| CS 839 Probability and Learning in High Dimension | Lecture 22-04/20/2022 |
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|  | Lecture 22: Linear MDPS II |
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In this lecture ${ }^{1}$ we recap the structure of Linear MDPs and the episodic setting. We then introduce the Least-Squares Value Iteration with Upper Confidence Bound (LSVI-UCB) algorithm and the regret bound for this algorithm. We also discuss the proof of the regret bound.

## 1 Recap: Linear structure

We assume that both the reward function and transition kernel have a linear structure, with respect to some known feature map. In the sequel $\|\cdot\|$ denotes the $\ell_{2}$ norm on $\mathbb{R}^{d}$.

Assumption 1 (Linearity and Boundedness). For each $h \in[H]$ and $\left(x, a, x^{\prime}\right) \in \mathcal{S} \times \mathcal{A} \times \mathcal{S}$, it holds that

$$
\begin{aligned}
\mathbb{P}_{h}\left(x^{\prime} \mid x, a\right) & =\left\langle\phi(x, a), \mu_{h}\left(x^{\prime}\right)\right\rangle \quad \text { and } \\
r_{h}(x, a) & =\left\langle\phi(x, a), \theta_{h}\right\rangle
\end{aligned}
$$

where

- $\phi_{h}: \mathcal{S} \times \mathcal{A} \rightarrow \mathbb{R}^{d}$ is a known feature map,
- $\mu_{h}=\left(\mu_{h}^{(i)}\right)_{i \in[d]}$ is a vector of d unknown (signed) measures on $\mathcal{S}$, and
- $\theta_{h}=\left(\theta_{h}^{1}, \ldots, \theta_{h}^{d}\right) \in \mathbb{R}^{d}$ is a vector of d unknown weights.

We assume that $\max _{(x, a) \in \mathcal{S} \times \mathcal{A}}\left\|\phi_{h}(x, a)\right\| \leq 1,\left\|\mu_{h}(\mathcal{S})\right\| \leq \sqrt{d},\left\|\theta_{h}\right\| \leq \sqrt{d}$ for all $h \in[H] \stackrel{2}{2}^{2}$

The above assumption implies that the Q-function is linear for any policy (including the optimal policy).
Lemma 1 (Linearity of Q). For any policy $\pi$ and $h \in[H]$, there exists a weight vector $w_{h}^{\pi} \in \mathbb{R}^{d}$ such that

$$
Q_{h}^{\pi}(x, a)=\left\langle\phi(x, a), w_{h}^{\pi}\right\rangle, \quad \forall(x, a) \in \mathcal{S} \times \mathcal{A}
$$

In particular, the optimal $Q$-function satisfies $Q_{h}^{*}(x, a)=\left\langle\phi(x, a), w_{h}^{*}\right\rangle, \forall x$, a for some $w_{h}^{*} \in \mathbb{R}^{d}$.

## 2 Episodic setting and regrets

The agent interacts with the MDP in $K$ episodes. At the beginning of episode $k$, the agent picks a policy $\pi^{k}=\left(\pi_{1}^{k}, \ldots, \pi_{H}^{k}\right)$ and receives an (arbitrary) initial state $x_{1}^{k}$. The agents then executes the policy for $H$ steps, resulting in the trajectory

$$
x_{1}^{k}, a_{1}^{k}, r_{1}^{k}, \ldots, x_{H}^{k}, a_{H}^{k}, r_{H}^{k}
$$

where $a_{h}^{k} \sim \pi_{h}^{k}\left(x_{h}^{k}\right), r_{h}^{k}=r\left(x_{h}^{k}, a_{h}^{k}\right)$ and $x_{h+1}^{k} \sim \mathbb{P}_{h}\left(\cdot \mid x_{h}^{k}, a_{h}^{k}\right)$. The system then resets and episode $(k+1)$ begins.

The regret over $K$ episodes is defined as

$$
\operatorname{Regret}(K):=\sum_{k=1}^{K}\left[V_{1}^{*}\left(x_{1}^{k}\right)-V_{1}^{\pi^{k}}\left(x_{1}^{k}\right)\right]
$$

which is the difference between the total value of the agent's policy $\pi^{1}, \ldots, \pi^{K}$ and that of the optimal policy. We want to find an algorithm that achieves a low regret.


Figure 1: Illustration of value functions across episodes. The dashed line represents $V_{1}^{*}\left(x_{1}\right)$, indicating the highest possible return one can get in each episode. The black solid line represents the value function as the episode $k$ progresses. The yellow line represents the value function by another possible algorithm, which has higher regret.

Suppose all $K$ episodes start at the same initial state $x_{1}$. The regret correspond to the shaded area in Figure 1 which we aim to minimize.
Remark 2. In Figure 1, both algorithms (black curve and yellow curve) eventually converge to the optimal value when $k$ is sufficiently large. However, the algorithm corresponding to the black curve has smaller regret during the learning process.
Remark 3. Suppose the (total) regret grows sublinearly, i.e., $\operatorname{Regret}(K)=o(K)$. In this case, the average regret $\operatorname{AvgRegret}(K):=\frac{1}{K} \sum_{k=1}^{K}\left[V_{1}^{*}\left(x_{1}^{k}\right)-V_{1}^{\pi^{k}}\left(x_{1}^{k}\right)\right]=o(1)$ ultimately goes to 0 . For AvgRegret( $K$ ), it's possible for a few episodes to have very bad value functions, but the effect of the bad episodes won't matter as it will eventually average out.

## 3 Algorithm and guarantees

The algorithm, Least-Squares Value Iteration with Upper Confidence Bound (LSVI-UCB), is given in Algorithm 1

```
Algorithm 1 LSVI-UCB
for episode \(k=1,2, \ldots, K\) do
    1. (Value estimation) for step \(h=H, H-1, \ldots, 1\) do
    (a) (Gram matrix) \(\Lambda_{h}^{k} \leftarrow \sum_{\tau \in[k-1]} \phi\left(x_{h}^{\tau}, a_{h}^{\tau}\right) \phi\left(x_{h}^{\tau}, a_{h}^{\tau}\right)^{\top}+I\)
    (b) (Least squares)
\[
\begin{aligned}
w_{h}^{k} & \leftarrow \arg \min _{w \in \mathbb{R}^{d}} \sum_{\tau \in[k-1]}\left[r_{h}\left(x_{h}^{\tau}, a_{h}^{\tau}\right)+V_{h+1}^{k}\left(x_{h+1}^{\tau}\right)-\left\langle w, \phi\left(x_{h}^{\tau}, a_{h}^{\tau}\right)\right\rangle\right]^{2}+\|w\|^{2} \\
& \left.=\left(\Lambda_{h}^{k}\right)^{-1} \sum_{\tau \in[k-1]} \phi\left(x_{h}^{\tau}, a_{h}^{\tau}\right) \cdot\left[r_{h}\left(x_{h}^{\tau}, a_{h}^{\tau}\right)+V_{h+1}^{k}\left(x_{h+1}^{\tau}\right)\right)\right] .
\end{aligned}
\]
(c) (Q estimate with UCB) \(Q_{h}^{k}(\cdot, \cdot)=\left\langle w_{h}^{k}, \phi(\cdot, \cdot)\right\rangle+\beta \sqrt{\phi(\cdot, \cdot)^{\top}\left(\Lambda_{h}^{k}\right)^{-1} \phi(\cdot, \cdot)}\)
(d) (From Q to value function) \(V_{h}^{k}(\cdot)=\max _{a} Q_{h}^{k}(\cdot, a)\).
2. Receive initial state \(x_{1}^{k}\)
3. (Policy execution) for step \(h=1,2, \ldots, H\) do
Take action \(a_{h}^{k} \leftarrow \arg \max _{a} Q_{h}^{k}\left(x_{h}^{k}, a\right)\); observe reward \(r_{h}^{k}=r_{h}\left(x_{h}^{k}, a_{h}^{k}\right)\) and next state \(x_{h+1}^{k}\).
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Below we discuss and provide intuition for the steps in Algorithm 1.

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### 3.1 Least squares estimation (Step 1(a)-(b))

Recall that $Q_{h}^{*}(x, a)=\left\langle\phi(x, a), w_{h}^{*}\right\rangle$. Our first goal is to estimate the unknown $w_{h}^{*}$ associated with the optimal $Q_{h}^{*}$. Under the linear assumption, Lemma 1 guarantees that

$$
\left\langle\phi(x, a), w_{h}^{*}\right\rangle=r_{h}(x, a)+\left(\mathbb{P}_{h} V_{h+1}^{*}\right)(x, a)
$$

The next-step value function $V_{h+1}^{*}$ on the RHS is unknown, so we may replace it by $V_{h+1}^{k}$, which is our current estimate of the next-step value function. The conditional distribution $\mathbb{P}_{h}\left(x^{\prime} \mid x, a\right)$ is also unknown, but we can estimate it using empirical observation of the next state $x^{\prime}$ (this is called a one-sample estimate). Combining, we see that $w_{h}^{*}$ satisfies the following approximate relationship:

$$
\left\langle\phi(x, a), w_{h}^{*}\right\rangle \approx r_{h}(x, a)+V_{h+1}^{k}\left(x^{\prime}\right)
$$

Therefore, we can estimate $w_{h}^{*}$ by finding a vector $w_{h}^{k}$ that minimizes the difference between the LHS and RHS of the above, over the data from episode $1, \cdots, k-1$. That is,

$$
w_{h}^{k} \leftarrow \arg \min _{w \in \mathbb{R}^{d}} \sum_{\tau \in[k-1]}\left[r_{h}\left(x_{h}^{\tau}, a_{h}^{\tau}\right)+V_{h+1}^{k}\left(x_{h+1}^{\tau}\right)-\left\langle w, \phi\left(x_{h}^{\tau}, a_{h}^{\tau}\right)\right\rangle\right]^{2}+\|w\|^{2}
$$

The regularization term $\|w\|^{2}$ ensures uniqueness of the solution $w_{h}^{k}$. The optimal solution of $w_{h}^{k}$ can be written in a closed-form as shown in step 1(b) in Algorithm 1 .

### 3.2 Value estimation and bonus term (Step 2(c))

Given the estimate $w_{h}^{k}$, we can calculate $Q_{h}^{k}(\cdot, \cdot)$ and $V_{h}^{k}(\cdot)$, which are estimates of the true value and Q functions $Q_{h}^{*}$ and $V_{h}^{*}$, in Steps 1(c) and 1(d) respectively.

In step $1(\mathrm{c})$ above, we add a "bonus" term $\beta \sqrt{\phi(\cdot, \cdot)^{\top}\left(\Lambda_{h}^{k}\right)^{-1} \phi(\cdot, \cdot)}$ that accounts for the uncertainty in the least square estimate $w_{h}^{k}$, thereby ensuring that with high probability $Q_{h}^{k}(x, a)$ is an upper confidence bound (UCB) of the true Q function $Q_{h}^{*}(x, a)$ for all $(x, a)$. This upper bound is larger for state-action pairs $(x, a)$ that are infrequently visited in the past, so it encourages exploration of these pairs in Step 3. This idea, as well as the particular form of the bonus, are a generalization of the UCB algorithm for multi-arm bandit.

Recall that tabular MDP is a special case of the linear MDP. It is instructive to look at the particular form of the "bonus" term in the tabular setting. In this setting, $d=|\mathcal{S} \| \mathcal{A}|$ and the feature map $\phi(s, a)=\mathbf{e}_{x a}$, which takes 1 at the $x a$ entry and 0 at the other entries. Consequently, the gram matrix

$$
\Lambda_{h}^{k}=I+\sum_{\tau}^{k-1} \mathbf{e}_{x_{h}^{\tau} a_{h}^{\tau}}^{\mathbf{e}_{x_{h}^{\tau}}^{\top} a_{h}^{\tau}} \in \mathbb{R}^{|\mathcal{S} \| \mathcal{A}| \times|\mathcal{S}||\mathcal{A}|}
$$

is a diagonal matrix, where each diagonal entry is

$$
\begin{aligned}
\Lambda_{h}^{k}(x a, x a) & =1+\sum_{\tau=1}^{k-1} \mathbb{1}\left\{\left(x_{h}^{\tau}, a_{h}^{\tau}\right)=(x, a)\right\} \\
& =1+\# \text { of visits to }(x, a) \text { pair in step } h \text { of episode } 1, \cdots, k-1 \\
& =: 1+N^{k-1}(x, a)
\end{aligned}
$$

Thus, the bonus term in the tabular case takes the form

$$
\sqrt{\phi(x, a)^{\top}\left(\Lambda_{h}^{k}\right)^{-1} \phi(x, a)}=\sqrt{\frac{1}{1+N^{k-1}(x, a)}} .
$$

This implies that the fewer visits to the state-action pair $(x, a)$, the larger the "bonus" term for that pair, which aligns with our initial intention for the "bonus" term.

In the general linear MDP case, we may have two different state-action pairs with similar features $\phi(x, a) \approx \phi\left(x^{\prime}, a^{\prime}\right)$. In this case, they will have similar "bonus" terms.

### 3.3 Policy Execution (Step 2 \& 3)

Given $Q_{h}^{k}(\cdot, \cdot), h \in[H]$, we can construct the corresponding (deterministic) greedy policy $\pi^{k}=\left(\pi_{h}^{k}\right)_{h \in[H]}$, where $\pi_{h}^{k}(x)=\arg \max _{a} Q_{h}^{k}(x, a)$. Starting from the initial state $x_{1}^{k}$ in Step 2, we can then execute the policy $\pi^{k}$ to play out episode $k$, as shown in Step 3.

### 3.4 Computational complexity

Regarding the computational complexity of Algorithm 1 it cannot be implemented if we naively follow the steps since when $|\mathcal{S}|=\infty$, we have infinite state-action pairs and it's impossible to calculate $Q_{h}^{k}(\cdot, \cdot)$ and $V_{h}^{k}(\cdot)$ for each pair in each $k$ and $h$.

A closer inspection of Algorithm 1 shows that we only need to calculate $Q_{h}^{k}(x, \cdot), V_{h}^{k}(x)$ for the states $x$ that we actually encounter during the episodes.

## 4 Regret bound

We establish the following regret bound for LSVI-UCB. Here $T:=K H$ is the total number of steps over all episodes.
Theorem 4. Set $\beta=c d H \sqrt{\iota}$ with $\iota:=\log (2 d T / p)$. With probability at least $1-p$, we have

$$
\operatorname{Regret}(K):=\sum_{k=1}^{K}\left[V_{1}^{*}\left(x_{1}^{k}\right)-V_{1}^{\pi^{k}}\left(x_{1}^{k}\right)\right] \lesssim \sqrt{d^{3} H^{3} T \iota^{2}}=\sqrt{d^{3} H^{4} K \iota^{2}} .
$$

Remark 5. We can compare the above regret bound with some known minimax bounds. For tabular MDP, the minimax regret bound is regret $\gtrsim \sqrt{d H^{3} K}$. Therefore, the dependence on $H$ in Theorem 4 is off by a factor of $\sqrt{H}$.

For linear bandit problem where $H=1$, the minimax regret bound is regret $\gtrsim \sqrt{d^{2} K}$. Therefore, the dependence on $d$ in Theorem 4 is off by a factor of $\sqrt{d}$. We will point out later where this additional $d$ factor comes from in the proof.

### 4.0.1 Sample complexity bound

From the regret bound in Theorem 4, we can derive a sample complexity bound for finding an $\epsilon$-optimal policy. Algorithm 1 outputs $K$ policies $\pi^{1}, \pi^{2}, \cdots, \pi^{K}$. Among them we can randomly pick a policy: $\hat{\pi} \sim \operatorname{uniform}\left\{\pi^{1}, \cdots, \pi^{K}\right\}$. For a given $\epsilon>0$, we have

$$
\begin{array}{rlrl}
\mathbb{P}\left(V_{1}^{*}\left(x_{1}\right)-V_{1}^{\hat{\pi}}\left(x_{1}\right) \geq \epsilon\right) & \leq \frac{\mathbb{E}\left[V_{1}^{*}\left(x_{1}\right)-V_{1}^{\hat{\pi}}\left(x_{1}\right)\right]}{\epsilon} & \text { (by Markov Inequality) } \\
& =\frac{\frac{1}{K} \sum_{k=1}^{K}\left[V_{1}^{*}\left(x_{1}\right)-V_{1}^{\hat{\pi}}\left(x_{1}\right)\right]}{\epsilon} & \\
& \leq \frac{\frac{1}{K} \sqrt{d^{3} H^{4} K}}{\epsilon} & & \text { (ignore } \iota \text { in Theorem 4 } \\
& =\sqrt{\frac{d^{3} H^{4}}{K \epsilon}} &
\end{array}
$$

It follows that when $K \geq \frac{d^{3} H^{4}}{0.1^{2} \epsilon}$, we have $\mathbb{P}\left(V_{1}^{*}\left(x_{1}\right)-V_{1}^{\hat{\pi}}\left(x_{1}\right) \geq \epsilon\right) \leq 0.1$. This means that with probability $\geq 0.9, \hat{\pi}$ is an $\epsilon$-optimal policy.
Remark 6 . When specialized to the tabular setting where $d=|\mathcal{S} \| \mathcal{A}|$, the above sample complexity bound for LSVI-UCB becomes $K \gtrsim \frac{|\mathcal{S}|^{3}|\mathcal{A}|^{3} H^{4}}{\epsilon^{2}}$. One may compare this bound with the minimax sample complexity bound we get last week, which reads $K \gtrsim \frac{1}{(1-\gamma)^{3}} \cdot \frac{|\mathcal{S} \||\mathcal{A}|}{\epsilon^{2}} \approx H^{3}$. $\frac{|\mathcal{S} \| \mathcal{A}|}{\epsilon^{2}}$, where we treats the effective horizon $\frac{1}{1-\gamma}$ as the horizon. We see that the above sample complexity bound for LSVI-UCB is sub-optimal by a factor of $H \cdot|\mathcal{S}|^{2}|\mathcal{A}|^{2}$ not an optimal bound. The additional $H$ factor can be removied by changing the current "Hoeffding bonus" term to a "Bernstein Bonus" term. It is not clear yet whether the difference in $|\mathcal{S} \| \mathcal{A}|$ can be removed.

## 5 Proof of Theorem 4

The proof proceeds in 5 steps.

1. Concentration
2. Least-squares estimation error
3. UCB property
4. Regret decomposition
5. Final regret bound

Today we will cover Step 1.
Define the shorthand $\phi_{h}^{\tau}:=\phi\left(x_{h}^{\tau}, a_{h}^{\tau}\right)$.

### 5.1 Concentration

We present a concentration result, which is the crucial step of the proof. For a given positive definite matrix $A$, define the weighted norm $\|u\|_{A}:=\sqrt{u^{\top} A u}$.

Lemma 7 (Concentration of empirical measure). For each $p$, the following event $\mathfrak{E}$ holds with probability at least $1-p / 2$ :

$$
\left\|\sum_{\tau \in[k-1]} \phi_{h}^{\tau}\left[V_{h+1}^{k}\left(x_{h+1}^{\tau}\right)-\left(\mathbb{P}_{h} V_{h+1}^{k}\right)\left(x_{h}^{\tau}, a_{h}^{\tau}\right)\right]\right\|_{\left(\Lambda_{h}^{k}\right)^{-1}} \lesssim d H \sqrt{\log (d T / p)}, \quad \forall k, h
$$

Roughly speaking, this lemma says that the empirical sum $\sum_{\tau} \phi_{h}^{\tau} \cdot V\left(x_{h+1}^{\tau}\right)$ approximates the true expectation $\sum_{\tau} \phi_{h}^{\tau} \cdot\left(\mathbb{P}_{h} V\right)\left(x_{h}^{\tau}, a_{h}^{\tau}\right)$. The approximation error is measured in the norm $\|\cdot\|_{\left(\Lambda_{h}^{k}\right)^{-1}}$ weighted by the Gram matrix $\Lambda_{h}^{k}:=I+\sum_{\tau \in[k-1]} \phi_{h}^{\tau}\left(\phi_{h}^{\tau}\right)^{\top}$, where we recall that $\phi_{h}^{\tau}=\phi\left(x_{h}^{\tau}, a_{h}^{\tau}\right)$ are feature vectors of the previous visited state-action pairs $\left(x_{h}^{\tau}, a_{h}^{\tau}\right)$. Therefore, we have better approximation in the directions that are better covered by the previous data. Here we crucially exploit the linear structure: we care about coverage w.r.t. the feature space rather than w.r.t. individual state-action pairs.

Proof Fix $k$ and $h$. For each $\tau \in[k]$, define the sigma-algebra $\mathcal{F}_{\tau-1}=\sigma\left(x_{1: H}^{1}, \ldots x_{1: H}^{\tau-1}, x_{1}^{\tau}, \ldots, x_{h}^{\tau}\right)$, which includes everything up to step $h$ of episode $\tau$. Note that $\phi_{h}^{\tau}, x_{h}^{\tau} \in \mathcal{F}_{\tau-1}$ and $x_{h+1}^{\tau} \in \mathcal{F}_{\tau}$.

Consider $V_{h+1}^{k}$ as fixed first. Note that $V_{h+1}^{k}\left(x_{h+1}^{\tau}\right)-\mathbb{P}_{h} V_{h+1}^{k}\left(x_{h}^{\tau}\right) \mid \mathcal{F}_{\tau-1}$ is zero-mean and $H$-bounded. Applying the concentration inequality for self-normalized processes (Lemma 10), we obtain that with probability at least $1-\delta$ :

$$
\begin{equation*}
\left\|\sum_{\tau \in[k-1]} \phi_{h}^{\tau}\left[V_{h+1}^{k}\left(x_{h+1}^{\tau}\right)-\mathbb{P}_{h} V_{h+1}^{k}\left(x_{h}^{\tau}, a_{h}^{\tau}\right)\right]\right\|_{\left(\Lambda_{h}^{k}\right)^{-1}} \lesssim H \sqrt{\log \frac{(k+1)^{d / 2}}{\delta}} \tag{1}
\end{equation*}
$$

Note that the log factor on the RHS comes from the bound $\operatorname{det} \Lambda_{h}^{k} \leq\left(\left\|\Lambda_{h}^{k}\right\|_{\mathrm{op}}\right)^{d} \leq(k+1)^{d}$.
In reality, $V_{h+1}^{k}$ is random. By construction and Lemma 9, $V_{h+1}^{k}$ must lie in the set

$$
\begin{aligned}
\mathcal{V}:=\{V: V(\cdot) & =\max _{a}\left[\phi(\cdot, a)^{\top} w+\beta \sqrt{\phi(\cdot, a) \Lambda^{-1} \phi(\cdot, a)}\right] \\
& w \in \mathbb{R}^{d} \text { with }\|w\| \leq H \sqrt{d k}, \Lambda \in \mathbb{R}^{d \times d} \text { with } \lambda_{\min }(\Lambda) \geq 1
\end{aligned}
$$

The $\epsilon$-covering number of $\mathcal{V}$ is $N \approx\left(\frac{H \sqrt{d k}}{\epsilon}\right)^{d}\left(\frac{\beta \sqrt{d}}{\epsilon}\right)^{d^{2}}$, since we need to cover the sets $\left\{w \in \mathbb{R}^{d}:\|w\| \leq H \sqrt{d k}\right\}$ and $\left\{A \in \mathbb{R}^{d \times d}: A=\Lambda^{-1}, \lambda_{\max }(A)=\lambda_{\min }(\Lambda)^{-1} \leq 1\right\}$.

Applying 11 with $\delta=\frac{p / 2}{N}$ and running an $\epsilon$-net $\operatorname{argument~to~all~possible~} V_{h+1}^{k}$ in $\mathcal{V}$, we obtain the desired inequality.

Note that the $d$ factor on the RHS of the lemma statement comes from $\sqrt{\log N} \approx d$.

## Appendices

## A Technical lemmas

We begin with a simple upper bound on the Gram matrix.
Lemma 8 (Simple upper bound). If $\Lambda_{t}=\lambda I+\sum_{i \in[t]} \phi_{i} \phi_{i}^{\top} \in \mathbb{R}^{d}$ and $\lambda>0$, then

$$
\sum_{i \in[t]} \phi_{i}^{\top} \Lambda_{t}^{-1} \phi_{i} \leq d
$$

Proof If $\lambda=0$, then it is easy to see that $\sum_{i \in[t]} \phi_{i}^{\top} \Lambda_{t}^{-1} \phi_{i}=\operatorname{tr}\left(I_{d}\right)=d$. The regularization $\lambda>0$ only makes the LHS smaller.

The next lemma ensures boundedness of the linear weights.
Lemma 9 (Weights are bounded). (i) For each policy $\pi$ and its $Q$ function $Q_{h}^{\pi}(x, a)=\left\langle\phi(x, a)\right.$, $\left.w_{h}^{\pi}\right\rangle$, we have $\left\|w_{h}^{\pi}\right\| \leq 2 H \sqrt{d}, \forall h$. (ii) The weights $\left\{w_{h}^{k}\right\}$ in the LSVI-UCB algorithm satisfies $\left\|w_{h}^{k}\right\| \leq 2 H \sqrt{d k}, \quad \forall k, h$.

Proof Part (i) follows from Assumption 1 on linearity and boundedness.Part (ii) holds since the Gram $\operatorname{matrix} \Lambda_{h}^{k}$ has minimum eigenvalue $\geq 1$ and satisfies the bound in Lemma 8 .

Lemma 10 (Concentration for self-normalized processes Abbasi-Yadkori et al., 2011, Theorem 1]). Suppose $\left(\epsilon_{s}\right)_{s=1,2, \ldots}$ is a scalar stochastic process adapted to the filtration $\left(\mathcal{F}_{s}\right)$, and $\epsilon_{s} \mid \mathcal{F}_{s-1}$ is zero mean and $\sigma$-subGaussian. Let $\left(\phi_{s}\right)_{s=1,2, \ldots}$ be an $\mathbb{R}^{d}$-valued stochastic process with $\phi_{s} \in \mathcal{F}_{s-1}$. Let $\Lambda_{t}=I+\sum_{s=1}^{t} \phi_{s} \phi_{s}^{\top} \in$ $\mathbb{R}^{d \times d}$. Then with probability at least $1-\delta$, we have

$$
\left\|\sum_{s=1}^{t} \phi_{s} \epsilon_{s}\right\|_{\Lambda_{t}^{-1}}^{2} \leq 2 \sigma^{2} \log \left[\frac{\operatorname{det}\left(\Lambda_{t}\right)^{1 / 2}}{\delta}\right], \quad \forall t \geq 0
$$

## References

[Abbasi-Yadkori et al., 2011] Abbasi-Yadkori, Y., Pál, D., and Szepesvári, C. (2011). Improved algorithms for linear stochastic bandits. Advances in Neural Information Processing Systems, pages 2312-2320.
[Jin et al., 2019] Jin, C., Yang, Z., Wang, Z., and Jordan, M. I. (2019). Provably efficient reinforcement learning with linear function approximation.


[^0]:    ${ }^{1}$ Reading: Jin et al., 2019
    ${ }^{2}$ This normalization ensures consistency when reducing to tabular case.

