

- 9.1 Review: One-Way ANOVA
- 9.2 Pairwise Comparisons in One-Factor Experiments
- 9.3 Two-Factor Experiments (part 1 of 2)

# 9.1 Review: One-Way ANOVA

e.g. Make a quick one-way ANOVA table for these data to test  $H_0: \mu_1 = \mu_2 = \mu_3$ :

Sample	Da	ata			$\bar{x}_{i.}$
1	0	1	2		
2	1	2	3		
3	5	6	7		
	Overall $\bar{x}_{} =$				

,					
Source	DF	SS	MS	F	P
Treatment	I-1	SSTr = $\sum_{i=1}^{I} J_i (\bar{X}_{i.} - \bar{X}_{})^2$	$MSTr = \frac{SSTr}{I-1}$	$F = \frac{\text{MSTr}}{\text{MSE}}$	$P(F_{I-1,N-I} > \underline{\hspace{1cm}})$
Error	N-I	SSE = $\sum_{i=1}^{I} \sum_{j=1}^{J_i} (X_{ij} - \bar{X}_{i.})^2$			
Total	N-1	$ SST  = \sum_{i=1}^{I} \sum_{j=1}^{J_i} (X_{ij} - \bar{X}_{})^2$ $= SSTr + SSE$			

Conclusion:

# 9.2 Pairwise Comparisons in One-Factor Experiments

A problem:

- One-way ANOVA tests  $H_0: \mu_1 = \cdots = \mu_I$  (for I samples), but rejecting  $H_0$  doesn't tell are different.
- If we calculate a level- $(1-\alpha)$  confidence interval for the difference of each pair of means,  $\mu_i \mu_j$ , the probability of \_\_\_\_\_ intervals containing their respective differences \_\_\_\_ is usually \_\_\_\_\_

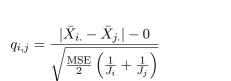
The Tukey-Kramer Method of Multiple Comparisons

• Experts say the Tukey-Kramer level- $(1-\alpha)$  \_\_\_\_\_ confidence intervals for  $\mu_i - \mu_j$  are

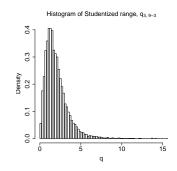
$$\bar{X}_{i.} - \bar{X}_{j.} \pm q_{I,N-I,\alpha} \sqrt{\frac{\text{MSE}}{2} \left(\frac{1}{J_i} + \frac{1}{J_j}\right)}$$

We have  $(1-\alpha)$  confidence that these intervals contain  $\mu_i - \mu_j$ 

• To test all  $H_0: \mu_i - \mu_j = 0$  \_\_\_\_\_, use the test statistics



The P-value for test i, j is \_\_\_\_\_ To \_\_\_\_ calculating  $\binom{I}{2}$  P-values, here's a \_\_\_\_\_:



e.g. Which pairs of population means from the "9.1 Review" example, above, differ at the 5% level?

# 9.3 Two-Factor Experiments (part 1 of 2)

#### Example

e.g. (p. 437 #6) A tool's lifetime was studied under three settings for feed rate and three for speed. Four tools were tested under each combination of settings. The lifetimes (in hours) were:

	Speed		
Feed Rate	Slow	Medium	Fast
Light	60.6 57.0 61.4 59.7	57.8 59.4 62.8 58.2	56.5 52.3 58.1 53.9
Medium	51.2 53.1 48.3 51.6	$49.6\ 48.1\ 49.8\ 51.1$	$45.7\ 48.6\ 45.0\ 49.2$
Heavy	44.8 46.7 41.9 51.3	$46.6\ 41.4\ 38.3\ 37.9$	$37.2\ 32.8\ 39.9\ 35.9$

Does feed rate influence lifetime? Does speed influence lifetime?

#### Vocabulary

Recall (§9.1) that a factor is an	_ variable in an experiment; its values are
A two-factor experiment has a with $J$ levels, for a total of con	$\_$ factor with $I$ levels and a $\_\_\_\_$ factor mbinations called $\_\_\_\_$ .
	measurements are made are units design uses the same number of replicates designs.)
	oservations are taken on $\_$ of the $IJ$ possible at for some treatments, are hard to interpret.)
	hose mean outcome is $\mu_{ij}$ . Replicates corresponding amn factor are denoted $X_{ij1}, \dots, X_{ijK}$ and represent ulation.
A two-way analysis of variance tests whether $\{\mu\}$ or or both.	$\{x_{ij}\}$ are affected by varying the

### Parameterization for Two-Way Analysis of Variance

- $\mu_{ij} =$  \_\_\_\_\_ mean corresponding to treatment in row \_\_\_\_ and column \_\_\_\_
- $\bar{\mu}_{i.} = \frac{1}{J} \sum_{j=1}^{J} \mu_{ij} = \text{average of means in}$
- $\bar{\mu}_{.j} = \frac{1}{I} \sum_{i=1}^{I} \mu_{ij}$  = average of means in \_\_\_\_\_\_
- $\mu = \frac{1}{IJ} \sum_{i=1}^{I} \sum_{j=1}^{J} \mu_{ij} \left( = \frac{1}{I} \sum_{i=1}^{I} \bar{\mu}_{i.} = \frac{1}{J} \sum_{j=1}^{J} \bar{\mu}_{.j} \right) = population grand mean = average of treatment means$

Here's a table showing these (unknown) population means:

	Column level			
Row level	1		J	Row mean
1	$\mu_{11}$	• • •	$\mu_{1J}$	$\bar{\mu}_{1.}$
:	:		:	:
I	$\mu_{I1}$	• • •	$\mu_{IJ}$	$ar{\mu}_{I.}$
Column mean	$\bar{\mu}_{.1}$	• • •	$\bar{\mu}_{.J}$	$\mu$

• Write  $\mu_{ij}$  as

$$\mu_{ij} = \mu + (\bar{\mu}_{i.} - \mu) + (\bar{\mu}_{.j} - \mu) + (\mu_{ij} - \bar{\mu}_{i.} - \bar{\mu}_{.j} + \mu)$$

$$= \mu + \alpha_{i} + \beta_{j} + \gamma_{ij}$$

where

-  $\alpha_i = \bar{\mu}_i$ . -  $\mu = i^{\text{th}}$  row effect, which indicates how much \_\_\_\_\_ tends to produce outcomes larger or smaller than  $\mu$  (note:  $\sum_{i=1}^{I} \alpha_i =$  \_\_\_\_\_)

-  $\beta_j = \bar{\mu}_{.j} - \mu = j^{\text{th}}$  column effect, which indicates how much \_\_\_\_\_ tends to produce outcomes larger or smaller than  $\mu$  (note:  $\sum_{j=1}^{J} \beta_j =$ \_\_\_\_\_)

-  $\gamma_{ij} = \mu_{ij} - \bar{\mu}_{i.} - \bar{\mu}_{.j} + \mu$  [=  $\mu_{ij} - (\mu + \alpha_i + \beta_j)$ ], the ij \_\_\_\_\_\_, which measures the degree to which the effect of a row (or column) factor depends on the level of the column (or row) factor it is paired with (note:  $\sum_{i=1}^{I} \gamma_{ij} =$ \_\_\_\_\_ and  $\sum_{j=1}^{J} \gamma_{ij} =$ \_\_\_\_\_)

Main effects consist of the I row effects and the J column effects; they don't include the \_\_\_\_\_\_. If the interactions  $\gamma_{ij}$  are all 0, then  $\mu_{ij} = \mu + \alpha_i + \beta_j + \gamma_{ij}$  becomes

$$\mu_{ij} = \mu + \alpha_i + \beta_j$$

In this case we have an \_\_\_\_\_ model, such that the treatment mean  $\mu_{ij}$  is the population grand mean  $\mu$  plus the amount  $\alpha_i$  resulting from using row level i plus the amount  $\beta_j$  resulting from using column level j. (If some interactions are not zero, the combined effect of a row and column can't be determined from the \_\_\_\_\_ alone.)

### Using Two-Way ANOVA to Test Hypotheses

Two-Way ANOVA tests

- Does the \_\_\_\_\_ model hold? That is, is it plausible that  $\gamma_{ij} =$  \_\_\_\_ for all i and j?  $H_0: \gamma_{11} = \cdots = \gamma_{IJ} = 0$
- If the additive model holds  $(H_0 \text{ is } \underline{\hspace{1cm}})$ ,
  - is the mean outcome \_\_\_\_\_ for all levels of the row factor?

 $H_0: \alpha_1 = \cdots = \alpha_I = 0$ 

- is the mean outcome \_\_\_\_\_ for all levels of the column factor?

 $H_0: \beta_1 = \cdots = \beta_J = 0$ 

Retain all  $\implies$  " $\mu_{ij} = \underline{\hspace{1cm}}$  for all i, j" is plausible  $\implies$  treatment differences  $\underline{\hspace{1cm}}$