CS 839 Probability and Learning in High Dimension

Lecture 4 - 02/07/2022

Lecture 4: Matrix Concentration II

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In this lecture,¹ we will complete the proof of Matrix Bernstein's inequality. We will also introduce the scalar Hoeffding and Bernstein's inequalities.

1 Notation

A quick summary of the notation:

For a matrix $A \in \mathbb{R}^{d_1 \times d_2}$, we use $||A||_{op}$ to denote its operator/spectral norm (i.e., the largest singular value of A.

For two $d \times d$ symmetric matrices A, B, the positive-semidefinite ordering $A \succeq B$ means that A - B is positive-semidefinite matrix, i.e., $\lambda_i(A - B) \ge 0, \forall i$. It implies that $\lambda_i(A) \ge \lambda_i(B)$, for i = 1, ..., d.

2 Recap

We have introduced the statement of Matrix Bernstein's Inequality, and covered several theorems which would be leveraged for the proof of Matrix Bernstein's Inequality.

Theorem 1 (Matrix Bernstein's Inequality). Suppose $X_1, ..., X_n \in \mathbb{R}^{d_1 \times d_2}$ are independent random matrices satisfying the following conditions:

- $\mathbb{E}[X_i] = 0$, for all $i \in \{1, 2, ..., n\}$,
- $||X_i||_{op} \leq b$ almost surely for all $i \in \{1, 2, ..., n\}$,

• max
$$\left[\left\| \mathbb{E} \sum_{i=1}^{n} X_{i} X_{i}^{\top} \right\|_{op}, \left\| \mathbb{E} \sum_{i=1}^{n} X_{i}^{\top} X_{i} \right\|_{op} \right] \leq \sigma^{2},$$

then for every t > 0, we have

$$\mathbb{P}\left[\left\|\sum_{i=1}^{n} X_i\right\|_{op} \ge t\right] \le (d_1 + d_2) \exp\left(\frac{-t^2/2}{\sigma^2 + bt/3}\right).$$
(1)

Lemma 1 (Matrix Laplace Transform). Suppose Y is a random symmetric matrix, then for any $t \in \mathbb{R}$,

$$\mathbb{P}(\lambda_{\max}(Y) \ge t) \le \inf_{\theta > 0} e^{-\theta t} \cdot \mathbb{E}[\mathsf{tr}(e^{\theta Y})],$$

where $\lambda_{\max}(Y)$ is the maximum eigenvalue of Y.

Theorem 2 (Lieb's Theorem). For a fixed symmetric matrix H, the matrix function f defined through

$$f(A) \triangleq \operatorname{tr} \exp(H + \log A)$$

is concave on the space of positive symmetric matrices with the same size as H.

The Lieb's Theorem together with Jensen's inequality implies that

$$\mathbb{E}[\operatorname{tr}\exp(H+X)] \le \operatorname{tr}\exp(h + \log \mathbb{E}[e^X]).$$
(2)

¹Reading: Section 6 in J. Tropp, An Introduction to Matrix Concentration Inequalities. https://arxiv.org/abs/1501.01571

3 Matrix Theory Background, cont'd

With Lieb's theorem, we can proof the following lemma, which is the key in the proof of Matrix Bernstein's Inequality.

Lemma 2 (Sub-additivity of Matrix MGF). Suppose $X_1, ..., X_n$ are independent symmetric matrices, then

$$\mathbb{E}\left[\operatorname{tr}\exp(\theta\sum_{i=1}^{n}X_{i})\right] \leq \operatorname{tr}\exp\left(\sum_{i=1}^{n}\log\mathbb{E}[e^{\theta X_{i}}]\right), \quad \text{for all } \theta.$$
(3)

Proof We note that the summation of $\theta \sum_{i=1}^{n} X_i$ can be decomposed and the Equation (2) implied by the Lieb's Theorem can be applied iteratively as follows:

$$\mathbb{E}\left[\operatorname{tr} \exp(\theta \sum_{i=1}^{n} X_{i})\right] = \mathbb{E}\left[\operatorname{tr} \exp\left(\theta \sum_{i=1}^{n-1} X_{i} + \theta X_{n}\right)\right]$$
$$= \mathbb{E}\left[\mathbb{E}\left[\operatorname{tr} \exp\left(\theta \sum_{i=1}^{n-1} X_{i} + \theta X_{n}\right) | X_{1}, ..., X_{n-1}\right]\right]$$
$$\stackrel{(i)}{\leq} \mathbb{E}\left[\operatorname{tr} \exp\left(\theta \sum_{i=1}^{n-1} X_{i} + \log \mathbb{E}[\exp(\theta X_{n}) | X_{1}, ..., X_{n-1}]\right)\right]$$
$$\stackrel{(ii)}{=} \mathbb{E}\left[\operatorname{tr} \exp\left(\theta \sum_{i=1}^{n-2} X_{i} + \log \mathbb{E}[\exp(\theta X_{n})] + \theta X_{n-1}\right)\right]$$
$$\leq ...$$
$$\leq \operatorname{tr} \exp\left(\sum_{i=1}^{n} \log \mathbb{E}[\exp(\theta X_{i})]\right),$$

where step (i) follows from invoking the inequality (2), and step (ii) follows from independence.

By combining Lemma 1 and Lemma 2, the following Theorem 3 (Master Bound) can be derived.

Theorem 3 (Master Bound). Suppose $X_1, ..., X_n$ are independent symmetric matrices, then for any $t \in \mathbb{R}$,

$$\mathbb{P}\left[\lambda_{\max}(Y) \ge t\right] \le \inf_{\theta > 0} e^{-\theta t} \cdot \operatorname{tr} \exp\left(\sum_{i=1}^{n} \log \mathbb{E}[e^{\theta X_i}]\right)$$

The master bound (and its variants) can be used to prove the matrix versions of different inequalities, such as Hoeffding, Bernstein, Chernoff, Azuma, Bounded Difference, Bennett, Freeman.

4 Proof of the Matrix Bernstein's Inequality

We shall prove the following symmetric version of the Matrix Bernstein's inequality.

Theorem 4 (Matrix Bernstein's Inequality: Symmetric Case). Suppose $X_1, ..., X_n \in \mathbb{R}^{d \times d}$ are independent symmetric random matrices with the following conditions:

- $\mathbb{E}[X_i] = 0$, for all $i \in \{1, 2, ..., n\}$,
- $\lambda_{\max}(X_i) \leq b$ almost surely for all $i \in \{1, 2, ..., n\}$,
- $\|\sum_{i=1}^n X_i\|_{op} \le \sigma^2$,

then for every t > 0, we have

$$\mathbb{P}\left[\lambda_{\max}(\sum_{i=1}^{n} X_i) \ge t\right] \le d \exp\left(\frac{-t^2/2}{\sigma^2 + bt/3}\right).$$
(4)

Remark To prove the rectangular version of Matrix Bernstein's Inequality in Theorem 1, we can apply Theorem 4 to the symmetric matrix $Y = \begin{bmatrix} X \\ X^{\top} \end{bmatrix}$, where $X \in \mathbb{R}^{d_1 \times d_2}$ is an general rectangular matrix and $Y \in \mathbb{R}^{(d_1+d_2) \times (d_1+d_2)}$ called the symmetric dilation of X. We can then use $\lambda_{\max}(Y) = \|X\|_{op}$ to finish the proof.

Proof Define the scalar function f as

$$f(x) \triangleq \frac{e^{\theta x} - 1 - \theta x}{x^2}$$

which is an increasing function.

For any symmetric matrix X with $\lambda_{\max}(X) \leq b$, we have

$$e^{\theta X} = I + \theta X + X f(X) X$$
$$\leq I + \theta X + f(b) X^{2}.$$

We use a scalar inequality: for any $\theta : 0 < \theta < \frac{3}{b}$, it holds that

$$f(b) = \frac{e^{\theta b} - 1 - \theta b}{b^2}$$
$$= \frac{1}{b^2} \sum_{k=2}^{\infty} \frac{(\theta b)^k}{k!}$$
$$\leq \frac{\theta^2}{2} \sum_{k=2}^{\infty} \frac{(\theta b)^{k-2}}{3^{k-2}}$$
$$= \frac{\theta^2/2}{1 - \theta b/3}.$$

Combining the above inequalities, we have for any $\theta : 0 < \theta < \frac{3}{b}$, and any $X : \lambda_{\max}(X) \leq b$, it holds that

$$e^{\theta X} \preceq I + \theta X + \frac{\theta^2/2}{1 - \theta b/3} X^2.$$

It follows that

$$\mathbb{E}[e^{\theta X}] \leq I + 0 + \frac{\theta^2/2}{1 - \theta b/3} \mathbb{E}[X^2]$$
$$\leq \exp\left(\frac{\theta^2/2}{1 - \theta b/3} \mathbb{E}[X^2]\right),$$

Since matrix logarithm is operator monotone, we obtain that

$$\log \mathbb{E}[e^{\theta X}] \preceq \frac{\theta^2/2}{1 - \theta b/3} \mathbb{E}[X^2].$$
(5)

Letting $g(\theta) := \frac{\theta^2/2}{1-\theta b/3}$, we have

$$\mathbb{P}\left[\lambda_{\max}(\sum_{i=1}^{n} X_{i}) \geq t\right] \leq \inf_{\theta > 0} \frac{\operatorname{tr} \exp\left(\sum_{i=1}^{n} \log \mathbb{E}[e^{\theta X_{i}}]\right)}{e^{\theta t}}$$
$$\leq \inf_{0 < \theta < \frac{3}{b}} \frac{\operatorname{tr} \exp\left(\sum_{i=1}^{n} g(\theta) \mathbb{E}[X_{i}^{2}]\right)}{e^{\theta t}}$$
$$\leq \inf_{0 < \theta < \frac{3}{b}} \frac{d \cdot \exp\left(g(\theta)\sigma^{2}\right)}{e^{\theta t}},$$

Here the first inequality is due to the Theorem 3 (Master Bound), and the second inequality is due to the Equation (5).

Taking $\theta = t/(\sigma^2 + bt/3)$ and simplifying the expression, we have the desired bound

$$\mathbb{P}\left[\lambda_{\max}(\sum_{i=1}^{n} X_i) \ge t\right] \le d \cdot \exp\left(\frac{-t^2/2}{\sigma^2 + bt/3}\right).$$

Remark

- The proof is quite similar to that of the scalar Bernstein's Inequality. Most of the hard work is done by the Lieb's Theorem.
- The dimension factor on the right hand side of Equation (4) leads to the $\sqrt{\log d}$ factor in the user-friend form of the matrix Bernstein's inequality (see last lecture).
- To prove a tighter bound and relax this dimension dependence, one must better capture non-commutativity. There is active research on achieving such improvement but it is outside the scope of this course.

5 Sub-Gaussian/Exponential Random Variables and Scalar Hoeffding/Bernstein Inequalities

In this section, 2,3,4 we will briefly introduce the scalar versions of Hoeffding's Inequality and Bernstein's Inequality.

Definition 1. A variable X is called sub-Gaussian with parameter σ^2 , denoted as sub-Gaussian(σ^2), if

$$\mathbb{E} e^{\lambda (X - \mathbb{E}[X])} \le e^{\lambda^2 \sigma^2 / 2}, \qquad \text{for all } \lambda \in \mathbb{R}.$$
(6)

Note that the right hand side above is the MGF of a zero-mean Gaussian random variable with variance σ^2 .

Example 1 (Rademacher). If a random variable $X \in \{-1, +1\}$ with the equal probability 1/2 for -1 and 1, then X is sub-Gaussian (1^2) .

²Reading: Martin J. Wainwright. High-Dimensional Statistics: A Non-Asymptotic Viewpoint. Section 2.1

 $^{^{3}}Reading:$ Roman Vershynin. High-Dimensional Probability: An Introduction with Applications in Data Science. Section 2

⁴Reading: John Duchi's Lecture Notes, Section 3.1 https://web.stanford.edu/class/stats311/lecture-notes.pdf

Proof For all λ , we have

$$\mathbb{E} e^{\lambda X} = \frac{1}{2} e^{\lambda} + \frac{1}{2} e^{-\lambda}$$

$$= \frac{1}{2} \sum_{k=0}^{\infty} \frac{\lambda^k}{k!} + \frac{1}{2} \sum_{k=0}^{\infty} \frac{(-\lambda)^k}{k!}$$

$$= \sum_{k=0}^{\infty} \frac{\lambda^{2k}}{(2k)!}$$

$$\leq \sum_{k=0}^{\infty} \frac{(\lambda^2)^k}{2^k k!}$$

$$= e^{\lambda^2/2}.$$

Example 2 (Bounded RV). If a random variable $X \in [a, b]$ with probability 1, then X is sub-Gaussian $((b - a)^2)$.

Proof We prove this by a symmetrization argument. Let $\varepsilon \in \{-1, +1\}$ be a random variable with equal probability 1/2 for -1 and +1, and X' be an independent copy of X. Then we have $\mathbb{E}[X'] = \mathbb{E}[X]$ and

$$\mathbb{E} e^{\lambda(X - \mathbb{E}[X])} = \mathbb{E} e^{\lambda(X - \mathbb{E}[X'])}$$

$$\leq \mathbb{E} e^{\lambda(X - X')}$$

$$= \mathbb{E} e^{\lambda \varepsilon (X - X')}$$

$$= \mathbb{E} [\mathbb{E} [e^{\lambda \varepsilon (X - X')} | X, X']]$$

$$\leq \mathbb{E} [e^{\lambda^2 (X - X')^2} / 2]$$

$$\leq e^{\lambda^2 (b - a)^2 / 2}.$$

The first inequality follows from the Jensen's Inequality, and the second inequality follows from the previous example. $\hfill \Box$

The lemma below provides an equivalent characterization of a sub-Gaussian random variable in terms of its tail probability.

Lemma 3. A variable X is sub-Gaussian(σ^2) if and only if for some universal constant c > 0,

$$\mathbb{P}[|X| \ge t] \le 2e^{-t^2/c\sigma^2}, \quad \text{for all } t \ge 0.$$
(7)

We now state the Hoeffding's inequality for sum of independent sub-Gaussian random variables. It generalizes the more commonly known Hoeffding's inequality for sum of bounded random variables.

Theorem 5 (Hoeffding's Inequality). If X_i 's are independent sub-Gaussian (σ_i^2) random variables, then for any $t \ge 0$,

$$\mathbb{P}\left[\left|\sum_{i} (X_i - \mathbb{E} X_i)\right| \ge t\right] \le 2 \exp\left(-\frac{t^2}{2\sum_{i} \sigma_i^2}\right),$$

i.e., $\sum_{i} X_i$ is sub-Gaussian $(\sum_{i} \sigma_i^2)$.

One consequence of Hoeffding's Inequality is

$$\left|\frac{1}{n}\sum_{i}(X_{i}-\mathbb{E}X_{i})\right| \lesssim \frac{\sigma\sqrt{\log 1/\delta}}{\sqrt{n}},$$

with probability $1 - \delta$.

Definition 2. A variable X is called sub-exponential with parameters (τ^2, b) , denoted as sub-exponential (τ^2, b) , if

$$\mathbb{E} e^{\lambda(X - \mathbb{E}[X])} \le e^{\lambda^2 \tau^2/2}, \quad \text{for all } \lambda \in \mathbb{R} \text{ with } |\lambda| \le \frac{1}{b}.$$
(8)

Example 3 (Gaussian Squared). If a variable Z follows the normal distribution N(0,1) and $X \triangleq Z^2$, then X is sub-exponential (2,4).

Example 4 (Bounded RV). If a variable $X \in [-b, b]$, with mean $\mathbb{E}[X] = 0$ and variance $var(X) = \sigma^2$, then X is sub-Gaussian($(2b)^2$), and is also sub-exponential($6\sigma^2/5, 2b$).

We now state the Bernstein's inequality for sum of independent sub-exponential random variables. It generalizes the more commonly known Bernstein's inequality for sum of bounded random variables.

Theorem 6 (Bernstein's Inequality). If X_i 's are independent sub-exponential(σ_i^2, b_i) random variables with $\mathbb{E}[X_i] = 0$, then for any $t \ge 0$,

$$\mathbb{P}\left[\left|\sum_{i} X_{i}\right| \geq t\right] \leq 2 \exp\left(-\frac{1}{2} \min\left\{\frac{t^{2}}{\sum_{i} \sigma_{i}^{2}}, \frac{t}{\max_{i} b_{i}}\right\}\right).$$