
Certifying Robustness to Programmable Data Bias in Decision Trees

Anna P. Meyer, Aws Albarghouthi,* and Loris D’Antoni

Department of Computer Sciences
University of Wisconsin–Madison
Madison, WI 53706

{anameyer, aws, loris}@cs.wisc.edu

Abstract

Datasets can be biased due to societal inequities, human biases, under-representation of minorities, etc. Our goal is to *certify* that models produced by a learning algorithm are *pointwise-robust* to potential dataset biases. This is a challenging problem: it entails learning models for a large, or even infinite, number of datasets, ensuring that they all produce the same prediction. We focus on decision-tree learning due to the interpretable nature of the models. Our approach allows programmatically specifying *bias models* across a variety of dimensions (e.g., missing data for minorities), composing types of bias, and targeting bias towards a specific group. To certify robustness, we use a novel symbolic technique to evaluate a decision-tree learner on a large, or infinite, number of datasets, certifying that each and every dataset produces the same prediction for a specific test point. We evaluate our approach on datasets that are commonly used in the fairness literature, and demonstrate our approach’s viability on a range of bias models.

1 Introduction

The proliferation of machine-learning algorithms has raised alarming questions about fairness in automated decision-making [4]. In this paper, we focus our attention on bias in training data. Data can be biased due to societal inequities, human biases, under-representation of minorities, malicious data *poisoning*, etc. For instance, historical data can contain human biases, e.g., certain individuals’ loan requests get rejected, although (if discrimination were not present) they should have been approved, or women in certain departments are consistently given lower performance scores by managers.

Given biased training data, we are often unable to de-bias it because we do not know which samples are affected. This paper asks, *can we certify (prove) that our predictions are robust under a given form and degree of bias in the training data?* We aim to answer this question without having to show which data are biased (i.e., poisoned). Techniques for certifying poisoning robustness (*i*) focus on specific poisoning forms, e.g., label-flipping [31], or (*ii*) perform certification using defenses that create complex, uninterpretable classifiers, e.g., due to randomization or ensembling [23, 24, 31]. To address limitation (*i*), we present *programmable bias definitions* that model nuanced biases in practical domains. To address (*ii*), we target *existing* decision-tree learners—considered interpretable and desirable for sensitive decision-making [32]—and exactly certify their robustness, i.e., provide proofs that the bias in the data will not affect the outcome of the trained model on a given point.

We begin by presenting a *language* for programmatically defining *bias models*. A bias model allows us to flexibly specify what sort of bias we suspect to be in the data, e.g., up to $n\%$ of the women *may have* wrongly received a negative job evaluation. Our bias-model language is generic, allowing

* Author’s name in native alphabet: أوس البرغوثي

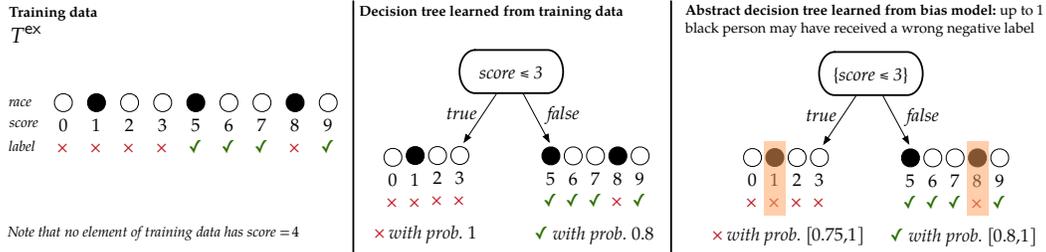


Figure 1: A simple, hypothetical running example

us to *compose* simpler bias models into more complex ones, e.g., up to $n\%$ of the women may have wrongly received a negative evaluation *and* up to $m\%$ of Black men’s records may have been completely missed. The choice of bias model depends on the provenance of the data and the task.

After specifying a bias model, our goal is to certify *pointwise robustness to data bias*: Given an input x , we want to ensure that no matter whether the training data is biased or not, the resulting model’s prediction for x remains the same. Certifying pointwise robustness is challenging. One can train a model for every perturbation (as per a bias model) of a dataset and make sure they all agree. But this is generally not feasible, because the set of possible perturbations can be large or infinite. Recall the bias model where up to $n\%$ of women may have wrongly received a negative label. For a dataset with 1000 women and $n = 1\%$, there are more than 10^{23} possible perturbed datasets.

To perform bias-robustness certification on decision-tree learners, we employ *abstract interpretation* [12] to symbolically run the decision-tree-learning algorithm on a large or infinite set of datasets simultaneously, thus learning a *set* of possible decision trees, represented compactly. The crux of our approach is a technique that lifts operations of decision-tree learning to symbolically operate over a *set of datasets* defined using our bias-model language. As a starting point, we build upon Drews et al.’s [16] demonstration of poisoning-robustness certification for the simple bias model where an adversary may have added fake training data. Our approach completely reworks and extends their technique to target the bias-robustness problem and handle complex bias models, including ones that may result in an infinite number of datasets.

Contributions. We make three contributions: (1) We formalize the bias-robustness-certification problem and present a language to compositionally define bias models. (2) We present a symbolic technique that performs decision-tree learning on a set of datasets defined by a bias model, allowing us to perform certification. (3) We evaluate our approach on a number of bias models and datasets from the fairness literature. Our tool can certify pointwise robustness for a variety of bias models; we also show that some datasets have *unequal robustness-certification rates* across demographics groups.

Running example. Consider the example in Fig. 1; our goal is to classify who should be hired based on a test score. A standard decision-tree-learning algorithm would choose the split (predicate) $score \leq 3$, assuming we restrict tree depth to 1.² As shown in Fig. 1 (middle), the classification depends on the data split; e.g., on the right hand side, we see that a person with $score > 3$ is accepted, because the proportion (“probability”) of the data with positive labels and $score > 3$ is $4/5 (> 1/2)$.

Now suppose that our bias model says that up to one Black person in the dataset may have received a wrongful rejection. Our goal is to show that even if that is the case, the prediction of a new test sample x will not change. As described above, training decision trees for all possible modified datasets is generally intractable. Instead, we symbolically learn a set of possible decision trees compactly, as illustrated in Fig. 1 (right). In this case the learning algorithm always chooses $score \leq 3$ (generally, our algorithm can capture all viable splits). However, the proportion of labels on either branch varies. For example, on the right, if the highlighted sample is wrongly labeled, then the ratio changes from 0.8 to 1. To efficiently perform this calculation, we lift the learning algorithm’s operations to *interval arithmetic* and represent the probability as $[0.8, 1]$. Given a new test sample $x = \langle race=Black, score=7 \rangle$, we follow the right branch and, since the interval is always larger than 0.5, we certify that

²Other predicates, e.g., $score \leq 4$, will yield the same split. We choose a single split for illustrative purposes here. (The implementation considers all possible splits that yield distinct partitions, so it would consider $score \leq 3$ and $score \leq 4$ as a single entity.)

the algorithm is robust for x . In general, however, due to the use of abstraction, our approach may fail to find tight intervals, and therefore be unable to certify robustness for all robust inputs.

2 Related work

Ties to poisoning. Our dataset bias language captures existing definitions of *data poisoning*, where an attacker is assumed to have maliciously modified training data. Poisoning has been studied extensively. Most works have focused on attacks [6, 11, 25, 33, 38, 39, 40] or on training models that are empirically less vulnerable (defenses) [3, 9, 19, 28, 31, 36]. Our work differs along a number of dimensions: (1) We allow programmatic, custom, composable definitions of bias models; notably, to our knowledge, no other work in this space allows for *targeted* bias, i.e., restricting bias to a particular subgroup. (2) Our work aims to certify and quantify robustness of an existing decision-tree algorithm, not to modify it (e.g., via bagging or randomized smoothing) to improve robustness [23, 24, 31].

Statistical defenses show that a learner is robust *with high probability*, often by modifying a base learner using, e.g., randomized smoothing [31], outlier detection [36], or bagging [22, 23]. Non-statistical certification (including abstract interpretation) has mainly focused on *test-time* robustness, where the vicinity (e.g., within an ℓ_p norm) of an input is proved to receive the same prediction [2, 21, 30, 35, 37, 1]. Test-time robustness is a simpler problem than our *train-time* robustness problem because it does not have to consider the mechanics of the learner on sets of datasets. The only work we know of that certifies train-time robustness of decision trees is by Drews et al. [16] and focuses on poisoning attacks where an adversary *adds fake data*. Our work makes a number of significant leaps beyond this work: (1) We frame data bias as programmable, rather than fixed, to mimic real-world bias scenarios, an idea that has gained traction in a variety of domains, e.g., NLP [41]. (2) We lift a decision-tree-learning algorithm to operate over sets of datasets represented via our bias-model language. (3) We investigate the bias-robustness problem through a fairness lens, particularly with an eye towards robustness rates for various demographic groups.

Ties to fairness. The notion of *individual fairness* specifies that *similar individuals* should receive similar predictions [18]; by contrast, we certify that no individual should receive different predictions under models trained by *similar datasets*. Black and Fredrickson explore the problem of how individuals’ predictions change under models trained by similar datasets, but their concept of similarity is limited to removing a *single* data point [7]. Data bias, in particular, has received some attention in the fairness literature. Chen et al. suggest adding missing data as an effective approach to remedying bias in machine learning [10], which is one operation that our bias language captures. Mandal et al. build on the field of *distributional robustness* [5, 27, 34] to build classifiers that are empirically group-fair across a variety of nearby distributions [26]. Our problem domain is related to distributional robustness because we certify robustness over a family of similar datasets; however, we define specific data-transformation operators to define similarity, and, unlike Mandal et al., we certify existing learners instead of building empirically robust models.

Ties to robust statistics. There has been renewed interest in robust statistics for machine learning [13, 14]. Much of the work concerns outlier detection for various learning settings, e.g., estimating parameters of a Gaussian. The distinctions are two-fold: (1) We deal with rich, nuanced bias models, as opposed to out-of-distribution samples, and (2) we aim to certify that predictions are robust for a specific input, a guarantee that cannot be made by robust-statistics-based techniques [15].

3 Defining data bias programmatically

We define the *bias-robustness problem* and a language for defining bias models *programmatically*.

Bias models. A dataset $T \subseteq \mathcal{X} \times \mathcal{Y}$ is a set of pairs of samples and labels, where $\mathcal{Y} = \{0, \dots, n-1\}$. For a dataset T , we will use $T_{\mathcal{X}}$ to denote $\{x \mid (x, y) \in T\}$. A *bias model* B is a function that takes a dataset and returns a *set of datasets*. We call $B(T) \subseteq 2^{\mathcal{X} \times \mathcal{Y}}$ a *bias set*. We assume that $T \in B(T)$. Intuitively, $B(T)$ represents *all datasets that could have existed had there been no bias*.

Pointwise data-bias robustness. Assume we have a learning algorithm A that, given a training dataset T , deterministically returns a classifier h_T from some hypothesis class \mathcal{H} . Fix a dataset T and bias model B . Given a sample $x \in \mathcal{X}$, we say that A is *pointwise robust* (or robust for short) on x iff

$$\text{there is a label } i \text{ such that for all } T' \in B(T), \text{ we have } h_{T'}(x) = i \quad (1)$$

Basic components of a bias model. We begin with basic bias models.

Missing data: A common bias in datasets is missing data, which can occur via poor historical representation of a subgroup (e.g., women in CS-department admissions data), or from present-day biases or shortsightedness (e.g., a survey that bypasses low-income neighborhoods). Using the parameter m as the maximum number of missing elements, we formally define:

$$\text{MISS}_m(T) = \{T' \mid T' \supseteq T, |T' \setminus T| \leq m\}$$

$\text{MISS}_m(T)$ defines an *infinite* number datasets when the sample space is infinite (e.g., \mathbb{R} -valued).

Example 3.1. Using T^{ex} from Fig. 1, $\text{MISS}_1(T^{\text{ex}})$ is the set of all datasets that are either T^{ex} or T^{ex} plus any new element (x, y) with arbitrary race, score, and label.

Label flipping: Historical data can contain human biases, e.g., in loan financing, certain individuals' loan requests get rejected due to discrimination. Or consider employee-performance data, where women in certain departments are consistently given lower scores by managers. We model such biases as label flipping, where labels of up to l individuals in the dataset may be incorrect:

$$\text{FLIP}_l(T) = \{T' \mid |T|=|T'|, |T \setminus T'| \leq l, T_x = T'_x\}$$

Example 3.2. Using T^{ex} from Fig. 1 and a bias model FLIP_1 , we have $\text{FLIP}_1(T^{\text{ex}}) = \{T_0^{\text{ex}}, \dots, T_3^{\text{ex}}, T_5^{\text{ex}}, \dots, T_9^{\text{ex}}, T^{\text{ex}}\}$, where T_i^{ex} is T^{ex} with the label of the element with score = i changed.

Fake data: Our final bias model assumes the dataset may contain *fake* data. One cause may be a malicious user who enters fraudulent data into a system (often referred to as *poisoning*). Alternatively, this model can be thought of as the inverse of MISS, e.g., we over-collected data for men.

$$\text{FAKE}_k(T) = \{T' \mid T' \subseteq T, |T \setminus T'| \leq k\}$$

Example 3.3. Using T^{ex} from Fig. 1 and a bias model FAKE_1 , we get $\text{FAKE}_1(T^{\text{ex}}) = \{T_{\downarrow 0}^{\text{ex}}, \dots, T_{\downarrow 3}^{\text{ex}}, T_{\downarrow 5}^{\text{ex}}, \dots, T_{\downarrow 9}^{\text{ex}}, T^{\text{ex}}\}$ where $T_{\downarrow i}^{\text{ex}}$ is T^{ex} such that the element with score = i has been removed.

Targeted bias models. Each bias model has a *targeted* version that limits the bias to a specified group of data points. For example, consider the missing data transformation. If we suspect that data about women is missing from an HR database, we can limit the MISS transformation to only add data points with *gender* = female. Formally, we define a predicate $g : \mathcal{X} \times \mathcal{Y} \rightarrow \mathbb{B}$, where $\mathbb{B} = \{\text{true}, \text{false}\}$.

$$\text{MISS}_m^g(T) = \{T' \mid T' \supseteq T, |T' \setminus T| \leq m, \text{ and } g(x, y) \text{ is true } \forall (x, y) \in (T' \setminus T)\}$$

Targeted versions of label-flipping and fake data can be defined in a similar way.

Example 3.4. In Fig. 1 (right), we used bias model FLIP_1^g , where g targets Black people with negative labels. This results in the bias set $\text{FLIP}_1^g(T^{\text{ex}}) = \{T_1^{\text{ex}}, T_8^{\text{ex}}, T^{\text{ex}}\}$, where T_i^{ex} is T^{ex} with the label of the element with score = i changed (recall that scores 1 and 8 belong to Black people in T^{ex}).

Composite bias models. We can compose basic components to generate a *composite model*. Specifically, we define a composite model B as a finite set of arbitrary basic components, that is,

$$B = [\text{MISS}_{m_1}^{g_1}, \dots, \text{MISS}_{m_j}^{g_j}, \text{FLIP}_{l_1}^{g_1}, \dots, \text{FLIP}_{l_p}^{g_j+p}, \text{FAKE}_{k_1}^{g_j+p+1}, \dots, \text{FAKE}_{k_q}^{g_j+p+q}] \quad (2)$$

$B(T)$ is generated from T by applying the basic components of B iteratively. We must apply the constituent components in an optimal order, i.e., one that generates all datasets that can be created through applying the transformers in *any* order. To do this, we apply components of the same type in any order and apply transformers of different types in the order MISS, FLIP, FAKE (see Appendix).

Example 3.5. Suppose $B = [\text{MISS}_2^{g_1}, \text{FAKE}_1]$. Then $B(T)$ is the set of all datasets obtained by adding up to 2 arbitrary data points that satisfy g_1 to T , and then removing any up to 1 data point.

4 Certifying robustness for decision-tree learning

We begin with a simplified version of the CART algorithm [8], which is our target for certification.

Given a dataset T and a Boolean function (predicate) $\phi : \mathcal{X} \rightarrow \mathbb{B}$, we define:

$$T_\phi = \{(x, y) \in T \mid \phi(x)\},$$

i.e., T_ϕ is the set of elements satisfying ϕ . Analogously, $T_{\neg\phi} = \{(x, y) \in T \mid \neg\phi(x)\}$.

Example 4.1. Using $\phi \triangleq \text{score} \leq 3$, we have $T_\phi^{\text{ex}} = \{0, 1, 2, 3\}$ and $T_{-\phi}^{\text{ex}} = \{5, 6, 7, 8, 9\}$.

Learning algorithm. To formalize our approach, it suffices to consider a simple algorithm that learns a decision stump, i.e., a tree of depth 1. Therefore, the job of the algorithm is to choose a predicate (splitting rule) ϕ from a set of predicates Φ that optimally splits the dataset T into two datasets. Formally, we define $\text{pr}_i(T)$ as the proportion of T with label i , i.e.,

$$\text{pr}_i(T) = |\{(x, i) \in T\}| / |T| \quad (3)$$

We use pr to calculate *Gini impurity* (imp), that is,

$$\text{imp}(T) = \sum_{i=0}^{n-1} \text{pr}_i(T)(1 - \text{pr}_i(T))$$

Using imp , we assign each dataset-predicate pair a cost, where a low value indicates that ϕ splits T *cleanly*, i.e., elements of T_ϕ (conversely, $T_{-\phi}$) have mostly the same label:

$$\text{cost}(T, \phi) = |T_\phi| \cdot \text{imp}(T_\phi) + |T_{-\phi}| \cdot \text{imp}(T_{-\phi})$$

Finally, we select the predicate that results in the lowest cost (we break ties arbitrarily), as defined by the split operator:

$$\text{split}(T) = \underset{\phi \in \Phi}{\text{argmin}} \text{cost}(T, \phi)$$

Example 4.2. For $\phi \triangleq \text{score} \leq 3$, $\text{cost}(T^{\text{ex}}, \phi) = 4 \times 0 + 5 \times 0.32 = 1.6$.

Inference. Given an optimal predicate ϕ and a new sample x to classify, we return the label with the highest proportion in the branch of the tree that x takes. Formally,

$$\text{infer}(T, \phi, x) = \underset{i}{\text{argmax}} \text{pr}_i(T_{\phi'}),$$

where ϕ' is ϕ if $\phi(x) = \text{true}$; otherwise, ϕ' is $-\phi$.

4.1 Certifying bias robustness with abstraction

Given a dataset T , bias model B , and sample x , our goal is to prove robustness (Eq. (1)): no matter which dataset in $B(T)$ was used to learn a decision tree, the predicted label of x is the same. Formally,

$$\text{there is a label } i \text{ s.t. for all } T' \in B(T), \text{infer}(T', \phi', x) = i, \text{ where } \phi' = \text{split}(T') \quad (4)$$

The naïve way to prove this is to learn a decision tree using each dataset in $B(T)$ and compare the results. This approach is intractable or impossible, as $|B(T)|$ may be combinatorially large or infinite.

Instead, we *abstractly evaluate* the decision-tree-learning algorithm on the entire bias set $B(T)$ in a symbolic fashion, without having to enumerate all datasets. Specifically, for each operator in the decision-tree-learning algorithm, we define an abstract analogue, called an *abstract transformer* [12], that operates over *sets of training sets* symbolically. An abstract transformer is an approximation of the original operator, in that it *over-approximates* the set of possible outputs on the set $B(T)$.

Sound abstract transformers. Consider the pr_i operator, which takes a dataset and returns a real number. We define an abstract transformer $\text{pr}_i^a(B(T))$ that takes a set of datasets (defined as a bias set) and returns an interval, i.e., a *subset* of \mathbb{R} . The resulting interval defines a range of possible values for the probability of class i . E.g., an interval $\text{pr}_i^a(B(T))$ may be $[0.1, 0.3]$, meaning that the proportion of i -labeled elements in datasets in $B(T)$ is between 0.1 and 0.3, inclusive.

Given intervals computed by pr_i^a , downstream operators will be lifted into interval arithmetic, which is fairly standard. E.g., for $a, b, c, d \in \mathbb{R}$, $[a, b] + [c, d] = [a + c, b + d]$. It will be clear from context when we are applying arithmetic operators to intervals. For a sequence $\{x_i\}_i$ of intervals, $\text{argmax}_i x_i$ returns a *set of possible indices*, as intervals may overlap and there may be no unique maximum.

Example 4.3. Let $I = \{[1, 2], [4, 8], [6, 7], [4, 5]\}$. Then $\max(I) = \{[4, 8], [6, 7]\}$ because 6 is the greatest lower bound of I , and $[4, 8]$ and $[6, 7]$ are the only intervals in I that contain 6.

For the entire certification procedure to be correct, pr^a and all other abstract transformers must be *sound*. That is, they should over-approximate the set of possible outputs. Formally, pr_i^a is a sound approximation of pr_i iff for all i and all $T' \in B(T)$, we have $\text{pr}_i(T') \in \text{pr}_i^a(B(T))$.

Certification process. To perform certification, we use an abstract transformer $\text{split}^a(B(T))$ to compute a *set* of best predicates Φ^a for $B(T)$. The reason split^a returns a set of predicates is because its input is a set of datasets that may result in different optimal splits. Then, we use an abstract transformer $\text{infer}^a(B(T), \Phi^a, x)$ to compute a *set* of labels for x . If infer^a returns a singleton set, then we have proven pointwise robustness for x (Eq. (4)); otherwise, we have an inconclusive result—we cannot falsify robustness because abstract transformers are over-approximate.

4.2 Abstract transformers for pr

We focus on the most challenging transformer, pr^a ; in § 4.3, we show the rest of the transformers.

Abstracting missing data. We begin by describing pr^a for missing data bias, $B = \text{MISS}_m$. From now on, we use c_i to denote the number of samples $(x, y) \in T$ with $y = i$. We define pr_i^a by considering how we can add data to *minimize* the fraction of i 's in T for the lower bound of the interval, and *maximize* the fraction of i 's in T for the upper bound. To minimize the fraction of i 's, we add m elements with label $j \neq i$; to maximize the fraction of i 's, we add m elements with label i .

$$\text{pr}_i^a(\text{MISS}_m(T)) = \left[\frac{c_i}{|T| + m}, \frac{c_i + m}{|T| + m} \right] \quad (5)$$

Example 4.4. Given $B = \text{MISS}_1$, we have $\text{pr}_i^a(B(T^{\text{ex}})) = \left[\frac{4}{10}, \frac{6}{10} \right]$.

Abstracting label-flipping. Next, we define pr_i^a for label-flipping bias, where $B = \text{FLIP}_l$. Intuitively, we can minimize the proportion of i 's by flipping l labels from i to $j \neq i$, and maximize the proportion of i 's by flipping l labels from $j \neq i$ to i . The caveat here is that if there are fewer than l of whichever label we want to flip, we are limited by c_i or $\sum_{j \neq i} c_j$, depending on flipping direction.

$$\text{pr}_i^a(\text{FLIP}_l(T)) = \left[\frac{c_i - \min(l, c_i)}{|T|}, \frac{c_i + \min(l, \sum_{j \neq i} c_j)}{|T|} \right] \quad (6)$$

Example 4.5. Given $B = \text{FLIP}_1$, we have $\text{pr}_i^a(B(T^{\text{ex}})) = \left[\frac{4}{9}, \frac{6}{9} \right]$.

Fake data bias models can be abstracted similarly (see Appendix).

Abstracting targeted bias models. We now show how to abstract targeted bias models, where a function g restricts the affected samples. To begin, we limit g to only condition on features, not the label. In the case of MISS^g , the definition of pr^a does not change, because even if we restrict the characteristics of the elements that we can add, we can still add up to m elements with any label.

In the case of label-flipping, we constrain the parameter l to be no larger than $|T_g|$. Formally, we define $l_i = \min(l, |\{(x, i) \in T : g(x)\}|)$ and then

$$\text{pr}_i^a(\text{FLIP}_l^g(T)) = \left[\frac{c_i - l_i}{|T|}, \frac{c_i + \min(\sum_{j \neq i} l_j, l)}{|T|} \right] \quad (7)$$

The definition for fake data is similar (see Appendix). The above definition is sound when g conditions on the label; however, the Appendix includes a more precise definition of pr^a for that scenario.

Abstracting composite bias models. Now consider a composite bias model consisting of all the basic bias models. Intuitively, pr_i^a will need to reflect changes in c_i that occur from adding data, flipping labels, and removing data. First, we consider a bias model with just one instance of each MISS , FLIP , and FAKE , i.e., $B = [\text{MISS}_m^{g_1}, \text{FLIP}_l^{g_2}, \text{FAKE}_k^{g_3}]$. We define auxiliary variables $l_i = \min(l, |\{(x, i) \in T : g_2(x, i)\}|)$ and $k_i = \min(k, |\{(x, i) \in T : g_3(x, i)\}|)$. Intuitively, these variables represent the number of elements with label i that we can alter. Conversely, to represent the elements with a label other than i , we will use $l'_i = \min(l, \sum_{j \neq i} l_j)$ and $k'_i = \min(k, \sum_{j \neq i} c_j)$.

$$\text{pr}_i^a(B(T)) = \left[\max \left(0, \frac{c_i - l_i - k_i}{|T| - k_i + m} \right), \min \left(1, \frac{c_i + l'_i + m}{|T| - k'_i + m} \right) \right] \quad (8)$$

Extending the above definition to allow multiple uses of the same basic model, e.g., $\{\text{FLIP}_{l_1}^{g_1}, \text{FLIP}_{l_2}^{g_2}\}$ is simple: essentially, we just sum l_1 and l_2 . A full formal definition is in the Appendix.

Theorem 1. pr^a is a sound abstract transformer. (In the Appendix, we also show that pr^a is precise.)

4.3 An abstract decision-tree algorithm

We define the remaining abstract transformers, with the goal of certification. Our definitions are based on Drews et al. [16]; the key difference is the T_ϕ operation, which is dependent on the bias model.

Filtering. We need $B(T)_\phi$, the abstract analogue of T_ϕ . For FLIP and FAKE, we define $B(T)_\phi = B(T_\phi)$. But for MISS, we have to alter the bias model, too, since after filtering on ϕ we only want to add new elements that satisfy ϕ . We define $\text{MISS}_m^g(T)_\phi = \text{MISS}_m^{g \wedge \phi}(T_\phi)$. Filtering composite bias models applies these definitions piece-wise (see a full definition and soundness proof in the Appendix).

Gini impurity. We lift imp to interval arithmetic: $\text{imp}^a(T) = \sum_{i=1}^n \text{pr}_i^a(T)([1, 1] - \text{pr}_i^a(T))$.

Cost. Recall that cost relies on $|T_\phi|$. We want an abstract analogue of $|T_\phi|$ that represents the range of sizes of datasets in $B(T)_\phi$ and not the number of datasets in $B(T)_\phi$. To this end, we define an auxiliary function size where $\text{size}(B(T)_\phi) = [a, b]$ such that $a = \min\{|T'| : T' \in B(T)_\phi\}$ and $b = \max\{|T'| : T' \in B(T)_\phi\}$.

Then, we define the cost of splitting on ϕ as follows (recall that the operators use interval arithmetic):

$$\text{cost}^a(B(T), \phi) = \text{size}(B(T)_\phi) \times \text{imp}^a(B(T)_\phi) + \text{size}(B(T)_{-\phi}) \times \text{imp}^a(B(T)_{-\phi}) \quad (9)$$

Since size and imp^a return intervals, cost^a will be an interval, as well.

Best split. To find the set of best predicates, we identify the *least upper bound (lub)* of any predicate’s cost. Then, any predicate whose cost overlaps with *lub* will be a member of the set of best predicates, too. Formally, $\text{lub} = \min_{\phi \in \Phi} b_\phi$, where $\text{cost}^a(\phi) = [a_\phi, b_\phi]$. Then, we define $\text{split}^a(B(T)) = \{\phi \in \Phi \mid a_\phi \leq \text{lub}\}$.

Inference. Finally, for inference, we evaluate every predicate computed by split^a on x and collect all possible prediction labels. Intuitively, we break the problem into two pieces: first, we evaluate all predicates ϕ that satisfy $\phi(x)$ (i.e., when x is sent down the left branch of the tree), and then predicates that satisfy $\neg\phi(x)$, (i.e., when x is sent down the right branch of the tree). Formally, we compute:

$$\text{infer}^a(B(T), \Phi^a, x) = \underbrace{\bigcup_{\phi(x)} \text{argmax}_i \text{pr}_i^a(B(T)_\phi)}_{\text{labels for predicates } \phi \text{ s.t. } \phi(x)} \cup \underbrace{\bigcup_{\neg\phi(x)} \text{argmax}_i \text{pr}_i^a(B(T)_{-\phi})}_{\text{labels for predicates } \phi \text{ s.t. } \neg\phi(x)} \quad (10)$$

where the range of \cup is over predicates in Φ^a . Since our goal is to prove robustness, we only care whether $|\text{infer}^a(B(T), \Phi^a, x)| = 1$, i.e., all datasets produce the same prediction.

Theorem 2. If $|\text{infer}^a(B(T), \Phi^a, x)| = 1$, where $\Phi^a = \text{split}^a(B(T))$, then x is robust (Eq. (4)).

Example 4.6. Recall Fig. 1 with bias model $B = \text{FLIP}_1^g$, where g targets Black people with \times label. $\text{split}^a(B(T^{\text{ex}}))$ returns the singleton set $\Phi^a = \{\text{score} \leq 3\}$. Then, given input $x = \langle \text{race} = \text{Black}, \text{score} = 7 \rangle$, $\text{infer}^a(B(T^{\text{ex}}), \Phi^a, x) = \{\checkmark\}$, since $\text{pr}_\checkmark^a(B(T^{\text{ex}})_{\text{score} > 3}) = [0.8, 1]$, which is greater than $\text{pr}_\times^a(B(T^{\text{ex}})_{\text{score} > 3}) = [0, 0.2]$. Therefore, the learner is robust on x .

5 Experimental evaluation

We implement our certification technique in C++ and call it Antidote-P, as it extends Antidote [16] to programmable bias models. To learn trees with depth > 1 , we apply the presented procedure recursively. We use Antidote’s *disjunctive domain*, which is beneficial for certification [16] but requires a large amount of memory because it keeps track of many different datasets on each decision-tree path. We evaluate on Adult Income [17] (training $n=32,561$), COMPAS [29] ($n=4629$), and Drug Consumption [20] ($n=1262$). A fourth dataset, MNIST 1/7 ($n=13,007$), is in the Appendix. For all datasets, we use the standard train/test split if one is provided; otherwise, we create our own train/test splits, which are available in our code repository at <https://github.com/annapmeyer/antidote-p>.

Table 1: Certification rates for various bias models. Targeted bias models use predicates ($race = \text{Black}$ and $label = \text{positive}$) for COMPAS and ($gender = \text{female}$ and $label = \text{negative}$) for Adult Income. Composite models show cumulative bias, e.g., 0.2% MISS + FAKE bias equates to 0.1% bias of each MISS and FAKE. Empty entries indicate tests that failed due to memory constraints (96GB).

Bias type	Dataset	Bias amount as a percentage of training set					
		0.05	0.1	0.2	0.4	0.7	1.0
MISS (missing data)	Drug Consumption	94.5	94.5	94.5	94.5	85.1	85.1
	COMPAS	89.0	81.9	52.9	45.3	9.3	9.2
	Adult Income (AI)	96.0	86.9	72.8	60.9		
	COMPAS targeted	89.0	89.0	81.9	52.9	47.8	42.3
	AI targeted	98.8	97.2	86.6	73.0	62.0	31.6
FLIP (label-flipping)	Drug Consumption	94.5	94.5	94.5	92.1	85.1	7.1
	COMPAS	81.9	71.5	47.8	20.6	3.0	3.0
	Adult Income	95.8	72.9	70.2	34.8		
	COMPAS targeted	89.0	81.9	71.5	50.5	43.2	24.2
	AI targeted	98.6	97.0	74.4	71.0	45.4	25.8
MISS + FAKE (missing + fake)	Drug Consumption	94.5	94.5	94.5	94.5	85.1	85.1
	COMPAS	81.9	76.2	52.9	43.2	9.3	9.3
	Adult Income	96.0	95.6	72.8	68.3	36.2	
MISS + FLIP (missing + label-flipping)	Drug Consumption	94.5	94.5	92.1	92.1	85.1	38.0
	COMPAS	81.9	71.5	50.5	41.6	9.3	3.0
	Adult Income	95.9	74.3	71.1	49.0		
Bias-set size color scheme		< 10 ¹⁰	< 10 ⁵⁰	< 10 ¹⁰⁰	< 10 ⁵⁰⁰	> 10 ⁵⁰⁰	infinite

For each dataset, we choose the smallest tree depth where accuracy improves no more than 1% at the next-highest depth. For Adult Income and MNIST 1/7, this threshold is depth 2 (accuracy 83% and 97%, respectively); for COMPAS and Drug Consumption it is depth 1 (accuracy 64% and 76%, respectively). We run additional experiments on COMPAS and Drug Consumption at depths 2 and 3 to evaluate how tree depth influences Antidote-P’s efficiency (see Appendix).

A natural baseline is enumerating all datasets in the bias set but that is infeasible—see bias-set sizes in Table 1. To our knowledge, our technique (extended from [16]), is the only method to certify bias robustness of decision-tree learners.

5.1 Effectiveness at certifying robustness

Table 1 shows the results. Each entry in the table indicates the percentage of test samples for which Antidote-P can prove robustness with a given bias model and the shading indicates the size of the bias set, $|B(T)|$. We see that even though the perturbation sets are very large—sometimes infinite—we are able to certify robustness for a significant percentage of elements.

By dataset. Certification rates vary from 98.8% robustness for MISS_{0.05%} on Adult Income (i.e., the predictions of 98.8% of the points in the test set do not change if up to 0.05% new points are added to the training set) to just 3% robustness for FLIP_{1%} on COMPAS. Even for a single bias model, the certification rates vary widely: under FLIP_{0.2%}, we can verify 94.5% of samples as robust for Drug Consumption, but only 70.2% for Adult Income and 47.8% for COMPAS. We posit that these differences stem from inherent properties of the datasets. The normalized cost of the optimal top-most split is 0.30 for Adult Income, 0.35 for Drug Consumption, and 0.45 for COMPAS (recall that a lower cost corresponds to greater information gain). As a result, biasing a fixed percentage of data yields greater instability for COMPAS, since the data already exhibited poorer separation.

By bias model. There are also differences in certification rates between bias models. FLIP is more destructive to robustness because flipping a single label results in a symmetric difference of 2 from the original dataset (as if we removed an element from the set and then inserted a new one with a flipped label), while adding a single item results in a symmetric difference of 1.

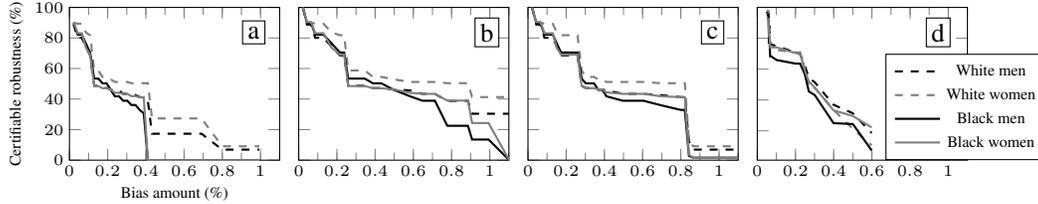


Figure 2: Left to right: Certifiable robustness by demographic group on (a) COMPAS under FLIP; (b) COMPAS under FLIP^g where $g \triangleq (\text{race} = \text{Black} \wedge \text{label} = \text{positive})$; (c) COMPAS under FLIP^g where $g \triangleq (\text{race} = \text{White} \wedge \text{label} = \text{negative})$; (d) Adult Income under FLIP.

The composite bias models display similar dataset- and bias model-based trends. Notably, MISS + FAKE yields a lower certifiable-robustness rate than FLIP. E.g., under $\text{FLIP}_{0.1\%}$, we can certify 71.5% of COMPAS test samples as robust. But for $\text{MISS}_{0.1\%} + \text{FAKE}_{0.1\%}$ (that is, 0.2% bias total), we are only able to certify 50.5% of test samples as robust. This shows that FLIP is a useful modeling tool for situations where we believe the features of all data points to be trustworthy, but suspect that some labels may be incorrect. The targeted bias models allow for greater certification rates than the non-targeted versions; this is expected because they result in smaller bias sets.

In summary, Antidote-P can effectively certify robustness across a variety of bias models, but its success depends on properties of the dataset such as separability.

5.2 Demographic variations

We evaluated differences in certifiable-robustness rates across demographic groups in all three datasets. We present results from COMPAS and Adult Income in Fig. 2 (results for Drug Consumption are in the Appendix; they are less interesting due to a lack of representation in the dataset).

COMPAS. Fig. 2a shows that under FLIP, White women are robust at a higher rate than any other demographic group, and that Black men and women are the least robust. Notably, for $\text{FLIP}_{0.4\%}$, we are able to certify robustness for 50.4% of White women, but 0% of Black people. There is also a significant gap between White women and White men at this threshold (50.4% vs. 38.8%). We can explain the gaps in certification rates of different subgroups by looking at the training data. In the COMPAS dataset, the same predicate provides the optimal split for every race-gender subgroup, but for White women the resulting split has cost = 0.41 versus cost = 0.46 for Black people. It is not clear whether this difference stems from sampling techniques or inherent differences in the population, but regardless, the end result is that **predictions made about Black people from decision trees trained on COMPAS are more likely to be vulnerable to data bias.**

To validate that the disparities in certifiable-robustness rates are due to inherent dataset properties rather than the abstraction, we performed random testing by perturbing the COMPAS dataset to try to find robustness counterexamples, i.e., datasets in the bias set that yield conflicting predictions on a given input. We found more bias counterexamples to robustness for Black people than for White people, which is further evidence for our claim that the robustness disparities are inherent to the dataset.

Targeted bias models (COMPAS). If we choose $g \triangleq (\text{race} = \text{Black} \wedge \text{label} = \text{positive})$ (Fig. 2b) in FLIP^g to model the real-world situation where structural or individual racism can lead to increased policing and convictions among Black people in the U.S., then there are generally higher robustness rates at moderate bias levels (e.g., ~50% robustness for all demographic groups at 0.4% poisoning). However, as the amount of bias increases, a gap between White and Black certification rates emerges (in exact terms, 32.9% of White test samples are certifiably robust versus 0% of Black test samples starting at 1.1% bias and continuing through, at least, 10.8% bias). It is unclear whether this trend stems from inherent dataset properties, or is due to the over-approximate nature of the abstraction.

By contrast, using $g \triangleq (\text{race} = \text{White} \wedge \text{label} = \text{negative})$ (Fig. 2c) to describe that White people may be under-policed or under-convicted due to White privilege nearly eliminates discrepancies between demographic groups. In particular, Black men (previously the least-robust subgroup) are the most robust of any population. FLIP^g and $\text{FLIP}^{g'}$ differ only on how they describe societal inequities: are White people under-policed, or are Black people over-policed? However, the vast differences in

demographic-level robustness rates between FLIP^g and $\text{FLIP}^{g'}$ shows that **the choice of predicate is crucial when using targeted bias models**. More experimentation is needed to understand why these results occur, and how consistent they are across different train/test splits of the data. However, our preliminary results indicate that Antidote-P could be a useful tool for social scientists to understand how data bias can affect the reliability of machine-learning outcomes.

Adult Income. Fig. 2d shows robustness by demographic group for FLIP. We see that Black men have about a 5% lower robustness rate than other demographic groups and that at higher bias levels, White women also have about a 5% lower robustness rate than White men or Black women. Using FLIP^g where $g = (\text{race} = \text{Female} \wedge \text{Label} = \text{negative})$ led to similar results (see Appendix).

6 Conclusions and broader impacts

We saw that our decision-tree-learner abstraction is able to verify pointwise robustness over large and even infinite bias sets. These guarantees permit increased confidence in the trees’ outputs because they certify that data bias has not affected the outcome (within a certain threshold). However, a couple of tricky aspects—and ones that we do not attempt to address—are knowing whether the assumptions underlying the bias model are correct, or whether our bias framework is even capable of representing all instances of real-world bias. If the user does not specify the bias model faithfully, then any proofs may not be representative. Also, our tool only certifies robustness, not accuracy. Therefore, it may certify that a model will always output the *wrong* label on a given data point. This behavior is linked to a shortcoming of many machine-learning audits: our tool cannot determine what is an appropriate use of machine learning. Machine learning is often used to promote and legitimize uses of technology that are harmful or unethical. In particular, we want to call out our use of the COMPAS dataset: we feel that it is illustrative to show how certifiable-robustness rates can vary widely between different demographic groups and be sensitive to subtle shifts in the bias model. However, this use should not be taken as an endorsement for the deployment of recidivism-prediction models.

Another limitation is that our framework can only certify decision-tree learners. In practice, many machine learning applications use more sophisticated algorithms that we do not address here. Future work to generalize our ideas to other machine learning architectures would increase the utility of this style of robustness certification.

Returning to our work, Antidote-P has a place in data scientists’ tool-kits as a powerful technique to understand robustness, and potential vulnerabilities, of data bias in decision-tree algorithms. An important direction for future work is to develop effective techniques for falsification of robustness (i.e., techniques to find minimal dataset perturbations that break robustness). We performed initial experiments in this area using brute-force techniques (i.e., randomly perturb data points, train a new decision tree and see whether the test sample’s classification changes under the new tree—see the Appendix for more details). The results were promising in that we were able to find counter-examples to robustness for some data points, but there remain many data points that are neither certifiably robust via Antidote-P nor falsified as robust using random testing. Random testing was an interesting proof of concept, but we recommend that the future focus be on developing techniques to identify these dataset perturbations in a more scalable and intelligent way. Other future work could also improve our approach’s utility through tightening the analytical bounds, such as by abstracting over a more complex domain than intervals.

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A Additional details and definitions

Throughout the appendices, we use square brackets, rather than braces, to denote composite bias models: this is to emphasize that the transformers are ordered, and that alternate orderings often result in distinct bias sets.

Filtering composite bias models. Filtering a composite bias model requires us to apply filter piece-wise, i.e., $[\text{MISS}_m^{g_1}, \text{FLIP}_l^{g_2}, \text{FAKE}_k^{g_3}](T)_\phi = [\text{MISS}_m^{g_1 \wedge \phi}, \text{FLIP}_l^{g_2}, \text{FAKE}_k^{g_3}](T_\phi)$.

pr^a for FAKE. Given c_i samples in T with label i , we use $k_i = \min(k, c_i)$ and then define

$$\text{pr}_i^a(\text{FAKE}_k(T)) = \left[\frac{c_i - k_i}{|T| - k_i}, \frac{c_i}{|T| - \sum_{j \neq i} k_j} \right] \quad (11)$$

For the edge case where $c_i = |T|$ and $c_i \leq k$ for any i , we define $\text{pr}_j^a(T) = [0, 1]$ for all $j \in [1, n]$. A similar edge case applies, when necessary, to the composite definition.

Optimizing pr^a when g looks at the label. If g conditions on the label, then we can improve the precision of pr^a by defining each component individually. Suppose $g(x, y) = y \in S \wedge g'(x)$, where $S \subset \{1, \dots, n\}$ and g' is a predicate that only conditions on features.

For MISS_m^g , we define

$$\text{pr}_i^a(\text{MISS}_m^g) = \begin{cases} \left[\frac{c_i}{|T|}, \frac{c_i+m}{|T|+m} \right] & \text{if } i \in S \text{ and } |S| = 1 \\ \left[\frac{c_i}{|T|+m}, \frac{c_i+m}{|T|+m} \right] & \text{if } i \in S \text{ and } |S| \geq 2 \\ \left[\frac{c_i}{|T|+m}, \frac{c_i}{|T|} \right] & \text{else} \end{cases} \quad (12)$$

For FLIP_l^g , we use $l_{a_i} = \min(l, |\{(x, y) \in T \mid y = i \wedge g(x, y)\}|)$ and $l_{b_i} = \min(l, |\{(x, y) \in T \mid y \neq i \wedge g(x, y)\}|)$. Then, we define

$$\text{pr}_i^a(\text{FLIP}_l^g) = \begin{cases} \left[\frac{c_i - l_{a_i}}{|T|}, \frac{c_i}{|T|} \right] & \text{if } i \in S \text{ and } |S| = 1 \\ \left[\frac{c_i - l_{a_i}}{|T|}, \frac{c_i + l_{b_i}}{|T|} \right] & \text{if } i \in S \text{ and } |S| \geq 2 \\ \left[\frac{c_i}{|T|}, \frac{c_i + l_{b_i}}{|T|} \right] & \text{else} \end{cases} \quad (13)$$

For FAKE_k^g , we use $k_{a_i} = \min(k, |\{(x, y) \in T \mid y = i \wedge g(x, y)\}|)$ and $k_{b_i} = \min(k, |\{(x, y) \in T \mid y \neq i \wedge g(x, y)\}|)$. Then, we define

$$\text{pr}_i^a(\text{FAKE}_k^g) = \begin{cases} \left[\frac{c_i - k_{a_i}}{|T| - k_{a_i}}, \frac{c_i}{|T|} \right] & \text{if } i \in S \text{ and } |S| = 1 \\ \left[\frac{c_i - k_{a_i}}{|T| - k}, \frac{c_i}{|T| - k_{b_i}} \right] & \text{if } i \in S \text{ and } |S| \geq 2 \\ \left[\frac{c_i}{|T|}, \frac{c_i}{|T| - k_{b_i}} \right] & \text{else} \end{cases} \quad (14)$$

We prove that the above definitions are sound and precise in Appendix D. If desired, the above definitions can be pieced together to provide a more precise definition for composite bias models. However, we limit ourselves to just the singleton transformers because notation becomes very messy, as we have to keep track of many variables indicating how many data elements satisfy the various conditions.

pr^a for composite bias models with multiple versions of the same transformer. If a bias model contains multiple instances of the same transformer, e.g., $B = [\text{FLIP}_{l_1}^{g_1}, \text{FLIP}_{l_2}^{g_2}]$, we can combine everything into a single transformer. Formally, given

$$B = [\text{MISS}_{m_1}^{g'_1}, \dots, \text{MISS}_{m_j}^{g'_j}, \text{FLIP}_{l_1}^{g'_{j+1}}, \dots, \text{FLIP}_{l_p}^{g'_{j+p}}, \text{FAKE}_{k_1}^{g'_{j+p+1}}, \dots, \text{FAKE}_{k_q}^{g'_{j+p+q}}] \quad (15)$$

we define

$$m = m_1 + \dots + m_j$$

$$\begin{aligned}
g_1 &= g'_1 \vee \dots \vee g'_j \\
l_i &= \min(l_1 + \dots + l_p, |\cup_{i \in [1,p]} T_{g_{j+i}(x,y) \wedge y=i}|) \\
g_2 &= g'_{j+1} \vee \dots \vee g'_{j+p} \\
k_i &= \min(k_1 + \dots + k_p, |\cup_{i \in [1,q]} T_{g_{j+p+i}(x,y) \wedge y=i}|)
\end{aligned}$$

and

$$g_3 = g'_{j+p+1} \vee \dots \vee g'_{j+p+1}$$

Then, we can use the formula shown in Equation 8 to compute pr^a . We show in Appendix C that these definitions are sound.

Size (size). We define $\text{size}(\text{MISS}_m^g) = [|T|, |T| + m]$, $\text{size}(\text{FLIP}_l^g) = [|T|, |T|]$, and $\text{size}(\text{FAKE}_k^g) = [|T| - k, |T|]$. Putting this all together, we have $\text{size}([\text{MISS}_m^{g_1}, \text{FLIP}_l^{g_2}, \text{FAKE}_k^{g_3}]) = [|T| - k, |T| + m]$.

B Proof of optimal composition of transformers

As stated in § 3, when composing transformers we want to apply them in an order that results in the largest composite bias model. To illustrate the concept of composite bias models' relative size, consider $B = [\text{MISS}_1^{g_1}, \text{FLIP}_1^{g_2}]$ and $B' = [\text{FLIP}_1^{g_2}, \text{MISS}_1^{g_1}]$ where $g_1 \triangleq (\text{gender}=\text{female} \wedge \text{label}=1)$ and $g_2 \triangleq (\text{gender}=\text{female})$. I.e., B adds one data point subject to g_1 and then flips the label of one data point subject to g_2 , whereas B' performs these two operations in the opposite order. Under B , we can use MISS to add the data point $(x, y) = \langle \text{gender}=\text{female}, \text{label}=1 \rangle$ and then use FLIP to change y to 0. However, under B' , we cannot alter the point that MISS adds, so B and B' are not equivalent. In this case, B can construct every dataset that B' can construct (but not vice-versa), so we write $B' \subset B$ and say that B is larger than B' .

First we consider the case when there are multiple transformers of the same type.

Theorem 3. *The bias models $B_1 = [\text{MISS}_{m_1}^{g_1}, \text{MISS}_{m_2}^{g_2}]$ and $B_2 = [\text{MISS}_{m_2}^{g_2}, \text{MISS}_{m_1}^{g_1}]$ are equivalent (and likewise for FAKE and FLIP, as long as no FLIP predicate conditions on the label).*

Proof. Missing data. The choice of what missing data to add has no bearing on what is already in (or not in) the dataset. Thus we can add m_1 elements that satisfy g_1 followed by m_2 elements that satisfy g_2 , or do the operators in the reverse order, but the end result is the same.

Label-flipping. Suppose $B = [\text{FLIP}_{l_1}^{g_1}, \text{FLIP}_{l_2}^{g_2}]$, where g_1 and g_2 do not condition on the label. We want to show that B is equivalent to $B' = [\text{FLIP}_{l_2}^{g_2}, \text{FLIP}_{l_1}^{g_1}]$.

Consider an arbitrary $T' \in B(T)$. Each data point in T' is either (1) untouched by $\text{FLIP}_{l_1}^{g_1}$ and $\text{FLIP}_{l_2}^{g_2}$, (2) flipped only by $\text{FLIP}_{l_1}^{g_1}$, (3) flipped only by $\text{FLIP}_{l_2}^{g_2}$, or (4) flipped by both $\text{FLIP}_{l_1}^{g_1}$ and $\text{FLIP}_{l_2}^{g_2}$. If (1), clearly this is obtainable by B' since we did nothing. If (2), then since the data point is untouched by $\text{FLIP}_{l_2}^{g_2}$, the data point can be flipped uninterrupted by $\text{FLIP}_{l_1}^{g_1}$ (similarly for (3)). If (4), then – since neither g_1 nor g_2 conditions on the label nor specifies what the new label can be – we can still flip the label twice and end up with the same configuration. The same arguments hold had we started with $T'' \in B'$. Therefore, B and B' are equivalent.

Fake data. The argument for fake data is similar. □

We can extend the proof of Theorem 3 to arbitrarily many transformers of the same type.

Note that if FLIP conditions on the label, this proof does not hold. To continue with the terminology from the proof, if $g_i \triangleq (\text{label}=a)$, then applying $\text{FLIP}_{l_j}^{g_j}$ first to some element (x, a) yields (x, a') , which may no longer eligible to be flipped by $\text{FLIP}_{l_i}^{g_i}$.

Next, we show that there is an optimal way to compose transformers of different types. We define *optimal* as largest, that is, some B' is optimal compared to B if $B \subseteq B'$. In other words, this notation says that every dataset created by B can also be created by B' . For the next theorem and its proof we assume there is only one instance of each transformer type; however, in conjunction with Theorem 3 we can extend it to include multiple instances of the same transformer type.

Theorem 4. *$B = [\text{MISS}, \text{FLIP}, \text{FAKE}]$ is the optimal order to apply the transformers MISS, FLIP, and FAKE (i.e., any other ordering B' of these transformers will satisfy $B' \subseteq B$).*

Proof. We will show that other orderings of MISS, FLIP, FAKE do not produce any biased datasets that do not also occur in [MISS, FLIP, FAKE]. For conciseness, we will write [MISS, FLIP, FAKE] as OPT.

1. [MISS, FAKE, FLIP]: We consider the set of datasets S achieved after applying MISS, FAKE, and then FLIP. Fix an arbitrary $T' \in S$. T' was constructed from T by some sequence of adding, removing, and flipping data points. We have these categories for (potential) data points in T' : (1) untouched data points, (2) added data points, (3) added then removed data points, (4) added then flipped data points, (5) removed data points, and (6) flipped data points. (1), (2), (5), and (6) apply single (or no) operators, so clearly are also attainable through OPT. MISS occurs before both FLIP and FAKE in OPT, so (3) and (4) are attainable, as well.

2. [FLIP, MISS, FAKE]: We consider the set of datasets S achieved after applying FLIP, MISS, and then FAKE. Fix an arbitrary $T' \in S$. T' was constructed from T by some sequence of flipping, adding, and removing data points. We have these categories: (1) untouched data points, (2) flipped data points, (3) flipped then removed data points, (4) added data points, (5) added then removed data points, and (6) removed data points. (1), (2), (4), and (6) apply single (or no) operators, so clearly they are also attainable through OPT. Since flipping and adding each come before removing in OPT, (3) and (5) are obtainable as well.

3. [FLIP, FAKE, MISS]: We consider the set of datasets S achieved after applying FLIP, FAKE, and then MISS. Fix an arbitrary $T' \in S$. T' was constructed from T by some sequence of flipping, removing, and adding data points. We have these categories: (1) untouched data points, (2) flipped data points, (3) flipped then removed data points, (4) removed data points, and (5) added data points. (1), (2), (4), and (5) apply single (or no) operators, so clearly they are also attainable through OPT. Since flipping comes before removing in OPT, (3) is obtainable as well.

4. [FAKE, MISS, FLIP]: We consider the set of datasets S achieved after applying FAKE, MISS, and then FLIP. Fix an arbitrary $T' \in S$. T' was constructed from T by some sequence of removing, adding, and flipping data points. We have these categories: (1) untouched data points, (2) removed data points, (3) added data points, (4) added then flipped data points, and (5) added data points. (1), (2), (4), and (5) apply single (or no) operators, so clearly they are also attainable through OPT. Since flipping comes before removing in OPT, (3) is obtainable as well.

5. [FAKE, FLIP, MISS]: We consider the set of datasets S achieved after applying FAKE, FLIP, and then MISS. Fix an arbitrary $T' \in S$. T' was constructed from T by some sequence of removing, flipping, and adding data points. We have these categories: (1) untouched data points, (2) removed data points, (3) flipped data points, and (4) added data points. Each of these apply single (or no) operators, so clearly they are also attainable through OPT.

We were not able to construct a dataset not also in OPT through any other ordering of the operators, therefore, OPT is optimal. \square

C Proofs of soundness

Proof of Theorem 4.2. pr^a is sound.

Proof. We show MISS as a simple example to illustrate our approach, and then we show the proof for composite bias. We omit the proofs for FLIP and FAKE because they (like MISS) are special cases of composite.

Missing data. Given a dataset T with n classes, suppose our bias set is $\text{MISS}_m^g(T)$. Furthermore, suppose that c_i samples in T have label i . We define $m_i \in [0, m]$ to be the number elements we add with label i , and $m'_i = \sum_{j \neq i} m_j$. Then, we can write the proportion of i 's as a function

$$F(m_i, m'_i) = \frac{c_i + m_i}{|T| + m_i + m'_i} \quad (16)$$

The partial derivatives of F have values $\frac{\delta F}{\delta m_i} > 0$ and $\frac{\delta F}{\delta m'_i} < 0$ over the entire domain $[0, m]$, therefore, any conclusions we draw over \mathbb{R} will also apply over the discrete integer domain. Therefore, to minimize F we choose $m_i = 0$ and $m'_i = m$, and do the reverse to maximize F . Thus, $F_{\min} = \frac{c_i}{|T|+m}$ and $F_{\max} = \frac{c_i+m}{|T|+m}$. Since $[F_{\min}, F_{\max}] \subseteq \text{pr}_i^a(\text{MISS}_m^g(T))$, pr^a is sound.

Composite. Given a dataset T with n classes, suppose that our bias model is $B = [\text{MISS}_{m_i}^{g_1}, \text{FLIP}_{l_i}^{g_2}, \text{FAKE}_{k_i}^{g_3}]$. Furthermore, suppose that in T , c_i samples have label i .

First, we consider how many elements we can add, flip, or remove of each label. Under MISS, we can add m of label i for all i . Under FLIP, we can flip up to l labels from label i to some $j \neq i$, assuming that there are at least l elements (x, y) that satisfy $g_2(x, y)$ and $y = i$. Similarly, the maximum number of labels we can flip from any label $j \neq i$ to i is bounded both by l and by the number of elements that satisfy g_2 and have label j . Note that the elements we can flip are not just limited to T : we can also flip any of the newly-added m elements. Formally, we define $l_{a_i} = \min(|\{(x, y) \in T \mid y = i \wedge g_2(x, y)\}| + m, l)$, and $l_{b_i} = \min(|\{(x, y) \in T \mid y \neq i \wedge g_2(x, y)\}| + m, l)$. Similarly, under FAKE we must also consider the elements added by MISS and those flipped from $j \neq i$ to i by FLIP. Therefore, we define $k_{a_i} = \min(|\{(x, y) \in T \mid y = i \wedge g_3(x, y)\}| + m + l_{b_i}, k)$ and $k_{b_i} = \min(|\{(x, y) \in T \mid y \neq i \wedge g_3(x, y)\}| + m + l_{a_i}, k)$.

We will show that under the specified bias model, the proportion of i 's in the dataset is always between $(c_i - l_{a_i} - k_{a_i})/(|T| - k_{a_i} + m)$ and $(c_i + m + l_{b_i})/(|T| - k_{b_i} + m)$.

Intuitively, we consider how to modify the proportion of i 's in T . This proportion decreases by (1) flipping elements from class i to some class $j \neq i$, (2) removing elements of class i , and (3) adding elements of a class other than i . Therefore, the fraction is minimized by doing (1), (2), and (3) as much as the bias model allows.

Formally, let $m_i \in [0, m]$ be the number of elements added with label i , $m'_i \in [0, m] = \sum_{j \neq i} m_j$, $l_i \in [0, l_{a_i}]$ be the number of elements flipped from class i to any other class, $l'_i \in [0, l_{b_i}] = \sum_{j \neq i} l_j$, $k_i \in [0, k_{a_i}]$ be the number of elements removed with label i , $k'_i \in [0, k_{b_i}] = \sum_{j \neq i} k_j$.

Then, we can write the proportion of i 's as a function

$$F(m_i, m'_i, l_i, l'_i, k_i, k'_i) = \frac{c_i + m_i - l_i + l'_i - k_i}{|T| + m_i + m'_i - k_i - k'_i} \quad (17)$$

We consider the partial derivative of F with respect to each variable. For all input in the domain, we have $\frac{\delta F}{\delta m_i} > 0$, $\frac{\delta F}{\delta m'_i} < 0$, $\frac{\delta F}{\delta l_i} < 0$, $\frac{\delta F}{\delta l'_i} > 0$, $\frac{\delta F}{\delta k_i} < 0$, and $\frac{\delta F}{\delta k'_i} > 0$.

Note that each partial derivatives is monotone over all values in the domain. Thus, they are also monotone over integers, so any conclusions we yield over the real numbers can be relaxed to integers, as well. To minimize F , we will maximize each variable whose partial derivative is negative, and minimize each variable whose partial derivative is positive. That is, we choose $m_i = 0$, $m'_i = m$, $l_i = l_{a_i}$, $l'_i = 0$, $k_i = k_{a_i}$, and $k'_i = 0$ to minimize F , yielding

$$F_{\min} = \frac{c_i - l_{a_i} - k_{a_i}}{|T| + m - k_{a_i}}$$

Conversely, to maximize F we maximize each variable whose partial derivative is positive and minimize each variable whose partial derivative is negative, yielding

$$F_{\max} = \frac{c_i + m + l_{b_i}}{|T| + m - k_{b_i}}$$

$[F_{\min}, F_{\max}] \subseteq \text{pr}_i^a(B(T))$, therefore, pr^a is sound.

Multiple composite bias models. Suppose

$$B = [\text{MISS}_{m_1}^{g'_1}, \dots, \text{MISS}_{m_j}^{g'_j}, \text{FLIP}_{l_1}^{g'_{j+1}}, \dots, \text{FLIP}_{l_p}^{g'_{j+p}}, \text{FAKE}_{k_1}^{g'_{j+p+1}}, \dots, \text{FAKE}_{k_q}^{g'_{j+p+q}}]$$

and

$$B' = \left[\text{MISS}_{\sum_{i \in [1, j]} m_i}^{\bigvee_{i \in [1, j]} g_i}, \text{FLIP}_{\sum_{i \in [1, p]} l_i}^{\bigvee_{i \in [j+1, j+p]} g_i}, \text{FAKE}_{\sum_{i \in [1, q]} k_i}^{\bigvee_{i \in [j+p+1, j+p+q]} g_i} \right]$$

We want to show that if $(x, y) \in T \cup T'$ is altered by B , it can be altered by B' (in other words, we want to show that $B \subseteq B'$).

Case 1: (x, y) was added by a transformer $\text{MISS}_{m_i}^{g_i}$ for $i \in [1, j]$. Therefore, $g_i(x)$ and $(\bigvee_{i \in [1, j]} g_i)(x)$, as well, so x can be added by $\text{MISS}_{\sum_{i \in [1, j]} m_i}^{\bigvee_{i \in [1, j]} g_i}$.

Case 2: $(x, y) \in T$ was flipped to $(x, y') \in T'$ by $\text{FLIP}_{l_i}^{g_{j+i}}$ for $i \in [1, p]$. This means that $g_{j+i}(x)$, so by extension, $(\bigvee_{i \in [j+1, j+p]} g_i)(x)$, which means that x 's label can be flipped by $\text{FLIP}_{\sum_{i \in [1, p]} l_i}^{\bigvee_{i \in [j+1, j+p]} g_i}$.

Case 3: $(x, y) \in T$ was removed by $\text{FLIP}_{k_i}^{g_{j+p+i}}$ for $i \in [1, q]$. This means that $g_{j+p+i}(x)$, and thus $\bigvee_{i \in [j+p+1, j+p+q]} g_i(x)$. Therefore (x, y) can be removed by $\text{FAKE}_{\sum_{i \in [1, q]} k_i}^{\bigvee_{i \in [j+p+1, j+p+q]} g_i}$.

To conclude, any modification to T that we can make under B is also attainable under B' , therefore, if B' is sound, then B is sound, as well. \square

Proposition 1. *Abstract filtering is sound.*

Proof. To prove soundness for filtering, we need to show that $T' \in B(T) \implies T'_\phi \in B(T)_\phi$.

Missing data. Consider $T' \in \text{MISS}_m^g(T)$. Since $T' \in \text{MISS}_m^g$, we have $T' = T \cup S$ where $|S| \leq m$ and $\forall (x, y) \in S. g(x, y)$. Therefore, $T'_\phi = T_\phi \cup S_\phi$. Since $S_\phi \subseteq S$, then $|S_\phi| \leq m$ and $\forall (x, y) \in S_\phi. g(x, y)$. then $T_\phi \cup S_\phi \in \text{MISS}_m^{g \wedge \phi}$, satisfying the claim.

Label flipping. Consider $T' \in \text{FLIP}_l^g(T)$. Since $T' \in \text{FLIP}_l^g$, we know that $T' = R \cup S$ where $R \subseteq T$, $|S| \leq l$, and $T = R \cup \{(x, y) \mid (x, y') \in S\}$. Additionally, we have $(x, y) \in S \implies g(x, y)$. Consider $T'_\phi = R_\phi \cup S_\phi$. Since $R \subseteq T$, then $R_\phi \subseteq T_\phi$. Since $(x, y) \in S \implies (x, y') \in T$ and ϕ does not condition on the label, then $(x, y) \in S_\phi \implies (x, y') \in T_\phi$. Since $S_\phi \subseteq S$, we have $|S_\phi| \leq l$. In total, this means that $T_\phi \in \text{FLIP}_l^g(T_\phi)$.

Fake data. See [16].

Composite. Suppose $B = [\text{MISS}_m^{g_1}, \text{FLIP}_l^{g_2}, \text{FAKE}_k^{g_3}]$. We want to show that $B(T)_\phi \subseteq B'(T_\phi)$, where $B' = [\text{MISS}_m^{g_1 \wedge \phi}, \text{FLIP}_l^{g_2}, \text{FAKE}_k^{g_3}]$. Consider $T' \in B(T)$. Each $(x, y) \in T'$ satisfies either (i) $(x, y) \in T$, (ii) $(x, y') \in T$ (for $y' \neq y$), or (iii) $(x, y') \notin T$. If (i), then $x \in T'_\phi \implies x \in T_\phi$, and likewise for $\neg\phi$. If (ii), $x \in T'_\phi \implies x \in T_\phi$ (since ϕ ignores the label), and likewise for $\neg\phi$. If (iii), the x was added by MISS . $x \in T'_\phi \implies \phi(x) \implies x$ can be added by $\text{MISS}_m^{g_1 \wedge \phi}$, and if $x \notin T_\phi$, this means that x cannot be added by $\text{MISS}_m^{g_1 \wedge \phi}$. Finally, there is a fourth category of elements: those in $T \setminus T'$. If $x \in T \setminus T'$, then x was removed by FAKE . If $\phi(x)$, then x can be removed from T to make T' , otherwise, x cannot be removed from T , so it must also be contained in T' .

Thus we have shown that T'_ϕ can be constructed from T_ϕ using B' , therefore, filtering is sound. \square

Proposition 2. *imp^a is sound*

Proof. To show that imp^a is sound, we must show that $T' \in B(T) \implies \text{imp}(T') \in \text{imp}^a(B(T))$. By § 4.2, $\text{pr}(T') \in \text{pr}(B(T))$. It follows from interval arithmetic $\text{imp}(T') \in \text{imp}^a(B(T))$. \square

Proposition 3. *size is sound.*

Proof. Given a bias model $B = [\text{MISS}_m^{g_1}, \text{FLIP}_l^{g_2}, \text{FAKE}_k^{g_3}]$ and a dataset T , we can write the size of $T' \in B(T)$ as $|T| - k' + m'$, where $k' \in [0, k]$ and $m' \in [0, m]$. Clearly, $|T| - k' + m'$ is minimized by choosing $k' = k$ and $m' = 0$, and maximized by choosing $k' = 0$ and $m' = m$. Since $|T| - k \in \text{size}(B)$ and $|T| + m \in \text{size}(B)$, we see that size is sound. \square

Proposition 4. *cost^a is sound.*

Proof. We need to show that if $T' \in B(T)$ and $\phi \in \Phi$, then $\text{cost}(T', \phi) \in \text{cost}^a(B(T), \phi)$. By Proposition 1, we know that $T'_\phi \in B(T)_\phi$, which means that by Proposition 3, $|T'_\phi| \in \text{size}(B(T)_\phi)$, and similarly we can derive that $|T'_{\neg\phi}| \in \text{size}(B(T)_{\neg\phi})$. Additionally, by Proposition 2, we know that $\text{imp}(T'_\phi) \in \text{imp}^a(B(T)_\phi)$ and $\text{imp}(T'_{\neg\phi}) \in \text{imp}^a(B(T)_{\neg\phi})$. By the rules of interval arithmetic, if $a \in [a_0, a_1]$ and $b \in [b_0, b_1]$, then $ab \in [a_0, a_1] \times [b_0, b_1]$ and $a + b \in [a_0, a_1] + [b_0, b_1]$. Therefore we can conclude that $\text{cost}(T', \phi) \in \text{cost}^a(B(T), \phi)$, i.e., cost^a is sound. \square

Proposition 5. *split^a is sound.*

Proof. We want to show that if $T' \in B(T)$, then $\text{split}(T') \in \text{split}^a(B(T))$.

Suppose $\gamma = \text{split}(T')$. Then, $\forall \phi \in \Phi$, $\text{cost}(T', \phi) \geq \text{cost}(T', \gamma)$.

Define ϕ^* such that $\text{lub} = \text{ub}(\text{cost}(T, \phi^*))$, where ub takes the upper bound of an interval. Since $\phi^* \in \Phi$, $\text{cost}(T', \gamma) \leq \text{cost}(T', \phi^*)$. By Proposition 4, cost is sound, therefore $\text{cost}(T', \gamma) \in \text{cost}^a(B(T), \gamma)$ and $\text{cost}(T', \phi^*) \in \text{cost}^a(B(T), \phi^*)$. And thus we have $\text{lb}(\text{cost}^a(B(T), \gamma)) \leq \text{ub}(\text{cost}^a(B(T), \phi^*)) = \text{lub}$. Thus, $\gamma \in \text{split}^a(B(T))$. \square

Proof of Theorem 2.

Proof. Given $\Phi^a = \text{split}^a(B(T))$, if $|\text{infer}^a(B(T), \Phi^a, x)| = 1$, then we know

$$\exists y. \forall \phi \in \Phi^a. \begin{cases} \text{if } \phi(x) \text{ then } \text{argmax}_i \text{pr}_i^a(B(T)_\phi) = y \\ \text{if } \neg\phi(x) \text{ then } \text{argmax}_i \text{pr}_i^a(B(T)_{\neg\phi}) = y \end{cases}$$

Given $T' \in B(T)$, we know from Proposition 5 that $\text{split}(T') \in \text{split}^a(T) = \Phi^a(T)$. Therefore, $\text{infer}(T', \text{split}(T'), x) = y$, so the algorithm is robust on x . (Note that we defined $B(T)$ such that $T \in B(T)$, therefore, the original prediction is also y .) \square

D Precision

Intuitively, an abstraction is precise if the abstraction cannot be improved. Formally, our abstraction is precise iff, it is sound and given $\text{pr}_i^a(B(T)) = [a_i, b_i]$, then for each i there is some $T' \in B(T)$ such that $\text{pr}_i(T') = a_i$ and some $T'' \in B(T)$ such that $\text{pr}_i(T'') = b_i$.

Theorem 5. pr^a is precise for missing data, label-flipping, and fake data.

Proof. Missing data (non-targeted). In the proof of Theorem 4.2, we show that the minimum proportion of i 's is $\frac{c_i}{|T|+m}$ and the maximum proportion of i 's is $\frac{c_i+m}{|T|+m}$. Since these bounds are equal to pr^a 's minimum and maximum, the interval is precise.

Proofs for label-flipping and fake data (non-targeted) similarly follow from Theorem 4.2.

Targeted. Next, we will show that the definitions of pr^a provided in Equations 12-14 are precise. First, we must show that these definitions are sound (as soundness is a prerequisite for precision).

To show that Eq. (12) is sound, we need to consider three cases (we use S such that $g(x, y) = y \in S \wedge g'(x)$): first, $i \in S$ and $|S| = 1$. In this case, we can add up to m elements of class i , and no elements of class $j \neq i$. Therefore, the minimum proportion of i 's is the original proportion: $\frac{c_i}{|T|}$, and the maximum is $\frac{c_i+m}{|T|}$. Second, $i \in S$ and $|S| \geq 2$. In this case, we can add elements with label i but we can also add elements with label $j \neq i$. As such, the minimum proportion of i 's is achieved by adding m elements with label j , and the maximum proportion of i 's is achieved by adding m elements with label i . Therefore, the minimum proportion of i 's is $\frac{c_i}{|T|+m}$ and the maximum proportion of i 's is $\frac{c_i+m}{|T|+m}$.

Proofs of soundness for Equations 13 and 14 follow similarly.

To show precision, note that in the soundness proof we described exactly how to achieve the minimum and maximum bounds of pr_i^a . As such, we have shown that the pr^a definition in Eq. (12) is precise. (Similar conclusions can be drawn based on the proofs for label-flipping and fake data.) \square

pr^a is not precise for composite bias because the auxiliary variables $l_{a_i}, l_{b_i}, k_{a_i}$, and k_{b_i} are over-approximate. As a motivating example, consider a bias model $[\text{MISS}_1, \text{FLIP}_2^g]$ over dataset T with classes $\{0, 1\}$ where $c_0 = 2$, $c_1 = 5$, and $|\{(x, y) \in T : g(x, y) \wedge i = 0\}| = 1$. By definition, $l_{a_0} = 2$, yielding a minimum proportion of 0's to be $\frac{2-2}{7+2} = 0$. However, the precise lower bound is $\frac{1}{9}$: to minimize the proportion of 0's, we add 2 elements with label 1 and flip the element that satisfies g from label 0 to label 1.

Table 2: Certification rates of MNIST 1/7 binary for various bias models, using decision trees of depth 2. The composite bias models show cumulative bias, e.g., 0.2% MISS + FAKE bias equates to 0.1% bias of each MISS and FAKE. Note that the scale for perturbation set size is slightly different (larger) than that in Table 1.

Bias type	Dataset	Bias amount as a percentage of training set size					
		0.05	0.1	0.2	0.4	0.7	1.0
MISS	MNIST-1-7 binary	100.0	100.0	93.0	88.0	85.0	68.0
FLIP	MNIST-1-7 binary	100.0	95.0	88.0	70.0	63.0	38.0
MISS + FAKE	MNIST-1-7 binary	100.0	100.0	93.0	88.0	85.0	68.0
MISS + FLIP	MNIST-1-7 binary	100.0	95.0	91.0	87.0	67.0	62.0
Perturbation set size:		< 10^{10}	< 10^{100}	< 10^{1000}	< 10^{10000}	> 10^{10000}	infinite

Table 3: Time and memory requirements for certifying a single test sample under different FLIP bias models on Adult Income with depth=2.

	Poisoning Amount (%)					
	0.1	0.2	0.3	0.4	0.5	0.6
Time (s.)	0.60	73.9	210	810	1800	5200
Memory (GB)	0.01	0.8	3.6	9.7	21	60

E Additional experimental data

E.1 MNIST 1/7 Binary

We used MNIST 1/7 (the limitation of MNIST to just 1’s and 7’s, with training $n=13,007$, as has been used in works including [16, 36]). We round each pixel to 0 or 1 (i.e., used a black-and-white image rather than a grayscale one). The accuracy of MNIST 1/7 binary (97.4% at depth 2) is comparable to that of MNIST 1/7, but the time and memory requirements on Antidote-P are much less.

Table 2 shows effectiveness data for MNIST 1/7 binary. We see that we are able to achieve high robustness certification rates, despite incredibly large perturbation set sizes. Notable, for $\text{MISS}_{0.1\%}$, we achieve 100% robustness even with a perturbation set size of over 10^{3058} , and for $\text{MISS}_{1\%}$, we achieve 68% robustness with a perturbation set size larger than 10^{30460} .

E.2 Performance

We performed experiments on an HTCCondor system, allowing us to perform many experiments in parallel. Each experiment ran robustness tests on a given bias model and dataset for between one and 1000 test samples (depending on the bias model and dataset). We used a single CPU for each experiment, and requested between 1 and 96GB of memory, depending on the bias model and dataset.

Bias model. Time and memory requirements increase exponentially as the amount of bias increases, as shown in Table 3 for the Adult Income dataset under the FLIP bias model. Other datasets typically required less than 100 ms per test sample. Additionally, bias models that yield lower certifiable robustness for a given bias threshold have correspondingly larger time and memory requirements (e.g., 810s and 9.7GB of memory to yield 34.8% robustness for $\text{FLIP}_{0.4\%}$ as compared with 77s and 1.3GB of memory to yield 60.3% robustness for $\text{MISS}_{0.4\%}$).

Datasets. The size and complexity of the feature space is most influential in determining time and memory requirements. Experiments on the Adult Income dataset were more resource-intensive than those on Drug Consumption or COMPAS, a fact that can be explained by Adult Income having more unique feature values than the other datasets (22,100 for Adult Income vs. 219 for Drug Consumption and 53 for COMPAS). For each unique value of any feature, the algorithm checks an additional predicate, which explains the additional time and memory needs.

Table 4: Robustness certification rates of COMPAS and Drug-Consumption datasets under FLIP for different decision-tree depths.

Poisoning amount (%)	COMPAS			Drug Consumption		
	depth 1	2	3	depth 1	2	3
0.10	71.5	51.5	34.0	94.5	83.8	8.5
0.20	47.8	27.7	23.9	94.5	55.9	4.5
0.50	9.3	2.5	0.7	85.1	27.5	0.5
1.00	3.0	0.7	0	7.1	0.8	0

Complexity of decision-tree algorithm. Increasing the depth of the decision tree not only requires additional time to essentially re-run the algorithm at each internal node, but also leads to lower certifiable-robustness rates, as shown in Table 4. This is because we must assume worst-case bias in each node. Intuitively, a depth 2 tree with 0.1% bias may initially split the data into two children nodes, each with 50% of the data. Our abstraction captures both the case where all bias occurs in the left child, and the case when all bias occurs in the right child. Therefore, we end up with an effective bias rate of 0.2% in either child, yielding lower robustness.

E.3 Additional experimental data

General data. Fig. 3 shows the certifiable robustness rates for each dataset and each main bias model (MISS, FLIP, MISS + FAKE, and MISS + FLIP).

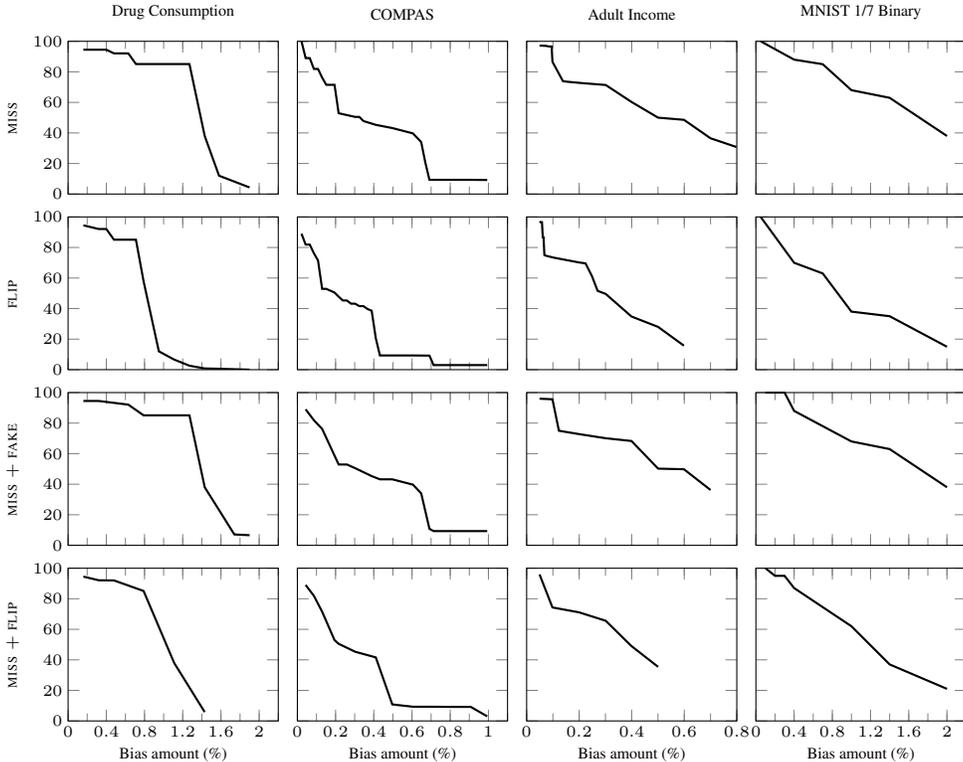


Figure 3: Certifiable robustness (shown on y-axis as a percentage of test data) for various datasets under different bias models.

Demographic data.

Fig. 4 shows robustness levels stratified by demographic groups on various bias models. We see that COMPAS under MISS (Fig. 4b) displays similar robustness gaps to FLIP; namely, White people, and particularly White women, are robust at a higher rate than Black people. Adult Income under

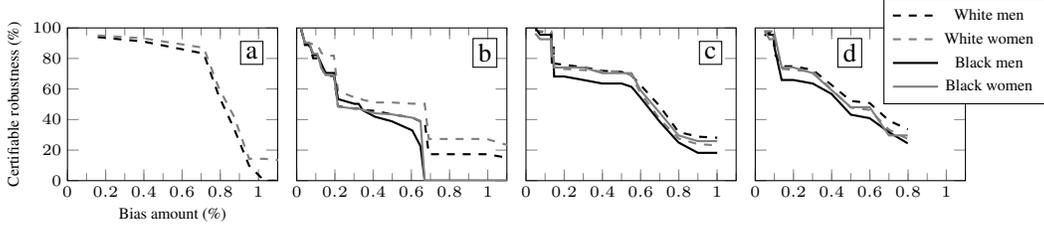


Figure 4: Left to right: Certifiable robustness by demographic group on (a) Drug Consumption under FLIP (NOTE: on this graph, the dark dotted line is all men, and the light dotted line is all women); (b) COMPAS under MISS; (c) Adult Income under FLIP^g where $g \triangleq (\text{gender} = \text{Female} \wedge \text{label} = \text{negative})$; (d) Adult Income under MISS

Table 5: Number of elements with counterexamples to robustness after 10,000 iterations of random testing from a subset of 100 test samples from COMPAS. All bias models are FLIP, and bias level refers to number of affected elements as a percentage of training dataset size.

Bias level	# of elements with counterexample
0.5	4
1.0	15
2.0	27
3.0	32

MISS (Fig. 4c) and under FLIP^g where $g \triangleq (\text{gender} = \text{female} \wedge \text{label} = \text{negative})$ (Fig. 4d) behaves similarly to Adult Income under FLIP (Fig. 2). That is, all demographic groups have roughly comparable robustness rates. Drug consumption under FLIP (Fig. 4a) yields comparable robustness rates between men and women (we do not graph robustness rates by racial group because the dataset is over 91% White).

E.4 Additional details on random testing

On a random subset of 100 test elements from the COMPAS dataset, we tested 10,000 dataset perturbations under FLIP_{0.5%}, FLIP_{1%}, FLIP_{2%}, and FLIP_{3%}. The number of elements for which we found a counter-example to robustness (i.e., a dataset perturbation that resulted in a different classification) is shown in Table 5. We see that we are able to find counterexamples to robustness for a non-trivial portion of test samples. However, the gap between certified-robustness and proved-non-robust rates is still wide (the gap ranges from 86.7% for FLIP_{0.5%} to 68.0% for FLIP_{3%}). As a result, there are many test samples that we cannot prove robustness for, but cannot find counterexamples for either. Future work to use a more precise abstract domain, or to better identify counterexamples to robustness could help to narrow this gap.

Breaking down the results for FLIP_{3%} further, we found counterexamples to robustness for 50% of Black women, 37% of Black men, 29% of White women, and 27% of White men. Similarly, we found more counter-examples to robustness for Black people using other bias models. The empirical result of having more counterexamples for test instances representing Black people combined with the fact that we are able to certify a smaller percentage of test instances representing Black people (§ 5.2) suggests that the robustness differences are inherent to the data, rather than a property of the abstraction.