

## 7.3 Small-Sample Inferences on the Difference Between Two Means

As in §5.4 and §6.4, we don't have elementary small-sample methods for \_\_\_\_\_ populations.

However, if  $X_1, \dots, X_{n_X}$  and  $Y_1, \dots, Y_{n_Y}$  are (possibly small) independent random samples from \_\_\_\_\_ populations with means  $\mu_X$  and  $\mu_Y$  and standard deviations  $\sigma_X$  and  $\sigma_Y$ , then

$$\bar{X} - \bar{Y} \sim N\left(\mu_X - \mu_Y, \frac{\sigma_X^2}{n_X} + \frac{\sigma_Y^2}{n_Y}\right)$$

(This " $\bar{X} - \bar{Y} \sim N(\dots)$ " statement is exact. We figured out an approximate version of it via the CLT for large samples in §7.1. The same reasoning, without \_\_\_\_\_ but with the assumption of \_\_\_\_\_, leads to §7.3's statement.)

But we don't know  $\sigma_X$  or  $\sigma_Y$ , and, with small samples, the approximations  $\sigma_X \approx s_X$  and  $\sigma_Y \approx s_Y$  are \_\_\_\_\_. We still use them to standardize  $\bar{X} - \bar{Y}$ , but, as in the one-small-sample case, we get a \_\_\_\_\_ statistic instead of a \_\_\_\_\_ statistic. Recall that the  $t_\nu$  distributions look like \_\_\_\_\_, but are \_\_\_\_\_ with \_\_\_\_\_.

Experts say  $T = \frac{(\bar{X} - \bar{Y}) - (\mu_X - \mu_Y)}{\sqrt{\frac{s_X^2}{n_X} + \frac{s_Y^2}{n_Y}}} \sim t_\nu (\approx)$ , where  $\nu = \frac{\left(\frac{s_X^2}{n_X} + \frac{s_Y^2}{n_Y}\right)^2}{\frac{(s_X^2/n_X)^2}{n_X-1} + \frac{(s_Y^2/n_Y)^2}{n_Y-1}}$ , rounded \_\_\_\_\_.

### Small-Sample Confidence Interval for the Difference of Two Means

Recall that many confidence intervals have the form

$$(\text{point estimate}) \pm (\text{margin of error}) = \hat{\theta} \pm (\text{table value for confidence}) \times \sigma_{\hat{\theta}}$$

To get a  $(100\%)(1 - \alpha)$  confidence interval for  $\mu_X - \mu_Y$ , start with  $t_{\nu, \alpha/2}$  such that

$$P(-t_{\nu, \alpha/2} < T < t_{\nu, \alpha/2}) = 1 - \alpha$$

Unstandardize  $T$  to get

$$P\left(-t_{\nu, \alpha/2} < \frac{(\bar{X} - \bar{Y}) - (\mu_X - \mu_Y)}{\sqrt{\frac{s_X^2}{n_X} + \frac{s_Y^2}{n_Y}}} < t_{\nu, \alpha/2}\right) = 1 - \alpha$$

Solve for  $(\mu_X - \mu_Y)$  in the middle:

$$P\left((\bar{X} - \bar{Y}) - t_{\nu, \alpha/2} \sqrt{\frac{s_X^2}{n_X} + \frac{s_Y^2}{n_Y}} < (\mu_X - \mu_Y) < (\bar{X} - \bar{Y}) + t_{\nu, \alpha/2} \sqrt{\frac{s_X^2}{n_X} + \frac{s_Y^2}{n_Y}}\right) = 1 - \alpha$$

That is,  $\boxed{(\bar{X} - \bar{Y}) \pm t_{\nu, \alpha/2} \sqrt{\frac{s_X^2}{n_X} + \frac{s_Y^2}{n_Y}}}$  contains \_\_\_\_\_ for a proportion \_\_\_\_\_ of samples.

## Small-Sample Hypothesis Test for the Difference of Two Means

Recall that many hypothesis tests use test statistics of the form

$$\frac{(\text{point estimate}) - (\text{parameter value under } H_0)}{(\text{estimated or true}) \text{ standard deviation of point estimate}}$$

which tells how far the estimate is from the parameter, in standard deviations.

Let  $X_1, \dots, X_{n_X}$  and  $Y_1, \dots, Y_{n_Y}$  be independent random samples from *normal* populations with means  $\mu_X$  and  $\mu_Y$  and standard deviations  $\sigma_X$  and  $\sigma_Y$ . To test  $H_0 : \mu_X - \mu_Y = \Delta_0$ ,

1. State null and alternative hypotheses,  $H_0$  and  $H_1$

2. Check assumptions

3. Find the test statistic  $t = \frac{(\bar{x} - \bar{y}) - \Delta_0}{\sqrt{\frac{s_X^2}{n_X} + \frac{s_Y^2}{n_Y}}}$

4. Find the degrees of freedom,  $\nu = \frac{\left(\frac{s_X^2}{n_X} + \frac{s_Y^2}{n_Y}\right)^2}{\frac{(s_X^2/n_X)^2}{n_X-1} + \frac{(s_Y^2/n_Y)^2}{n_Y-1}}$ , rounded down

5. Find the  $P$ -value, which is an area under the  $t_\nu$  curve depending on  $H_1$ :

$H_1 : \mu_X - \mu_Y > \Delta_0 \implies P\text{-value} = P(T > t)$ , the area right of  $t$

$H_1 : \mu_X - \mu_Y < \Delta_0 \implies P\text{-value} = P(T < t)$ , the area left of  $t$

$H_1 : \mu_X - \mu_Y \neq \Delta_0 \implies P\text{-value} = P(|T| > |t|)$ , the sum of the two tail areas

6. Draw a conclusion

## Examples

e.g. (Like p. 292 #17) While working an avalanche control route at Jackson Hole, Eric says his overhand spiral bomb throws penetrate the snowpack better, and blow a bigger hole, than Tyler's underhand lobbs. Tyler says throwing style doesn't matter (and thinks extra penetration could even shrink the hole). To test  $H_0 : \mu_{\text{spiral}} - \mu_{\text{lob}} = 0$  vs.  $H_1 : \mu_{\text{spiral}} - \mu_{\text{lob}} \neq 0$ , each man throws five bombs, and they measure the resulting hole diameters (in meters):

Tyler (spiral): 2.13 2.14 2.10 2.09 2.07

Eric (lob): 2.09 2.15 2.07 2.13 2.07

Can Tyler conclude that hole diameters are different for overhand spiral throws? The argument's loser must \_\_\_\_\_ for the entire patrol.

e.g. A study on logging in Borneo counted the number of tree species in 12 randomly chosen unlogged forest plots and in 9 similar plots logged 8 years earlier.

	#Tree species											
Unlogged	22	18	22	20	15	21	13	13	19	13	19	15
Logged	17	4	18	14	18	15	15	10	12			

Does logging reduce the mean number of species in a plot after 8 years? Use  $\alpha = .10$ .

e.g. Find a 90% confidence interval for the difference in mean number of species between unlogged and logged plots.

## Caution

Statistics programs offer the option to find an interval or run a test assuming  $\sigma_X = \sigma_Y$ . This option is \_\_\_\_\_ when the assumption is incorrect, and should usually be \_\_\_\_\_.