldsymbol{\theta}^*$  is also possible w.r.t.  $\boldsymbol{\theta}$ . Recall that a list is generated by a random walk. Let  $\mathbf{a} \in \text{supp}(\mathbf{q}^*)$ . There exists a random walk under  $\boldsymbol{\theta}^*$  that generates  $\mathbf{a}$ . Then, the same random walk is also possible under  $\boldsymbol{\theta}$ , which implies  $\mathbf{a} \in \text{supp}(\mathbf{q})$ .

( $\impliedby$ ) Suppose the LHS is false. Then, there exists  $(i, j)$  s.t.  $P_{ij} = 0$  and  $P_{ij}^* > 0$ . Consider a list  $\mathbf{a}$  such that it has nonzero probability w.r.t.  $\boldsymbol{\theta}^*$  (that is,  $q_{\mathbf{a}}^* > 0$ ), and its first two items are  $i$  then  $j$ . Since  $P_{ij} = 0$ ,  $q_{\mathbf{a}} = 0$ . However, the RHS implies that  $q_{\mathbf{a}} > 0$  since  $q_{\mathbf{a}}^* > 0$ : a contradiction.

For (8),

$$\text{supp}(\theta) \supseteq \text{supp}(\theta^*) \implies \mathcal{Q}^*(\theta) > -\infty \implies \widehat{\mathcal{Q}}_m(\theta) > -\infty, \forall m,$$

where the last implication is due to the fact that a censored list  $\mathbf{a}^{(i)}$  that appears in  $\widehat{\mathcal{Q}}_m(\theta)$  is generated by  $\theta^*$ , so the term  $\log \mathbb{P}(\mathbf{a}^{(i)}; \theta)$  also appears in  $\mathcal{Q}^*(\theta)$ . □

**Lemma 2.** *Assume A1. Then,  $\theta^*$  is the unique maximizer of  $\mathcal{Q}^*(\theta)$ .*

*Proof.* If  $\theta$  satisfies  $\text{supp}(\theta) \not\supseteq \text{supp}(\theta^*)$ , then  $\mathcal{Q}^*(\theta) = -\infty$  by Lemma 1, so such  $\theta$  cannot be a maximizer. Thus, it is safe to restrict our attention to  $\theta$ 's whose support include that of  $\theta^*$ :  $\text{supp}(\theta) \supseteq \text{supp}(\theta^*)$ .

Recall the definition of  $\mathcal{Q}^*(\theta)$ :

$$\mathcal{Q}^*(\theta) = \sum_{\mathbf{a} \in \mathcal{D}} \mathbb{P}(\mathbf{a}; \theta^*) \log \mathbb{P}(\mathbf{a}; \theta) \propto -\text{KL}(\theta^* || \theta),$$

where  $\text{KL}(\theta^* || \theta)$  is well defined since  $\text{supp}(\theta) \supseteq \text{supp}(\theta^*)$ . Due to the identifiability of the model (Theorem 5) and the unique minimizer property of the KL-divergence,  $\theta^*$  is the unique maximizer. □

We denote by  $\text{decomp}(\theta) = \{T, W_1, \dots, W_K\}$  the decomposition induced by  $\theta$  as in Theorem 1.

**Lemma 3.**  *$\text{supp}(\widehat{\theta}_m) \supseteq \text{supp}(\theta^*)$  for large enough  $m$ . Furthermore,  $\text{decomp}(\widehat{\theta}_m) = \text{decomp}(\theta^*)$  for large enough  $m$ .*

*Proof.* Note that due to the strong law of large numbers, a list  $\mathbf{a}$  is valid in the true model  $\theta^*$  must appear in  $D_m$  for large enough  $m$ . Since the number of censored lists that can be generated by  $\theta^*$  is finite, one observes every valid censored list in the true model  $\theta^*$ ; that is, there exists  $m'$  such that

$$m \geq m' \implies \{\mathbf{a} \mid \mathbf{a} \in D_m\} = \{\mathbf{a} \mid \mathbb{P}(\mathbf{a}; \theta^*) > 0\}.$$

For the first statement, assume that  $m \geq m'$ . Since we observe every valid list in  $\theta^*$ , by the definition of  $\widehat{\mathcal{Q}}_m(\theta)$ , the following holds true:

$$\forall \theta \in \Theta, \widehat{\mathcal{Q}}_m(\theta) > -\infty \iff \mathcal{Q}^*(\theta) > -\infty.$$

Then, using Lemma 1,

$$\widehat{\mathcal{Q}}_m(\widehat{\theta}_m) > -\infty \implies \mathcal{Q}^*(\widehat{\theta}_m) > -\infty \implies \text{supp}(\widehat{\theta}_m) \supseteq \text{supp}(\theta^*).$$

For the second statement, assume  $m \geq m'$ . Let  $\text{decomp}(\widehat{\theta}_m) = \{\widehat{T}, \widehat{W}_1, \dots, \widehat{W}_{\widehat{K}}\}$  and  $\text{decomp}(\theta^*) = \{T^*, W_1^*, \dots, W_{K^*}^*\}$ . Furthermore, define  $\widehat{\tau}(i)$  to be the index of the closed irreducible set in  $\text{decomp}(\widehat{\theta}_m)$  to which  $i$  belongs, and define  $\tau^*(i)$  similarly.

Suppose that the data  $D_m$  contains every valid list in  $\theta^*$ , but  $\text{decomp}(\widehat{\theta}_m) \neq \text{decomp}(\theta^*)$ . There are four cases. In each case, we show that there exists a list that is valid in  $\theta^*$  but not in  $\widehat{\theta}_m$ , which means that the log likelihood of  $\widehat{\theta}_m$  is  $-\infty$ . This is a contradiction in that  $\widehat{\theta}_m$  is the MLE.

**Case 1 :**  $\exists s_1$  s.t.  $s_1$  is transient in  $\widehat{\theta}$  but recurrent in  $\theta^*$ .

Let  $W_k^*$  be the closed irreducible set to which  $s_1$  belongs and  $L = |W_k^*|$ . Use  $\theta^*$  to start a random walk from  $s_1$  and generate a censored list  $\mathbf{a}$ , which consists of all states in  $W_k^*$ :  $\mathbf{a} = (s_1, s_2, \dots, s_L)$ . If  $\mathbf{a}$  is invalid in  $\widehat{\theta}_m$ , we have a contradiction. If not,  $s_L$  must be recurrent in  $\widehat{\theta}_m$  by Theorem 2. Use  $\theta^*$  to generate a censored list  $\mathbf{a}'$  that starts from  $s_L$ . Then,  $s_1$  must appear after  $s_L$  in  $\mathbf{a}'$ . However, this is impossible in  $\widehat{\theta}_m$  since  $s_1$  is transient and  $s_L$  is recurrent: a contradiction.

**Case 2 :**  $\exists t$  s.t.  $t$  is transient in  $\theta^*$  but recurrent in  $\widehat{\theta}_m$ .

For brevity, assume that  $t$  is the only transient state in  $\theta^*$ ; this can be easily relaxed. Use  $\theta^*$  to generate a censored list that starts with  $t$ , say  $\mathbf{a} = (t, s_1, \dots, s_L)$ . By Theorem 2,  $\{s_{1:L}\}$  is a closed irreducible set in  $\theta^*$ . Define  $\mathbf{a}' = (s_{1:L})$ , which is also valid in  $\theta^*$ . Now,  $\mathbf{a}$  may or may not be valid in  $\hat{\theta}_m$ . Assume that  $\mathbf{a}$  is valid in  $\hat{\theta}_m$  since otherwise we have a contradiction. Then, in  $\hat{\theta}_m$ ,  $\{t, s_{1:L}\}$  must be a closed irreducible set since  $t$  is recurrent. Then,  $\mathbf{a}' = (s_{1:L})$  is invalid in  $\hat{\theta}_m$  since  $t$  must be visited as well: a contradiction.

**Case 3.**  $\exists(i, j)$  s.t.  $\hat{\tau}(i) = \hat{\tau}(j)$ , but  $\tau^*(i) \neq \tau^*(j)$ .

Start a random walk from the state  $i$  w.r.t.  $\theta^*$  and generate a censored list  $\mathbf{a}$ . By Theorem 2, the censored list  $\mathbf{a}$  does not contain  $j$ . In  $\hat{\theta}_m$ , however, a censored list starting from  $i$  must also output  $j$  since  $i$  and  $j$  are in the same closed irreducible set. Thus,  $\mathbf{a}$  is invalid in  $\hat{\theta}_m$ : a contradiction.

**Case 4.**  $\exists(i, j)$  s.t.  $\tau^*(i) = \tau^*(j)$ , but  $\hat{\tau}(i) \neq \hat{\tau}(j)$ .

Start a random walk from the state  $i$  w.r.t.  $\theta^*$  and generate a censored list  $\mathbf{a}$ . By Theorem 2, the censored list  $\mathbf{a}'$  must also contain  $j$ . In  $\hat{\theta}_m$ , however, a censored list starting from  $i$  cannot output  $j$  since  $j$  is in a different closed irreducible set. Thus,  $\mathbf{a}'$  is invalid in  $\hat{\theta}_m$ : a contradiction.  $\square$

**Lemma 4.** Assume A1. Let  $\{\hat{\theta}_{m_j}\}$  be a convergent subsequence of  $\{\hat{\theta}_m\}$  and  $\theta'$  be its limit point:  $\theta' = \lim_{j \rightarrow \infty} \hat{\theta}_{m_j}$ . Then,  $\lim_{j \rightarrow \infty} \mathbb{P}(\mathbf{a}; \hat{\theta}_{m_j}) = \mathbb{P}(\mathbf{a}; \theta')$  for all  $\mathbf{a}$  that is valid in  $\theta^*$ .

*Proof.* There are exactly two case-by-case operators which causes the likelihood function to be discontinuous. The operators appear in (3) and (1), which respectively rely on the following conditions w.r.t. a list  $\mathbf{a} = (a_{1:M})$ :

$$(\mathbf{I} - \mathbf{Q}^{(k)})^{-1} \text{ exists, } \forall k \in [M-1] \quad (9)$$

$$\mathbb{P}(s \mid a_{1:M}; \theta) = 0, \forall s \in S \setminus \{a_{1:M}\}. \quad (10)$$

**Step 1:** claim that  $\forall \theta \in \Theta$ ,

$$\text{supp}(\theta) \supseteq \text{supp}(\theta^*) \text{ and } \text{decomp}(\theta) = \text{decomp}(\theta^*) \implies \forall \mathbf{a} \text{ valid in } \theta^* \text{ (9) and (10)}$$

To show (9), suppose it is false for some  $k \in [M-1]$  and some censored list  $\mathbf{a} = (a_{1:M})$  valid in  $\theta^*$ . The nonexistence of  $(\mathbf{I} - \mathbf{Q}^{(k)})^{-1}$  implies that there is no path from  $a_k$  to a state that is outside of  $\{a_{1:k}\}$  whereas there is such a path w.r.t.  $\theta^*$ . This contradicts  $\text{supp}(\theta) \supseteq \text{supp}(\theta^*)$ .

To show (10), consider a censored list  $\mathbf{a} = (a_{1:M})$  that is valid in  $\theta^*$ . By Theorem 2, the last state  $a_M$  must be a recurrent state in a closed irreducible set  $W$  w.r.t.  $\theta^*$ . Since  $\theta$  has the same decomposition as  $\theta^*$  and every state in  $W$  must be present in  $\mathbf{a}$ , no other state can appear after  $a_M$ . This implies (10).

Define

$$\Theta' = \{\theta \in \Theta \mid \|\theta - \theta'\|_\infty < \min_i \theta'_i, \text{decomp}(\theta) = \text{decomp}(\theta')\}.$$

**Step 2:** claim that  $\mathbb{P}(\mathbf{a}; \theta)$  is a continuous function of  $\theta$  in the subspace  $\Theta'$ ,  $\forall \mathbf{a}$  valid in  $\theta^*$ .

Note that  $\forall \theta \in \Theta'$ ,

$$\begin{aligned} \text{supp}(\theta) &\supseteq \text{supp}(\theta') \supseteq \text{supp}(\theta^*) \\ \text{decomp}(\theta) &= \text{decomp}(\theta') = \text{decomp}(\theta^*), \end{aligned}$$

where the first subset relation is due to the  $\infty$ -norm in the definition of  $\Theta'$ , the second subset relation and the last equality is due to Lemma 3 and  $\theta' = \lim_{j \rightarrow \infty} \hat{\theta}_{m_j}$ .

This implies, together with step 1, that  $\forall \theta \in \Theta'$ , (9) and (10) are satisfied, which effectively gets rid of the case-by-case operators in  $\Theta'$ . This concludes the claim.

**Step 3:**  $\lim_{j \rightarrow \infty} \mathbb{P}(\mathbf{a}; \hat{\theta}_{m_j}) = \mathbb{P}(\mathbf{a}; \theta')$  for all  $\mathbf{a}$  that is valid in  $\theta^*$ .

Since  $\widehat{\boldsymbol{\theta}}_{m_j} \rightarrow \boldsymbol{\theta}'$ , there exists  $J$  such that

$$j \geq J \implies \|\widehat{\boldsymbol{\theta}}_{m_j} - \boldsymbol{\theta}'\|_\infty < \min_i \theta'_i.$$

Thus, after  $J$ , the sequence enters the subspace  $\Theta'$  in which  $\mathbb{P}(\mathbf{a}; \boldsymbol{\theta})$  is continuous  $\forall \mathbf{a}$  valid in  $\boldsymbol{\theta}^*$ , which concludes the claim.  $\square$

**Lemma 5.** *Assume A1. Let  $\{\widehat{\boldsymbol{\theta}}_{m_j}\}$  be a convergent subsequence of  $\{\widehat{\boldsymbol{\theta}}_m\}$  and  $\boldsymbol{\theta}'$  be its limit point:  $\boldsymbol{\theta}' = \lim_{j \rightarrow \infty} \widehat{\boldsymbol{\theta}}_{m_j}$ . Then,  $\mathcal{Q}^*(\boldsymbol{\theta}') > -\infty$ .*

*Proof.* Suppose not:  $\mathcal{Q}^*(\boldsymbol{\theta}') = -\infty$ . Then, there exists a list  $\mathbf{a}'$  that is valid in  $\boldsymbol{\theta}^*$  whose likelihood w.r.t.  $\boldsymbol{\theta}'$  converges to 0:

$$\exists \mathbf{a}' \text{ s.t. } \mathbb{P}(\mathbf{a}'; \boldsymbol{\theta}^*) > 0 \text{ and } \mathbb{P}(\mathbf{a}'; \boldsymbol{\theta}') = 0,$$

By Lemma 4,  $\mathbb{P}(\mathbf{a}'; \boldsymbol{\theta}') = 0$  implies that  $\lim_{j \rightarrow \infty} \mathbb{P}(\mathbf{a}'; \widehat{\boldsymbol{\theta}}_{m_j}) = 0$ .

Let  $0 < \epsilon < \mathbb{P}(\mathbf{a}'; \boldsymbol{\theta}^*)$ . Denote by  $\#\{\mathbf{a}'\}$  the number of occurrences of the list  $\mathbf{a}'$  in  $\{\mathbf{a}^{(1)}, \dots, \mathbf{a}^{(m_j)}\}$ . Then, the following statements hold:

$$\exists J_1 \text{ s.t. } j > J_1 \implies \left| \frac{\#\{\mathbf{a}'\}}{m_j} - \mathbb{P}(\mathbf{a}'; \boldsymbol{\theta}^*) \right| < \epsilon \quad (11)$$

$$\exists J_2 \text{ s.t. } j < J_2 \implies \left| \widehat{\mathcal{Q}}_{m_j}(\boldsymbol{\theta}^*) - \mathcal{Q}^*(\boldsymbol{\theta}^*) \right| < \epsilon \quad (12)$$

$$\exists J_3 \text{ s.t. } j > J_3 \implies \log \mathbb{P}(\mathbf{a}'; \widehat{\boldsymbol{\theta}}_{m_j}) < \frac{\mathcal{Q}^*(\boldsymbol{\theta}^*) - \epsilon}{\mathbb{P}(\mathbf{a}'; \boldsymbol{\theta}^*) - \epsilon}. \quad (13)$$

The first two statements are due to the law of large numbers, and the last statement is due to the convergence of  $\mathbb{P}(\mathbf{a}'; \widehat{\boldsymbol{\theta}}_{m_j})$  to 0. Note that  $\widehat{\mathcal{Q}}_{m_j}(\boldsymbol{\theta}^*) \leq \widehat{\mathcal{Q}}_{m_j}(\widehat{\boldsymbol{\theta}}_{m_j})$  since  $\widehat{\boldsymbol{\theta}}_{m_j}$  is the maximizer of the function  $\widehat{\mathcal{Q}}_{m_j}(\boldsymbol{\theta})$ . Then, if  $j > \max\{J_1, J_2, J_3\}$ ,

$$\begin{aligned} \mathcal{Q}^*(\boldsymbol{\theta}^*) - \epsilon &\leq \widehat{\mathcal{Q}}_{m_j}(\boldsymbol{\theta}^*) \\ &\leq \widehat{\mathcal{Q}}_{m_j}(\widehat{\boldsymbol{\theta}}_{m_j}) \\ &= \left( \sum_{\mathbf{a} \neq \mathbf{a}'} \frac{\#\{\mathbf{a}\}}{m_j} \log \mathbb{P}(\mathbf{a}; \widehat{\boldsymbol{\theta}}_{m_j}) \right) + \frac{\#\{\mathbf{a}'\}}{m_j} \log \mathbb{P}(\mathbf{a}'; \widehat{\boldsymbol{\theta}}_{m_j}) \\ &< (\mathbb{P}(\mathbf{a}'; \boldsymbol{\theta}^*) - \epsilon) \log \mathbb{P}(\mathbf{a}'; \widehat{\boldsymbol{\theta}}_{m_j}) \\ &< \mathcal{Q}^*(\boldsymbol{\theta}^*) - \epsilon, \end{aligned}$$

where the last inequality is due to (13). This is a contradiction.  $\square$

**Lemma 6.** *Assume A1. Let  $\{\widehat{\boldsymbol{\theta}}_{m_j}\}$  be a convergent subsequence of  $\{\widehat{\boldsymbol{\theta}}_m\}$ . Let  $\boldsymbol{\theta}' = \lim_{j \rightarrow \infty} \widehat{\boldsymbol{\theta}}_{m_j}$ . Then,  $\lim_{j \rightarrow \infty} \widehat{\mathcal{Q}}_{m_j}(\widehat{\boldsymbol{\theta}}_{m_j}) = \mathcal{Q}^*(\boldsymbol{\theta}')$ .*

*Proof.* The idea is that we can have a compact ball around the limit point  $\boldsymbol{\theta}'$  and show that the log likelihood  $\widehat{\mathcal{Q}}_{m_j}(\boldsymbol{\theta})$  converges uniformly on the ball. Then, after the sequence  $\widehat{\boldsymbol{\theta}}_{m_j}$  gets in the ball, we can use the uniform convergence of the log likelihood.

Let  $B_{\boldsymbol{\theta}'}(r) = \{\boldsymbol{\theta} \in \Theta \mid \|\boldsymbol{\theta} - \boldsymbol{\theta}'\|_\infty \leq r\}$  be an  $\infty$ -norm ball around  $\boldsymbol{\theta}'$ . Choose  $\epsilon' < \min_{i \in \text{supp}(\boldsymbol{\theta}')} \theta'_i$ . We claim that

$$\forall \boldsymbol{\theta} \in B_{\boldsymbol{\theta}'}(\epsilon'), \mathcal{Q}^*(\boldsymbol{\theta}) > -\infty \text{ and } \widehat{\mathcal{Q}}_m(\boldsymbol{\theta}) > -\infty, \forall m. \quad (14)$$

Let  $\boldsymbol{\theta} \in B_{\boldsymbol{\theta}'}(\epsilon')$ . By the definition of the ball  $B_{\boldsymbol{\theta}'}(\epsilon')$ ,  $\text{supp}(\boldsymbol{\theta}) \supseteq \text{supp}(\boldsymbol{\theta}')$ . Note that  $\mathcal{Q}^*(\boldsymbol{\theta}') > -\infty$  by Lemma 5. By Lemma 1,  $\text{supp}(\boldsymbol{\theta}') \supseteq \text{supp}(\boldsymbol{\theta}^*)$ :

$$\text{supp}(\boldsymbol{\theta}) \supseteq \text{supp}(\boldsymbol{\theta}') \supseteq \text{supp}(\boldsymbol{\theta}^*).$$

This then, again by Lemma 1, implies the claim. Now,  $\widehat{\mathcal{Q}}_{m_j}(\boldsymbol{\theta})$  converges to  $\mathcal{Q}^*(\boldsymbol{\theta})$  uniformly on the ball  $B_{\boldsymbol{\theta}'}(\epsilon')$  since the function is continuous on the ball that is compact.

Let  $0 < \epsilon < 2\epsilon'$ . Note

$$P\left(\|\widehat{\boldsymbol{\theta}}_{m_j} - \boldsymbol{\theta}'\|_\infty > \epsilon/2\right) \rightarrow 0 \quad (15)$$

$$P\left(\sup_{\boldsymbol{\theta} \in B_{\boldsymbol{\theta}'}(\epsilon')} \left| \widehat{\mathcal{Q}}_{m_j}(\boldsymbol{\theta}) - \mathcal{Q}^*(\boldsymbol{\theta}) \right| > \epsilon/2\right) \rightarrow 0 \quad (16)$$

$$P\left(\left| \mathcal{Q}^*(\widehat{\boldsymbol{\theta}}_{m_j}) - \mathcal{Q}^*(\boldsymbol{\theta}') \right| > \epsilon/2\right) \rightarrow 0. \quad (17)$$

(15) is due to the convergence of  $\{\widehat{\boldsymbol{\theta}}_{m_j}\}$ . (16) holds because of the uniform convergence on the ball  $B_{\boldsymbol{\theta}'}(\epsilon')$ . (17) holds because  $\mathcal{Q}^*(\boldsymbol{\theta})$  is continuous at  $\boldsymbol{\theta}'$ .

Recall we want to show  $\mathbb{P}(|\widehat{\mathcal{Q}}_{m_j}(\widehat{\boldsymbol{\theta}}_{m_j}) - \mathcal{Q}^*(\boldsymbol{\theta}')| > \epsilon) \rightarrow 0$ . Note:

$$\begin{aligned} & \mathbb{P}(|\widehat{\mathcal{Q}}_{m_j}(\widehat{\boldsymbol{\theta}}_{m_j}) - \mathcal{Q}^*(\boldsymbol{\theta}')| > \epsilon) \\ & \leq \mathbb{P}(|\widehat{\mathcal{Q}}_{m_j}(\widehat{\boldsymbol{\theta}}_{m_j}) - \mathcal{Q}^*(\widehat{\boldsymbol{\theta}}_{m_j})| > \epsilon/2) + \mathbb{P}(|\mathcal{Q}^*(\widehat{\boldsymbol{\theta}}_{m_j}) - \mathcal{Q}^*(\boldsymbol{\theta}')| > \epsilon/2). \end{aligned}$$

The second term goes to zero by (17). It remains to show that the first term goes to 0:

$$\begin{aligned} & \mathbb{P}(|\widehat{\mathcal{Q}}_{m_j}(\widehat{\boldsymbol{\theta}}_{m_j}) - \mathcal{Q}^*(\widehat{\boldsymbol{\theta}}_{m_j})| > \epsilon/2) \\ & \leq \mathbb{P}\left(|\widehat{\mathcal{Q}}_{m_j}(\widehat{\boldsymbol{\theta}}_{m_j}) - \mathcal{Q}^*(\widehat{\boldsymbol{\theta}}_{m_j})| > \epsilon/2 \mid \|\widehat{\boldsymbol{\theta}}_{m_j} - \boldsymbol{\theta}'\|_\infty > \epsilon/2\right) \mathbb{P}\left(\|\widehat{\boldsymbol{\theta}}_{m_j} - \boldsymbol{\theta}'\|_\infty > \epsilon/2\right) + \\ & \quad \mathbb{P}\left(\left\{|\widehat{\mathcal{Q}}_{m_j}(\widehat{\boldsymbol{\theta}}_{m_j}) - \mathcal{Q}^*(\widehat{\boldsymbol{\theta}}_{m_j})| > \epsilon/2\right\} \cap \left\{\|\widehat{\boldsymbol{\theta}}_{m_j} - \boldsymbol{\theta}'\|_\infty \leq \epsilon/2\right\}\right). \end{aligned}$$

The first term goes to zero by (15). The second term also goes to zero as follows, which completes the proof:

$$\begin{aligned} & \mathbb{P}\left(\left\{|\widehat{\mathcal{Q}}_{m_j}(\widehat{\boldsymbol{\theta}}_{m_j}) - \mathcal{Q}^*(\widehat{\boldsymbol{\theta}}_{m_j})| > \epsilon/2\right\} \cap \left\{\|\widehat{\boldsymbol{\theta}}_{m_j} - \boldsymbol{\theta}'\|_\infty \leq \epsilon/2\right\}\right) \\ & \leq \mathbb{P}\left(\left\{\sup_{\boldsymbol{\theta} \in B_{\boldsymbol{\theta}'}(\epsilon/2)} \left| \widehat{\mathcal{Q}}_{m_j}(\boldsymbol{\theta}) - \mathcal{Q}^*(\boldsymbol{\theta}) \right| > \epsilon/2\right\} \cap \left\{\|\widehat{\boldsymbol{\theta}}_{m_j} - \boldsymbol{\theta}'\|_\infty \leq \epsilon/2\right\}\right) \\ & \leq \mathbb{P}\left(\sup_{\boldsymbol{\theta} \in B_{\boldsymbol{\theta}'}(\epsilon/2)} \left| \widehat{\mathcal{Q}}_{m_j}(\boldsymbol{\theta}) - \mathcal{Q}^*(\boldsymbol{\theta}) \right| > \epsilon/2\right) \\ & \leq \mathbb{P}\left(\sup_{\boldsymbol{\theta} \in B_{\boldsymbol{\theta}'}(\epsilon')} \left| \widehat{\mathcal{Q}}_{m_j}(\boldsymbol{\theta}) - \mathcal{Q}^*(\boldsymbol{\theta}) \right| > \epsilon/2\right) \rightarrow 0, \end{aligned}$$

where the last line is due to (16). □