CS 839 Probability and Learning in High Dimension

Lecture 15 - 03/28/2022

Lecture 15: Statistical Learning

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In this lecture, We will derive the estimation error in the learning task and introduce a Rademacher complexity based technique to upper bound it. We will also see one example as the application of this bound.¹

1 Problem Set Up

Consider the following learning task. Let $f^*: \mathfrak{X} \to [0,1]$ being the unknown true regression function. We observe *n* data points $(x_1, f^*(x_1)), \ldots, (x_n, f^*(x_n))$, where the feature vectors x_i 's are sampled i.i.d. from some unknown distribution μ . The goal is to estimate f^* given the data.

For a given function $f: \mathfrak{X} \to [0, 1]$, define the population risk (a.k.a. test error):

$$\mathcal{L}(f) = \mathbb{E}_{x \sim \mu} \Big(f(x) - f^*(x) \Big)^2$$

and the empirical risk (a.k.a. training error):

$$\mathcal{L}_n(f) = \frac{1}{n} \sum_{i=1}^n \left(f(x_i) - f^*(x_i) \right)^2.$$

Ideally, we want to find the population risk minimizer $f_o = \arg \min_{f \in \mathcal{F}} \mathcal{L}(f)$, which is however not computable since f^* and μ . As a surrogate, we consider the empirical risk minimizer (ERM)

$$\hat{f} = \operatorname*{arg\,min}_{f \in \mathcal{F}} \mathcal{L}_n(f).$$

Our goal is to control the population risk of the ERM \hat{f} .

Risk decomposition: We can decompose the population risk of \hat{f} as follows

$$\underbrace{\mathcal{L}(\hat{f})}_{\text{test error}} = \underbrace{\left(\mathcal{L}(\hat{f}) - \mathcal{L}_n(\hat{f})\right)}_{\text{generalization gap}} + \underbrace{\mathcal{L}_n(\hat{f})}_{\text{training error}} \\ \leq \left(\mathcal{L}(\hat{f}) - \mathcal{L}_n(\hat{f})\right) + \mathcal{L}_n(f_o) \\ = \underbrace{\left(\mathcal{L}(\hat{f}) - \mathcal{L}_n(\hat{f})\right) + \left(\mathcal{L}_n(f_o) - \mathcal{L}(f_o)\right)}_{\text{statistical error}} + \underbrace{\mathcal{L}(f_o)}_{\text{approximation error}} + \underbrace{\mathcal{L}(f_o)}_{\text{a$$

where the inequality above holds since \hat{f} minimizes \mathcal{L}_n . We have mentioned that $\mathcal{L}(\hat{f}) - \mathcal{L}_n(\hat{f})$ is called the **generalization gap/error**, which is the gap between the test and training error of \hat{f} . The first two terms

• Section 3.3 in [Duchi, 2021]

 $^{^{1}}Reading$:

[•] Section 4.1 and 4.2 in [Wainwright, 2019]

[•] Section 8.4 in [Vershynin, 2018]

in the last line above are the differences between an empirical quantity (defined by \mathcal{L}_n) and its population counterpart (defined by \mathcal{L}). These two terms represent the **statistical/estimation error** due to having a finite number of data points. The last term $\mathcal{L}(f_o)$ measures how well the function class \mathcal{F} can approximate the true function f^* under the real data distribution μ ; this term represents the **approximation error**.

We can upper bound the statistical error error by the supremum of the difference:

$$\left(\mathcal{L}(\hat{f}) - \mathcal{L}_n(\hat{f})\right) + \left(\mathcal{L}_n(f_o) - \mathcal{L}(f_o)\right) \le 2\sup_{f \in \mathcal{F}} |\mathcal{L}_n(f) - \mathcal{L}(f)|,$$

which leads to the bound

$$\underbrace{\mathcal{L}(\hat{f}) - \mathcal{L}(f_o)}_{\text{excess risk}} \leq 2 \sup_{f \in \mathcal{F}} |\mathcal{L}_n(f) - \mathcal{L}(f)|.$$

In what follows, we establish upper bound on the above supremum using Rademacher complexity.

2 Upper Bound Using Rademacher Complexity

Assume \mathcal{F} is [0, 1]-bounded, i.e. $\forall f \in \mathcal{F}, \forall x \in \mathcal{X}: f(x) \in [0, 1]$. Also assume $f^* \in \mathcal{F}$. Recall that we consider the mean square loss. The supremum above can be written as

$$\sup_{f \in \mathcal{F}} |\mathcal{L}_n(f) - \mathcal{L}(f)| = \sup_{f \in \mathcal{F}} \left| \frac{1}{n} \sum_{i=1}^n \underbrace{\left(f(x_i) - f^*(x_i) \right)^2}_{g(x_i)} - \mathbb{E} \left(f(x) - f^*(x) \right)^2 \right|$$
$$= \sup_{g \in \mathcal{G}} \left| \frac{1}{n} \sum_{i=1}^n g(x_i) - \mathbb{E} [g(x_i)] \right|,$$

where $\mathcal{G} \stackrel{\Delta}{=} \left\{ x \mapsto \left(f(x) - f^*(x) \right)^2 \right\}$. Note that the quantity of the form $\sup -\mathbb{E} \sup$ can be bounded by concentration inequalities. Below we focus on bounding the expectation $\mathbb{E} \sup$.

2.1 Symmetrization

Let $(\epsilon_1, \ldots, \epsilon_n)$ be i.i.d Rademacher random variables. By the symmetrization argument in Theorem 8 of lecture 14, we have

$$\mathbb{E}\sup_{g\in\mathcal{G}}\Big|\frac{1}{n}\sum_{i=1}^{n}(g(x_i)-\mathbb{E}g(x_i))\Big|\leq 2\mathbb{E}_x\mathbb{E}_{\epsilon}\sup_{g\in\mathcal{G}}\Big|\frac{1}{n}\sum_{i=1}^{n}\epsilon_ig(x_i)\Big|.$$

We introduce some definitions. Define empirical Rademacher complexity of \mathcal{G} as

$$\mathcal{R}_n(\mathcal{G}|x) \stackrel{\Delta}{=} \mathbb{E}_{\epsilon} \sup_{g \in \mathcal{G}} \Big| \frac{1}{n} \sum_{i=1}^n \epsilon_i g(x_i) \Big|,$$

and the **Rademacher complexity** of \mathcal{G} as

$$\mathcal{R}_n(\mathcal{G}) \stackrel{\Delta}{=} \mathbb{E}_x[\mathcal{R}_n(\mathcal{G}|x)].$$

Using these notations, we have established the following:

Theorem 1 (Symmetrization Bound).

$$\mathbb{E}\sup_{g\in\mathcal{G}}\left|\frac{1}{n}\sum_{i=1}^{n}\left(g(x_i)-\mathbb{E}g(x_i)\right)\right| \leq 2\mathcal{R}_n(\mathcal{G}) = 2\mathbb{E}_x[\mathcal{R}_n(\mathcal{G}|x)].$$

2.2 Contraction

Recall that $g(x_i) := (f(x_i) - f^*(x_i))^2$, and that f and f^* are [0, 1]-bounded. Also note that the square function $\phi : \theta \mapsto \theta^2$ is 2-Lipschitz over [-1.1]. By the Rademacher contraction principle in Theorem 7 from Lecture 14, we have

$$\begin{aligned} \mathcal{R}_{n}(\mathcal{G}|x) &= \mathbb{E}_{\epsilon} \sup_{f \in \mathcal{F}} \left| \frac{1}{n} \sum_{i=1}^{n} \epsilon_{i} \left(f(x_{i}) - f^{*}(x_{i}) \right)^{2} \right| \\ &= \mathbb{E}_{\epsilon} \sup_{f \in \mathcal{F}} \left| \frac{1}{n} \sum_{i=1}^{n} \epsilon_{i} \phi \left(f(x_{i}) - f^{*}(x_{i}) \right) \right| \\ &\leq 2 \mathbb{E}_{\epsilon} \sup_{f \in \mathcal{F}} \left| \frac{1}{n} \sum_{i=1}^{n} \epsilon_{i} \left(f(x_{i}) - f^{*}(x_{i}) \right) \right| \qquad \text{contraction principle} \\ &\leq 4 \mathbb{E}_{\epsilon} \sup_{f \in \mathcal{F}} \left| \frac{1}{n} \sum_{i=1}^{n} \epsilon_{i} f(x_{i}) \right| \qquad f^{*} \in \mathcal{F} \\ &= 4 \mathcal{R}_{n}(\mathcal{F}|x). \end{aligned}$$

It follows that $\mathcal{R}_n(\mathcal{G}) \leq 4\mathcal{R}_n(\mathcal{F})$.

2.3 Putting Together

Recapping the arguments above, we have

$$\begin{split} \mathcal{L}_{n}(\hat{f}) - \mathcal{L}(f_{o}) &\lesssim \mathbb{E} \sup_{f \in \mathcal{F}} \left| \mathcal{L}_{n}(f) - \mathcal{L}(f) \right| & \text{risk decomposition} \\ &= \sup_{f \in \mathcal{F}} \left| \frac{1}{n} \sum_{i=1}^{n} \left(f(x_{i}) - f^{*}(x_{i}) \right)^{2} - \mathbb{E} \left(f(x) - f^{*}(x) \right)^{2} \right| \\ &= \sup_{g \in \mathcal{G}} \left| \frac{1}{n} \sum_{i=1}^{n} \left[g(x_{i}) - \mathbb{E} g(x_{i}) \right] \right| \\ &\lesssim \mathcal{R}_{n}(\mathcal{G}) & \text{symmetrization} \\ &\lesssim \mathcal{R}_{n}(\mathcal{F}) & \text{contraction principle} \\ &= \mathbb{E}_{x} \mathbb{E}_{\epsilon} \sup_{f \in \mathcal{F}} \left| \frac{1}{n} \sum_{i=1}^{n} \epsilon_{i} f(x_{i}) \right|. \end{split}$$

We have upper bounded the supremum of one empirical process by that of another, and both processes are indexed by $f \in \mathcal{F}$. We are doing this because $\mathcal{R}_n(\mathcal{F}) = \mathbb{E}_x[\mathcal{R}_n(\mathcal{F}|x)]$ is often easier to control. In particular, we can bound $\mathcal{R}_n(\mathcal{F}|x)$ conditioned on the data x. For fixed x, the quantity $\mathcal{R}_n(\mathcal{F}|x) = \mathbb{E}_{\epsilon} \sup_{f \in \mathcal{F}} |\frac{1}{n} \langle \epsilon, f(X) \rangle |$ is the supremum of a (canonical) Rademacher process. To control this supremum, we may use the following techniques:

- Union bound
- Dudley integral bound (e.g., when \mathcal{F} is the set of Lipschitz functions; see Lecture 14 for details.)
- VC-dimension (usually used for binary functions; not covered in this course)
- Talagrand comparison: $\mathbb{E}_{\epsilon} \sup_{f \in \mathcal{F}} |\langle \epsilon, f(X) \rangle| \lesssim \mathbb{E}_{g \sim \mathcal{N}(0,I)} \sup_{f \in \mathcal{F}} |\langle g, f(X) \rangle|$. Then we can use any techniques for Gaussian process to bound the RHS.

In the following section, we give an example for bounding $\mathcal{R}_n(\mathcal{F})$ and $\mathcal{R}_n(\mathcal{F}|x)$ using union bound.

3 Example: Glivenk-Cantelli Uniform Law of Large Number (ULLN)

Let x_1, \ldots, x_n be i.i.d random variables with distribution μ and Cumulative Distribution Function (CDF) $F(\theta) = \Pr[x_1 \le \theta] = \mathbb{E}[\mathbb{1}\{x_1 \le \theta\}].$

We can estimate the true CDF F using empirical CDF, defined as

$$\hat{F}(\theta) := \frac{1}{n} \sum_{i=1}^{n} \mathbb{1}\{x_1 \le \theta\} = \frac{1}{n} \sum_{i=1}^{n} g_{\theta}(x_i),$$

where we have Introduced the short hand $g_{\theta}(x) := \mathbb{1}\{x \leq \theta\}$. Denote the set of such indicator functions by $\mathcal{G} \stackrel{\Delta}{=} \{g_{\theta} : \theta \in \mathbb{R}\}$. Note that the functions g_{θ} are **not** Liptschitz.

We have

$$\mathbb{E}\sup_{\theta \in \mathbb{R}} \left| \hat{F}(\theta) - F(\theta) \right| = \mathbb{E}\sup_{g \in \mathcal{G}} \left| \frac{1}{n} \sum_{i=1}^{n} g(x_i) - \mathbb{E}g(x_1) \right|$$

$$\leq \mathcal{R}_n(\mathcal{G}) \qquad (\text{Theorem 1})$$

$$= \mathbb{E}_x \mathcal{R}_n(\mathcal{G}|x)$$

$$= \frac{1}{n} \mathbb{E}_x \mathbb{E}_\epsilon \sup_{\theta \in \mathbb{R}} \left| \sum_{i=1}^{n} \epsilon_i g_\theta(x_i) \right|.$$

Let us condition on fixed x_1, \ldots, x_n ; assume w.l.o.g. that $x_1 \le x_2 \le \cdots \le x_n$. Note that the *n*-dimensional vector $(g_{\theta}(x_1), \ldots, g_{\theta}(x_n)) \in \{0, 1\}^n$ can take on at most n + 1 values:

$$(0, 0, \dots, 0) (1, 0, \dots, 0) (1, 1, \dots, 0) \vdots (1, 1, \dots, 1)$$

Therefore, $\sup_{\theta \in \mathbb{R}} \left| \sum_{i} \epsilon_{i} g_{\theta}(x_{i}) \right|$ is the supremum of at most (n + 1) random variables. Moreover, for each $\theta \in \mathbb{R}$, the random variable $\epsilon_{i} g_{\theta}(x_{i})$ is zero-mean and lies in the interval $\in [-1, 1]$. It follows that the sum $\sum_{i} \epsilon_{i} g_{\theta}(x_{i})$ is a zero-mean, O(n)-sub-Gaussian random variable by Hoeffding inequality.

Using bound on the maximum of a finite number of sub-Gaussian random variables (Lecture 13, Lemma 3), we obtain

$$\mathbb{E}\sup_{\theta\in\mathbb{R}}\Big|\sum_{i=1}^{n}\epsilon_{i}g_{\theta}(x_{i})\Big|\lesssim\sqrt{n\log(n+1)}.$$

Combining pieces, we obtain the following upper bound on the expectation:

$$\mathbb{E}\sup_{\theta\in\mathbb{R}}\left|\hat{F}(\theta) - F(\theta)\right| \le \sqrt{\frac{\log n}{n}}.$$

We can further use the Bounded Difference Inequality (Lecture 14, Theorem 4) to prove concentration around the expectation. Together, we obtain the following theorem:

Theorem 2. With probability at least $1 - e^{-n\delta^2}$,

$$\sup_{\theta \in \mathbb{R}} \left| \hat{F}(\theta) - F(\theta) \right| \le \sqrt{\frac{\log n}{n}} + \delta.$$

Hence $\sup_{\theta \in \mathbb{R}} \left| \hat{F}(\theta) - F(\theta) \right| \xrightarrow{a.s.} 0.$

Remark: We can remove the $\sqrt{\log n}$ factor using Dudley's integral bound and VC-dimension.

References

- [Duchi, 2021] Duchi, J. (2021). Lecture notes in statistics 311/electrical engineering 377: Information theory and statistics. http://web.stanford.edu/class/stats311/lecture-notes.pdf.
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